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## AN AFFINE PROPERTY OF THE RECIPROCAL ASIAN OPTION PROCESS

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This note describes a property of the Asian option process, which neatly links the process at  $\mu$  with the one at  $-\mu$ , where  $\mu$  is the drift of the geometric Brownian motion. The proof is based on (i) a known result due to Yor, on the law of the Asian option process taken at an exponential time, and (ii) a recent result on beta and gamma distributions.

Suppose  $W$  is one-dimensional standard Brownian motion starting at the origin, and define what this author calls the *Asian option process*, for want of a better name:

$$A_t^{(\mu)} = \int_0^t e^{2\mu s + 2W_s} ds, \quad t \geq 0, \quad \mu \in \mathbb{R}.$$

Asian options have payoffs such as  $(A_t^{(\mu)} - K)_+$ , and have been studied by numerous authors in Finance and Mathematics; reciprocal Asian options have payoffs such as  $(K - 1/A_t^{(\mu)})_+$ , and have not received much attention so far; for more details and references, the reader is referred to [3] and [4].

**Theorem 1** ([6]). *Let  $T_\lambda$  be an exponentially distributed random variable, independent of  $W$ , with mean  $1/\lambda$ . Then*

$$2A_{T_\lambda}^{(\mu)} \stackrel{\mathcal{L}}{=} \frac{B_{1,\alpha}}{G_\beta},$$

where  $B_{1,\alpha} \sim \text{Beta}(1, \alpha)$  and  $G_\beta \sim \Gamma(\beta, 1)$  are independent,  $\alpha = \mu/2 + \sqrt{2\lambda + \mu^2}/2$ ,  $\beta = \alpha - \mu$ .

**Theorem 2** ([2]). *For any  $a, b, c > 0$ ,*

$$\frac{G_a}{B_{b,a+c}} + G'_c \stackrel{\mathcal{L}}{=} \frac{G_{a+c}}{B_{b,a}}.$$

where  $G_a \sim \text{Gamma}(a, 1)$ ,  $G'_c \sim \text{Gamma}(c, 1)$ ,  $B_{b,a+c} \sim \text{Beta}(b, a + c)$ ,  $B_{b,a} \sim \text{Beta}(b, a)$  and all variables are independent.

**Theorem 3.** For any  $\mu, t > 0$ ,

$$\frac{1}{2A_t^{(\mu)}} + G_\mu \stackrel{\mathcal{L}}{=} \frac{1}{2A_t^{(-\mu)}},$$

where  $G_\mu \sim \Gamma(\mu, 1)$  is independent of  $W$ .

The last result follows directly from the two previous ones, upon setting  $a = \beta$ ,  $b = 1$ ,  $c = \mu$ , and then inverting the Laplace transform represented by the exponential time  $T_\lambda$ . Theorem 3 gives an easy proof of the well-known formula in Corollary 4 (just observe that  $A_\infty^{(\mu)} = \infty$  a.s. if  $\mu \geq 0$ ).

**Corollary 4.** For any  $\mu > 0$ ,

$$\frac{1}{2A_\infty^{(-\mu)}} \sim \text{Gamma}(\mu, 1).$$

**Theorem 5.** Let  $\{U_k; k \geq 1\}$  be independent variables with the same distribution as

$$U \stackrel{\mathcal{L}}{=} \frac{B_1}{B_2}, \quad B_1 \sim \text{Beta}(\beta, \mu), \quad B_2 \sim \text{Beta}(1 + \beta, \mu),$$

with  $B_1, B_2$  independent,  $\mu > 0$ ,  $\beta = -\mu/2 + \sqrt{2\lambda + \mu^2}/2$ , and let  $\{G_\mu^{(k)}; k \geq 0\}$  be a sequence of independent variables with a common  $\text{Gamma}(\mu, 1)$  distribution, independent of  $\{U_k; k \geq 1\}$ . Then

- (a)  $\frac{1}{2A_{T_\lambda}^{(\mu)}} \stackrel{\mathcal{L}}{=} U \left( \frac{1}{2A_{T_\lambda}^{(\mu)}} + G_\mu \right)$
- (b)  $\frac{1}{2A_{T_\lambda}^{(\mu)}} \stackrel{\mathcal{L}}{=} \sum_{k=1}^\infty U_1 \cdots U_k G_\mu^{(k)}; \quad \frac{1}{2A_{T_\lambda}^{(-\mu)}} \stackrel{\mathcal{L}}{=} G_\mu^{(0)} + \sum_{k=1}^\infty U_1 \cdots U_k G_\mu^{(k)}$
- (c)  $\mathbf{E} e^{-s/2A_{T_\lambda}^{(\mu)}} = \mathbf{E} \left( \prod_{k=1}^\infty \frac{1}{1 + sU_1 \cdots U_k} \right)^\mu$
- (d)  $\mathbf{E} e^{-s/2A_{T_\lambda}^{(-\mu)}} = \mathbf{E} \left( \frac{1}{1 + s} \prod_{k=1}^\infty \frac{1}{1 + sU_1 \cdots U_k} \right)^\mu = \left( \frac{1}{1 + s} \right)^\mu \mathbf{E} e^{-s/2A_{T_\lambda}^{(\mu)}}.$

In (a),  $A_{T_\lambda}^{(\mu)}$  and  $G_\mu$  are independent; moreover, given  $G_\mu$  and  $U$  with the given distributions, the solution  $1/2A_{T_\lambda}^{(\mu)}$  is unique (in distribution).

Part (a) follows from computing the Mellin transform ( $s \mapsto \mathbf{E}(\cdot)^s$ ) of either side,

which from Theorem 1 is

$$\frac{\Gamma(1-s)\Gamma(\beta+s)\Gamma(1+\beta+\mu)}{\Gamma(1+\beta+\mu-s)\Gamma(\beta)} = \frac{\Gamma(\beta+s)\Gamma(\beta+\mu)}{\Gamma(\beta)\Gamma(\beta+\mu+s)} \frac{\Gamma(1+\beta-s)\Gamma(1+\beta+\mu)}{\Gamma(1+\beta)\Gamma(1+\beta+\mu-s)} \frac{\Gamma(1-s)\Gamma(\alpha+s)\Gamma(1+\alpha-\mu)}{\Gamma(1+\alpha-\mu-s)\Gamma(\alpha)}.$$

Uniqueness of the solution is essentially a consequence of  $E \log U < 0$ , see [5], Theorem 1.5. Part (b) results from iterating (a) (see also Theorems 3 and 3 in [2]). Parts (c) and (d) result from conditioning on  $\{U_k; k \geq 1\}$  in (b).

Theorem 5 (b) is another instance of the relationship between perpetuities and the Asian option process, observed in [1]. Finally, note that Theorem 5 (a) does not say anything about  $E 1/A_t^{(\mu)}$ , as  $EU = 1$  and (by (b))  $E 1/A_{T_\lambda}^{(\mu)} = \infty$ . The expectation of  $1/A_t^{(\mu)}$  is obtained by other means in [3].

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