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Author(s)	Hitsuda, Masuyuki
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REPRESENTATION OF GAUSSIAN PROCESSES EQUIVALENT TO WIENER PROCESS[1]

MASUYUKI HITSUDA

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1. Introduction

The purpose of this paper is to get a canonical representation of Gaussian processes which are equivalent (or mutually absolutely continuous) to Wiener process. The main result is this. Suppose we are given a Gaussian process Y_t on a probability space $(\Omega, \mathfrak{B}, P)$, which is equivalent to Wiener process. Then a Wiener process X_t is constructed on $(\Omega, \mathfrak{B}, P)$ as a functional of $\{Y_s; s\leqq t\}$ and, conversely, Y_t is represented as a measurable functional of $\{X_s; s\leqq t\}$ for each $t\in[0, T]$. In case of $E(Y_t)=0$, $t\in[0, T]$, Y_t is represented by the formula

$$Y_t = X_t-\int_0^t\left(\int_0^s l(s, u)dX_u\right)ds\,, \tag{1.1}$$

where $l(s, t)$ is a Volterra kernel belonging to $L^2([0, T]^2)$[2], and the representation of Y_t is unique (Theorem 1). Conversely, if a Wiener process X_t is given and if Y_t is represented by (1.1), Y_t is equivalent to Wiener process (Theorem 2). The density of the transformed measure with respect to which Y_t is a Wiener process is easily evaluated in the proof of Theorem 2. The proof of these facts is based heavily upon martingale theory due to Meyer [8] and Kunita-S. Watanabe [6].

The conditions for a Gaussian process to be equivalent to Wiener process have been obtained in terms of the mean and the covariance by Shepp [11] and Golosov [3]. Moreover, by the method of linear transformations of Wiener space, Shepp [11] and H. Sato [10] have obtained the following representation

$$Y_t = X_t-\int_0^t\left(\int_0^T m(s, u)dX_u\right)ds\,,$$

where $m(s, u)$ is a kernel of $L^2([0, T]^2)$ with some additional conditions. This representation involves the stochastic integral on the fixed time interval $[0, T]$, so that it does not assert even the fact that if the Gaussian process is equivalent

1) This work was partially supported by the Yukawa Foundation.
2) The Volterra kernel $l(s, t)\in L^2([0, T]^2)$ means $l(s, u)=0$ for $s<u$.

to Wiener process in the time interval $[0, T]$, it is so in any subinterval $[0, t_0]$ for each $t_0 \in [0, T]$. Such a fact is clarified in the representation (1.1).

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2. Preliminaries

Let $(\Omega, \mathfrak{B}, P)$ be a complete probability space, $\{\mathfrak{F}_t;\ t \in [0, T]\}$ a system of σ-subalgebras of $\mathfrak{B}$ which are increasing in t, and $\{X_t;\ t \in [0, T]\}$ a stochastic process on $(\Omega, \mathfrak{B}, P)$. In the following discussion, time interval $[0, T]$ will be fixed.

DEFINITION 1. When $(X_t, \mathfrak{F}_t, P)$ satisfies the following conditions 1), 2) and 3), it is called a *Wiener process*:

1) The sample paths of X_t are continuous in t, and $X_0 = 0$.

2) For $t \geqq s$, $t, s \in [0, T]$, $E(X_t | \mathfrak{F}_s) = X_s$ with P-measure 1, where $E(\cdot | \cdot)$ denotes the conditional expectation with respect to the measure P.

3) $E((X_t - X_s)^2 | \mathfrak{F}_s) = t - s$ with P-measure 1, for $t \geqq s$, $t, s \in [0, T]$.

This definition of Wiener process is due to Doob [1].

DEFINITION 2. A stochastic process Y_t, definied on $(\Omega, \mathfrak{B}, P)$ (or simply, (Y_t, P)) is called a *Gaussian process*, when the distribution of $(Y_{t_1}, Y_{t_2}, \cdots, Y_{t_N})$ with respect to P is subject to an N-dimensional Gaussian distribution.

Let (Y_t, P) be a stochastic process. Let $\tilde{P}$ be a probability measure on $(\Omega, \mathfrak{B})$ such that P and $\tilde{P}$ are mutually absolutely continous, that is, $\tilde{P}(d\omega) = \varphi(\omega) P(d\omega)$ with a strictly positive φ. Let $\mathfrak{Y}_t$ be the σ-subalgebra of $\mathfrak{B}$, generated by $\{Y_s, s \leqq t\}$, adjoined with all P-negligible sets. Note that the notion of negligible sets is identical for both P and $\tilde{P}$.

DEFNITION 3. A stochastic process (Y_t, P) is said to be *equivalent to Wiener process* when there is a probability measure

$$\tilde{P}(d\omega) = \varphi(\omega) P(d\omega),$$

such that P and $\tilde{P}$ are mutually absolutely continuous and such that $(Y_t, \mathfrak{Y}_t, \tilde{P})$ is a Wiener process.

REMARK 1. Suppose $\mathfrak{B} = \mathfrak{Y}_T$. Let $\tilde{P}$ be absolutely continous relative to P, that is, $\tilde{P}(d\omega) = \varphi(\omega) P(d\omega)$ with non-negative φ. If Y_t is Gaussian with respect to both P and $\tilde{P}$, then P and $\tilde{P}$ are mutually absolutely continous by Hajék and Feldman's result (see Rozanov [9]).

REMARK 2. When (Y_t, P) is equivalent to Wiener process, we can assume that the sample paths of Y_t are continous by choosing a suitable modification,

3. Necessary condition and uniqueness

The purpose of this section is to prove the following theorem.

Theorem 1. *Suppose that a Gaussian process* (Y_t, P), $t\in[0, T]$, *with mean* 0, *is equivalent to Wiener process. Then there exists a Wiener process* $(X_t, \mathfrak{Y}_t, P)$ *and a Volterra kernel* $l(s, u)\in L^2([0, T]^2)$ *such that* Y_t *is represented, with P-measure* 1, *by the formula*

$$Y_t = X_t-\int_0^t\left(\int_0^s l(s, u)\,dX_u\right)ds \qquad \textit{for every } t\in[0, T]\,. \tag{3.1}$$

Moreover, such X_t *and* $l(s, u)\in L^2([0, T]^2)$ *are unique.*

For the proof of Theorem 1, we need some lemmas. Let $\varphi(\omega)$ be the density $d\tilde{P}(\omega)/dP(\omega)$ in Definition 3.

Lemma 1. *Let* M_t *be a right continuous modification of the martingle* $\tilde{E}\left(\frac{1}{\varphi}\,|\,\mathfrak{Y}_t\right)$ *with respect to* $(\mathfrak{Y}_t, P)$. *Then,*

1) $$P(M_t>0 \quad \textit{for} \quad t\in[0, T]) = 1\,.$$

2) M_t *is represented by*

$$M_t = \exp\left\{\int_0^t f(s, \omega)\,dY_s-\frac{1}{2}\int_0^t f^2(s, \omega)\,ds\right\} \qquad \textit{for any} \quad t\in[0, T]\,,$$

where $f(s, \omega)$ *is a function which is* (i) (s, ω)*-measurable and* (ii) $\mathfrak{Y}_s$*-measurable for each* $s\in[0, T]$, *and which satisfies*

(iii) $$P\left(\int_0^T f^2(s, \omega)ds<\infty\right) = 1\,.$$

Proof. Let us put

$$\tau_0 = \begin{cases} \inf\{t;\, M_t = 0\} & \\ T & , \quad \text{if} \quad \{t;\, M_t = 0\} = \phi\,, \end{cases}$$

then τ_0 is a stopping time relative to $\{\mathfrak{Y}_t\}$. By the optional sampling theorem

$$\tilde{E}(M_T\,|\,\mathfrak{Y}_{\tau_0}) = M_{\tau_0} \quad \text{with } P\text{-measure } 1.$$

So we get

$$\int_{\{\tau_0<T\}} M_T\tilde{P}(d\omega) = \int_{\{\tau_0<T\}} M_{\tau_0}\,\tilde{P}(d\omega) = 0.$$

On the other hand, since $M_T(\omega)=\tilde{E}(\varphi^{-1}\,|\,\mathfrak{Y}_T)>0$, we get $P(\tau_0<T)=0$ from the above equality. The proof of 1) is finished.

Since $(Y_t, \mathfrak{Y}_t, \tilde{P})$ is a Wiener process, the martingle M_t is represented by

$$M_t = \int_0^t g(s, \omega) dY_s + 1 \qquad \text{for any} \quad t \in [0, T], ^{3)}$$

according to Kunita-S. Watanabe [6], where $g(s, \omega)$ satisfies (i), (ii) and (iii) of this lemma by replacing $f(s, \omega)$ with $g(s, \omega)$. We can now apply Itô's formula [5] and we get

$$\log M_t = \log(1 + \int_0^t g(s, \omega) dY_s)$$
$$= \int_0^t \frac{1}{M_s} g(s, \omega) dY_s - \frac{1}{2} \int_0^t \frac{1}{M_s^2} g(s, \omega)^2 ds .$$

Put

$$f(s, \omega) = \frac{1}{M_s} g(s, \omega),$$

then $f(s, \omega)$ satisfies (i), (ii) and (iii). The proof of this lemma is finished.

Lemma 2. (Girsanov [2]). *Let $f(s, \omega)$ be the function of Lemma* 1 *and let*

$$X_t = Y_t - \int_0^t f(s, \omega) ds . \tag{3.2}$$

Then $(X_t, \mathfrak{Y}_t, P(d\omega))$ is a Wiener process.

The next lemma plays an important role.

Lemma 3. *Under the same assumption as in Lemma* 1, *if (Y_t, P) is a Gaussian process, then it follows that*

$$E\left(\int_0^T f(s, \omega)^2 ds\right) < \infty \quad \textit{and} \quad \tilde{E}\left(\int_0^T f(s, \omega)^2 ds\right) < \infty .$$

Proof. First let us note the fact that

$$\tilde{E}(M_T \log M_T) = K < \infty,$$

which is due to Hajék-Feldman (see Rozanov [9] for a simple proof), because Y_t is a Gaussian process with respect to two measures P and $\tilde{P}$. Since $x \log x$ is a convex function, $\{M_t \log M_t\}_{t \in [0,T]}$ is a submartingale with respect to $\{\mathfrak{Y}_t\}$ in the probability space $\{\Omega, \mathfrak{B}, \tilde{P}\}$. Therefore, by the optional sampling theorem,

$$K \geqq \tilde{E}(M_{T_n \wedge T} \log M_{T_n \wedge T}),$$

where $\{T_n\}_{n=1,2,\dots}$ is an arbitrary sequence of stopping times increasing to T. Moreover,

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3) See Supplement.

$$\begin{aligned}\tilde{E}(M_{T_n\wedge T}\log M_{T_n\wedge T}) &= \tilde{E}(M_T\log M_{T_n\wedge T})\\ &= E\Big(\int_0^{T_n\wedge T} f\,dY_s - \frac{1}{2}\int_0^{T_n\wedge T} f^2 ds\Big)\\ &= E\Big(\int_0^{T_n\wedge T} f\,dX_s + \frac{1}{2}\int_0^{T_n\wedge T} f^2 ds\Big).\end{aligned}$$

The last equality holds by Girsanov [2]. The process $\left\{\int_0^t f(s,\omega)\,dX_s\right\}_{t\in[0,T]}$ is a local martingale[4] with respect to $\{\mathfrak{Y}_t\}$ in $(\Omega, \mathfrak{B}, P)$, so we can choose a sequence $\{T_n\}$ such that $\left\{\int_0^{t\wedge T_n} f(s,\omega)dX_s\right\}_{t\in[0,T]}$ is a martingale for every n. Then, the last expectation above is $\frac{1}{2}E\Big(\int_0^{T_n\wedge T} f^2 ds\Big)$ for every n, and we get the first part of this lemma as n tends to ∞. On the other hand,

$$M_t^{-1} = \exp\Big(-\int_0^t f dX_s - \frac{1}{2}\int_0^t f^2 ds\Big)$$

is a martingale with respect to $\{\mathfrak{Y}_t\}$ in $(\Omega, \mathfrak{B}, P)$. Therefore, we can similarly get the second part.

Lemma 4. *Under the same assumption as in Lemma 3, let $\mathfrak{M}_t$ be the linear manifold spanned by $\{Y_s;\, s\leqq t\}$ and let $\overline{\mathfrak{M}}_t^{(P)}$ and $\overline{\mathfrak{M}}_t^{(\tilde{P})}$ be the closure of $\mathfrak{M}_t$ by L^2-norm relative to the measure P and $\tilde{P}$, respectively. Then $\overline{\mathfrak{M}}_t^{(P)} = \overline{\mathfrak{M}}_t^{(\tilde{P})}$. Moreover,*

$$F_t(\omega) = \int_0^t f(s,\omega)ds$$

belongs to $\overline{\mathfrak{M}}_t^{(P)}$, where $f(s,\omega)$ is the function of Lemma 1.

Proof. To prove the first half, let $Z\in\overline{\mathfrak{M}}_t^{(P)}$ and $\{Z_n\}$ be a sequence of $\mathfrak{M}_t$ converging to Z in $L^2(P)$ sense. Then, there is a subsequence $\{Z_{n_k}\}$ of $\{Z_n\}$ converging to Z with P-measure 1. Since $\{Z_{n_k}\}$ is a Gaussian system relative to the measure $\tilde{P}$, so is $\{z_{n_k}\}\cup\{Z\}$, and the convergence $\{Z_{n_k}\}\to\{Z\}$ takes place in $L^2(\tilde{P})$-sense. Therefore, $Z\in\overline{\mathfrak{M}}_t^{(\tilde{P})}$ or equivalently $\overline{\mathfrak{M}}_t^{(P)}\subseteq\overline{\mathfrak{M}}_t^{(P)}$. The converse relation is carried out by the same way.

Since $(X_t, \mathfrak{Y}_t, P)$ is a martingale, the relation (3.2) implies that, for each $h>0$, the equality

$$\int_0^t \frac{E(F_{s+h}\,|\,\mathfrak{Y}_s)-F_s}{h}ds = \int_0^t \frac{E(Y_{s+h}\,|\,\mathfrak{Y}_s)-Y_s}{h}ds \quad \text{for any } t^{5)} \tag{3.3}$$

4) We say a stochastic process L_t, $t\in[0,T]$, is a local martingale with respect to $\{\mathfrak{Y}_t\}$ in $(\Omega, \mathfrak{B}, P)$ when there exists an increasing sequence of stopping times $\{T_n\}_{n=1,2,\dots}$ with respect to $\{\mathfrak{Y}_t\}$ such that $T_n\to T$ with P-measure 1 and $L_{t\wedge T_n}$ is a martingale with respect to $\{\mathfrak{Y}_t\}$ for each n in $(\Omega, \mathfrak{B}, P)$ (see [6]).

5) We define $F_t=F_T$ and $Y_t=Y_T$ for $t>T$, for convenient.

holds with P-measure 1. On the other hand, it is known that every $E[Y_{s+h}|\mathfrak{Y}_s]$, $s\leqq t$, belongs to $\mathfrak{M}_{h+t}^{(P)}$ because (Y_t, P) is Gaussian. Hence it is sufficient to show that the left hand of (3.3) converges to F_t in probability. Put

$$F_t^+ = \int_0^t (f\vee 0)ds, \quad F_t^- = -\int_0^t (f\wedge 0)ds$$

and denote by $F_{h,t}^{\pm}$ the left hand of (3.3) replacing F by $F^{\pm}$ there, respectively. Then $\{F_t^{\pm}\}$ is a continous and increasing process adapted to $\{\mathfrak{Y}_t\}$ in the sense of Meyer [8]. Moreover, $\{F_t^{\pm}\}$ is integrable by Lemma 3. Hence, $F_t^{\pm}$ converges to $F_t^{\pm}$ as $h\to 0$ in $L^1(P)$-sense ([8] p. 126), respectively. Similarly $F_{h,t}$ converges to F_t as $h\to 0$. Thus the proof is complete.

Proof of Theorem 1. 1°. Under the assumption of this theorem, we shall first prove that, with P-measure 1 the function $f(s, \omega)$ of Lemma 1 can be represented by

$$f(s, \omega) = \int_0^s k(s, u)dY_u \qquad \text{for almost all} \quad s\in[0, T],$$

where $k(s, u)$ is a Volterra kernel in $L^2([0, T]^2)$. It is well known that each element of $\overline{\mathfrak{M}}_t^{(P)}$ can be represented by a stochastic integral of the form $\int_0^t K(u)dY_u$. Hence F_t of Lemma 4 is represented by $\int_0^t K(t, u)\, dY_u$. Noting that F_t is continuous, we can choose $K(t, u)$ to be (t, u)-measurable by means of Slutsky's method [12]. Now, let Λ denote the (s, ω)-set

$$\{(s, \omega);\ \lim_{n\to\infty} n(F_s - F_{s-1/n}) \quad \text{does not exist or} \\ \lim_{n\to\infty} n(F_s - F_{s-1/n}) \neq f(s, \omega)\},$$

where we define $F_s=0$, for $s<0$. Then Λ is (s, ω)-measurable and $\mu(\Lambda)=0$, where μ is the product measure of $\tilde{P}(d\omega)$ and Lebesgue measure $m(ds)$ on $[0, T]$. In fact, $m(\Lambda_\omega)=0$ with P-measure 1, where $\Lambda_\omega=\{s;\ (s, \omega)\in\Lambda\}$. By Fubini's theorem it follows that $\tilde{P}(\tilde{\Lambda}_s)=0$ for almost all s, where $\Lambda_s=\{\omega;\ (s, \omega)\in\Lambda\}$. Therefore, for almost all s,

$$\lim_{n\to\infty} n(F_s - F_{s-1/n}) = f(s, \omega) \qquad \text{for} \quad \omega\notin\Lambda_s,$$

and $f(s, \omega)\in\overline{\mathfrak{M}}_s^{(P)}$ for such s. Hence for almost all $s\in[0, T]$,

$$f(s, \omega) = \int_0^s k(s, u)dY_u \qquad \text{with } P\text{-measure } 1,$$

where $k(s, u)$ belongs to $L^2(du)$ for such s. Moreover, we can choose $k(s, u)$ to be (s, u)-measurable. Put

$$\Lambda'_s = \left\{\omega ; f(s, \omega) \neq \int_0^s k(s, u) dY_u\right\} \cup \Lambda_s ,$$

then $P(\Lambda'_s)=0$ for almost all $s \in [0, T]$. Since

$$\Lambda' = \left\{(s, \omega); f(s, \omega) \neq \int_0^s k(s, u) dY_u\right\} \cup \Lambda$$

is (s, ω)-measurable,

$$f(s, \omega) = \int_0^s k(s, u) dY_u$$

holds for $s \in \Lambda_\omega' = \{s; (s, \omega) \in \Lambda'\}$ with P-measure 1. By this fact, we can get

$$X_t = Y_t - \int_0^t \Big(\int_0^s k(s, u) dY_s) ds .$$

By Lemma 2,

$$\begin{aligned}\tilde{E}\left(\int_0^T \left(\int_0^s k(s, u) dY_u \right)^2 ds \right) &= \int_0^T \tilde{E}\left(\int_0^s k(s, u) dY_u \right)^2 ds \\ &= \int_0^T \int_0^s k(s, u)^2 ds du < \infty ,\end{aligned}$$

therefore we can see that $k(s, u)$ is a Volterra kernel of $L^2([0, T]^2)$.

2°. Next, we want to represent Y_t in the form of (3.1), by constructing the kernel $l(s, t)$. For the Volterra kernel $k(s, t)$, there is a resolvent kernel $l(s, t)$ such that

$$\text{(3.4)} \qquad \begin{aligned} &l(s, t) + k(s, t) - \int_t^s l(s, u) k(u, t) du = 0 \qquad \text{in} \quad L^2([0, T]^2) \\ &l(s, t) + k(s, t) - \int_t^s k(s, u) l(u, t) du = 0 \qquad \text{in} \quad L^2([0, T]^2) . \end{aligned}$$

For, the Neumann series for the Volterra kernel $k(s, t)$ converges in the sense of $L^2([0, T]^2)$, and the limit is the kernel $l(s, t)$ (see Smithesis [13]). Thus the equations

$$\text{(3.5)} \qquad \begin{aligned} &X_t - \int_0^t \left(\int_0^s l(s, u) dX_u \right) ds \\ &= Y_t - \int_0^t \left(\int_0^s k(s, u) dY_u \right) ds - \int_0^t \left(\int_0^s l(s, u) dY_u \right) ds \\ &\qquad + \int_0^t \left(\int_0^s l(s, u) \int_0^u k(u, v) dY_v du \right) ds \\ &= Y_t - \int_0^t \left(\int_0^s k(s, u) dY_u \right) ds - \int_0^t \left(\int_0^s l(s, u) dY_u \right) ds \\ &\qquad + \int_0^t \left(\int_0^s \left(\int_v^s l(s, u) k(u, v) du \right) dY_v \right) ds \end{aligned}$$

hold with P-measure 1 for each $t\in[0, T]$. The last equality follows by using the formula

$$\int_0^s\left(\int_0^s m(u, v)dY_u\right)dv = \int_0^s\left(\int_0^s m(u, v)dv\right)dY_u$$

where m$(u, v)\in L^2[0, T]^2$). By the equation (3.4), the stochastic process

$$\int_0^t\left\{\int_0^s\left(-k(s, u)-l(s, u)+\int_u^s l(s, v)k(v, u)dv\right)dY_u\right\}ds$$

is identically 0 with P-measure 1. Therefore the right side of (3.5) is equal to Y_t with P-measure 1 for each $t\in[0, T]$.

This proves (3.1), for both side of (3.1) are continuous with P-measure 1.

3°. Finally we will discuss the uniqueness of the representation (3.1). In fact, we will prove a slightly stronger result than the uniqueness statement in the theorem. Suppose Y_t has two representations such as

$$\begin{aligned} Y_t &= X_t^1-\int_0^t\left(\int_0^s l(s, u)dX_u^1\right)ds \\ &= X_t^2-\int_0^t h(s, \omega)ds, \end{aligned}$$

where $(X_t^1, \mathfrak{Y}_t, P)$ and $(X_t^2, \mathfrak{Y}_t, P)$ are Wiener processes and $h(s, \omega)$ is a function satisfying (i), (ii) and (iii) in Lemma 1. Then,

$$X_t^1-X_t^2 = \int_0^t\left(h(s, \omega)-\int_0^s l(s, u)dX_u^1\right)ds$$

is a martingale with respect to $\{\mathfrak{Y}_t\}$. By the uniqueness of Meyer's decomposition ([7] p. 113), it follows easily that

$$X_t^1-X_t^2 = X_0^1-X_0^2 = 0,$$
$$\int_0^t h(s, \omega)ds = \int_0^t\left(\int_0^s l(s, u)dX_u^1\right)ds$$

hold for any $t\in[0, T]$ with P-measure 1. By Fubini's theorem

$$h(s, \omega) = \int_0^s l(s, u)dX_u^1, \qquad \text{for almost all} \quad (s, \omega)\in[0, T]\times\Omega.$$

Thus, the proof of this theorem is completed.

REMARK 3. Theorem 1 shows that $\mathfrak{X}_t=\mathfrak{Y}_t$ for each $t\in[0, T]$, where $\mathfrak{X}_t$ is the σ-algebra generated by $\{X_s; s\leqq t\}$ and P-negligible sets. Therefore (Y_t, P) has the proper canonical representation (3, 1) with respect to the Wiener process $(X_t, \mathfrak{X}_t, P)$ in the sense of Hida [4].

In case of $E(Y_t)\neq 0$, we get the following theorem by the same method as in Theorem 1.

Theorem 1′. *Suppose that a Gaussian process* (Y_t, P), $t \in [0, T]$, *is equivalent to Wiener process. Then there exists a Wiener process* (X_t, $\mathfrak{Y}_t$, P) *such that* Y_t *is represented, with P-measure* 1, *by the formula*

$$Y_t = X_t - \int_0^t \left(\int_0^s l(s, u) dX_u \right) ds - \int_0^t a(s) ds, \quad \text{for any } t \in [0, T]$$

where $l(s, u)$ *is a Volterra kernel belonging to* $L^2([0, T]^2)$ *and* $a(s) \in L^2([0, T])$. *Moreover such* X_t, $l(s, u)$ *and* $a(s)$ *are unique.*

4. Sufficient condition

In this section we will prove the converse of Theorem 1. This result is implied in Shepp [11] or H. Sato [10], if we note that the Volterra kernel has only zero as its eigenvalues. But, our proof is simpler than theirs and the density is explicitly evaluated.

Theorem 2. *If* (X_t, $\mathfrak{F}_t$, P) *is a Wiener process, then the Gaussian process with respect to the measure P*

$$Y_t = X_t - \int_0^t \left(\int_0^s l(s, u) dX_u \right) ds$$

is equivalent to Wiener process, where $l(s, u) \in L^2([0, T]^2)$ *is a Volterra kernel. In this case the density* $\varphi(\omega)$ *in Definition* 3 *can be taken as follows*[6];

$$(4.1) \qquad \varphi(\omega) = \exp\left\{ \int_0^T \left(\int_0^s l(s, u) dX_u \right) dX_s - \frac{1}{2} \int_0^T \left(\int_0^s l(s, u) dX_u \right)^2 ds \right\}.$$

Proof. This theorem is established if we can show that

$$E\left(\exp\left\{ \int_0^T \left(\int_0^s l(s, u) dX_u \right) dX_s - \frac{1}{2} \int_0^T \left(\int_0^s l(s, u) dX_u \right)^2 ds \right\} \right) = 1,$$

for then (Y_t, $\mathfrak{F}_t$, $N_T(\omega)P(d\omega)$) is a Wiener process by virtue of Girsanov [2]. Let

$$N_t = \exp\left\{ \int_0^t \left(\int_0^s l(s, u) dX_u \right) dX_s - \frac{1}{2} \int_0^t \left(\left(\int_0^s l(s, u) dX_u \right)^2 ds \right) \right\}.$$

Then the process (N_t, $\mathfrak{F}_t$, P) is a local martingale (see [6]) and we may consider N_t has continous paths. Therefore, there is an increasing sequence of stopping times $\{T_n\}$ which tends to T with P-measure 1 such that $\{N_t^n = N_{t \wedge T_n}, \mathfrak{F}_t, P\}$ is a martingale for each n. Hence, it is enough to prove that $\{N_T^n\}$ is uniformly integrable, because $E(N_T) = \lim_{n \to \infty} E(N_{T_n}) = 1$. Observing that

6) The density $\varphi(\omega)$ may not be unique in general. But, the function of (4.1) is the only density which is measurable with respect to $\{X_t;\ 0 \leqq t \leqq T\}$.

$$N_t^n = \exp\left\{\int_0^t \left(\chi_{[0,T_n]}(s)\int_0^s l(s, u)dX_u\right)dX_s - \frac{1}{2}\int_0^t \left(\chi_{[0,T_n]}(s)\int_0^s l(s, u)dX_u\right)^2 ds\right\},$$

where

$$\chi_{[0,T_n]}(s) = \begin{cases} 1 & \text{if } s \leqq T_n \\ 0 & \text{if } s > T_n, \end{cases}$$

define

$$\begin{aligned} Y_t^n &= X_t - \int_0^t \chi_{[0,T_n]}(s)\left(\int_0^s l(s, u)dX_u\right)ds \\ &= X_t - \int_0^{T_n \wedge t}\left(\int_0^s l(s, u)dX_u\right)ds. \end{aligned}$$

Then Girsanov's theorem, applied to N_T^n, tells us that $(Y_t^n, \mathfrak{F}_t, N_T^n P(d\omega))$ is a Wiener process for each n. On the other hand, a similar calculation as in (3.5) shows that

$$X_t = Y_t - \int_0^t \left(\int_0^s k(s, u)dY_u\right)ds$$

holds for any $t \in [0, T]$ with P-measure 1, where $k(s, u)$ is the resolvent kernel of $l(s, u)$ which satisfies (3.4). Therefore, it follows that

$$\int_0^t \left(\int_0^s k(s, u)dY_u\right)ds = \int_0^t \left(\int_0^s l(s, u)dX_u\right)ds, \qquad t \in [0, T],$$

or

$$\int_0^s k(s, u)dY_u = \int_0^s l(s, u)dX_u \qquad \text{for almost all} \quad s \in [0, T].$$

Then

$$\begin{aligned} \log N_t^n &= \int_0^{t\wedge T_n}\left(\int_0^s l(s, u)dX_u\right)dX_s - \frac{1}{2}\int_0^{t\wedge T_n}\left(\int_0^s l(s, u)dX_u\right)^2 ds \\ &= \int_0^{t\wedge T_n}\left(\int_0^s l(s, u)dX_u\right)dY_s + \frac{1}{2}\int_0^{t\wedge T_n}\left(\int_0^s l(s, u)dX_u\right)^2 ds \\ &= \int_0^{t\wedge T_n}\left(\int_0^s k(s,u)dY_u\right)dY_s + \frac{1}{2}\int_0^{t\wedge T_n}\left(\int_0^s k(s,u)dY_u\right)^2 ds \\ &= \int_0^{t\wedge T_n}\left(\int_0^s k(s, u)dY_u^n\right)dY_s^n + \frac{1}{2}\int_0^{t\wedge T_n}\left(\int_0^s k(s, u)dY_u^n\right)^2 ds, \end{aligned}$$

for any $t \in [0, T]$, with P-measure 1. The last equality follows from the fact that

$$\int_0^{t\wedge T_n} f(s, \omega)dY_s = \int_0^{t\wedge T_n} f(s, \omega)dY_s^n$$

holds for any $f(s, \omega)$ satisfying (i), (ii) and (iii) in Lemma 1. As $(Y_t^n, \mathfrak{F}_t, \tilde{P}^n)$, $n=1, 2, \cdots$, are Wiener processes, where $\tilde{P}^n(d\omega) = N_T^n(\omega)P(d\omega)$, we have

$$E((\log N_T^n)\, N_T^n) = \tilde{E}^n (\log N_T^n)$$
$$= \tilde{E}^n\Big(\int_0^{T\wedge T_n}\Big(\int_0^s k(s, u)dY_u^n\Big)dY_n^s + \frac{1}{2}\int_0^{T\wedge T_n}\Big(\int_0^s k(s, u)dY_u^n\Big)^2 ds\Big)$$
$$= \frac{1}{2}\tilde{E}^n\Big(\int_0^{T\wedge T_n}\Big(\int_0^s k(s, u)dY_u^n\Big)^2 ds\Big)$$
$$\leqq \frac{1}{2}\int_0^T\Big(\int_0^s k(s, u)^2 du\Big)ds = K.$$

Hence, the family $\{N_T^n\}_{n=1,2,\cdots}$ is uniformly integrable. This is our desired result.

More generally, the following theorem can be proved analogously.

Theorem 2′. *Let* $(X_t, \mathfrak{F}_t, P)$, $l(s, t)$ *be as in Theorem* 2 *and* $a(s)$ *be of* $L^2([0, T])$. *Then the Gaussian process with respect to the measure* P

$$Y_t = X_t - \int_0^t\Big(\int_0^s l(s, u)dX_u\Big)ds - \int_0^t a(s)ds$$

is equivalent to Wiener process. In this case, the function

$$\varphi(\omega) = \exp\Big\{\int_0^T\Big(\int_0^s l(s, u)dX_u + a(s)\Big)dX_s - \frac{1}{2}\int_0^T\Big(\int_0^s l(s, u)dX_u + a(s)\Big)^2 ds\Big\}$$

defines a density in Definition 3 *such that* $(X_t, \mathfrak{F}_t, \varphi P(d\omega))$ *is a Wiener process.*

5. Related topics

1. A decomposition of a positive definite operator $(I-H)$ on $L^2([0, T])$.

Proposition 1. *A Gaussain process* (Y_t, P), $t\in[0, T]$, *with mean* 0 *is equivalent to the Wiener process if and only if* (Y_t, P) *has the covariance*

$$E(Y_{t_1} Y_{t_2}) = (t_1\wedge t_2) - \int_0^{t_1\wedge t_2}\Big(\int_u^{t_1} l(s, u)ds\Big)du - \int_0^{t_1\wedge t_2}\Big(\int_u^{t_2} l(s, u)\,ds\Big)du + \int_0^{t_1}\int_0^{t_2}\Big(\int_0^{s_1\wedge s_2} l(s_1, u)\, l(s_2, u)du\Big)ds_1\, ds_2,$$

with a Volterra kernel $l(s, u)$ *in* $L^2([0, T]^2)$. *Moreover such* $l(s, u)$ *is unique.*

This proposition follows immediately from Theorem 1 and Theorem 2. As an application to L^2-theory, we can get

Proposition 2. *Let* H *be a symmetric integral operator on* $L^2([0, T])$.

Then $I-H$ is strictly positive definite if and only if there is an integral operator L of Volterra type such that

$$I\text{-}H=(I-L)(I-L^*)\,, \tag{5.1}$$

where L^* is the adjoint of L. Furthermore, such a decomposition is unique.

Proof. "If" part: Since L is of Volterra type, the integral equation

$$(I-L)f=0$$

has the unique solution $f=0$ in $L^2([0,\ T])$ (Simthesis [13]). Therefore

$$(\mathrm{I}-L^*)g=0$$

has the unique solution $g=0$ in $L^2([0,\ T])$. Hence, for $g\neq 0$,

$$\begin{aligned}((I-H)g,\ g)&=((I-L)(I-L^*)g,\ g)\\ &=((I-L^*)g,\ (I-L^*)g)>0\,.\end{aligned}$$

"Only if" part: Let $h(u,\ v)$ be the kernel which defines the operator H. Then, by a result of Shepp [11], there is a Gaussain process $(Y_t,\ P)$, equivalent to Wiener process, with covariance

$$E(Y_{t_1}\ Y_{t_2})=(t_1\wedge t_2)-\int_0^{t_1}\int_0^{t_2}h(u,\ v)\,du\,dv\,.$$

Hence, by Proposition 5.1, there is a unique Volterra kernel $l(u,\ v)$ such that

$$h(u,\ v)=l(u,\ v)+l(v,\ u)-\int_0^T l(u,\ w)\ l(u,\ w)\ dw\,.$$

If we define the operator $(I-L)$ by

$$(I-L)f(u)=f(u)-\int_0^T l(u,\ v)f(v)dv=f(u)-\int_0^u l(u,\ v)f(v)dv\,,$$

then

$$(I-L)\ (I-L^*)\,f(u)=(I-H)f(u).$$

2. Pinned Wiener process (Lévy [7] p. 318).
If $(X_t,\ \mathfrak{F}_t,\ P)$, $t\in[0,\ 1]$, is a Wiener process, then

$$\begin{aligned}Y_t&=(1-t)\int_0^t\frac{dX_u}{1-u}=X_t-\int_0^t\Big(\int_0^s\frac{-1}{1-u}dX_u\Big)ds \qquad 0\leqq t<1\\ &=0 \qquad\qquad t=1\end{aligned}$$

is the so-called pinned Wiener process with mean 0 and covariance

$E(Y_t Y_s)=(1-t)s$, for $t<s$. In this case,

$$l(s,u)=\begin{cases}\dfrac{-1}{1-u} & u\leqq s\\ 0 & u>s\end{cases},\quad k(s,u)=\begin{cases}\dfrac{1}{1-s} & u\leqq s\\ 0 & u>s\,.\end{cases}$$

Evidently, the Gaussian process (Y_t, P) is equivalent to Wiener process in $[0, t_0]$, $t_0<1$, by Theorem 2, while (Y_t, P) is not equivalent to Wiener process in $[0, 1]$, because $Y_1=0$, with P-measure 1. This phenomenon can be explained from that the kernel $l(s, u)$ does not belong to $L^2([0, 1]^2)$. The process Y_t is the unique solution of the stochastic integral equation

$$Y_t = X_t+\int_0^t \frac{1}{1-s}\int_0^s dY_u \qquad t<1\,,$$

with the initial condition $Y_0=0$.

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References

[1] J.L. Doob: Stochastic Processes, Wiley, New York, 1963.

[2] I.V. Girsanov: *On transforming a certain class of stochastic processes by absolutely continuous substitution of measures*, Ther. Probability Appl. **5** (1960), 285–301.

[3] Ju. I. Golosov: *Gaussian measures equivalent to Gaussian Markov measures*, Dokl. Akad. Nauk SSSR **166** (1966), 48–52.

[4] T. Hida: *Canonical representations of Gaussian processes and their applications*, Mem. Coll. Sci. Univ. Kyoto **33** (1960), 109–155.

[5] K. Itô: *On a formula concerning stochastic differentials*, Nagoya Math. J. **3** (1951), 55–65.

[6] H. Kunita and S. Watanabe: *On square integrable martingles*, Nagoya Math. J. **30** (1967), 209–245.

[7] P. Lévy: Processus Stochastiques et Mouvement Browinen, Gauthier-Villars, Paris, 1965.

[8] P.A. Meyer: Probability and Potentials, Blaisdell Publ. Comp., Waltham, 1966.

[9] Yu. A. Rozanov: *On the density of Gaussian measure with respect to another*, Theor. Probability Appl. **7** (1962), 82–87.

[10] H. Sato: *On the equivalence of Gaussian measures*, J. Math. Soc. Japan **19** (1967), 159–172.

[11] L.A. Shepp: *Randon-Nikodym derivatives of Gaussian measures*, Ann. Math. Statist. **37** (1966), 321–354.

[12] E. Slutsky: *Sur les fonctions aléatoires presque périodiques et sur la décomposition des fonctions aléatoires stationaires en composantes*, Actualités Sci. Indust. **738** (1938), 35–55.

[13] F. Smithies: *On the theory of linear integral equations*, Proc. Cambridge Philos. Soc. **31** (1935), 76–84.

Supplement for the proof of Lemma 1.

Professor T. Watanabe pointed out that, for the representation $M_t = \int_0^u g(s, \omega) dY_s + 1$, it is necessary to prove that there is an increasing sequence $\{T_n\}_{n=1,2,\cdots}$ of stopping times wh'ch converges to T and $\{M_{t\wedge T_n}\}_{t\in[0,T]}$ $(n=1, 2, \cdots)$ are square integrable martingales with respect to $\{\mathfrak{Y}_t\}$ in $(\Omega, \mathfrak{B}, P)$ (see Kunita-S. Watanabe [6]).

For the proof, we will first show that M_t has cotinuous paths. Set

$$M_t^N = \tilde{E}\left(\frac{1}{\varphi} \wedge N \mid \mathfrak{Y}_t\right),$$

then $M_t - M_t^N$ is a positive martingale and M_T^N converges to M_T in $L^1(\tilde{P})$ sense. Using Doob's inequality ([1] p. 353),

$$\tilde{P}(\sup_{0 \leqq t \leqq T} (M_t - M_t^N) \geqq \lambda) \leqq \frac{\tilde{E}(M_T - M_T^N)}{\lambda}.$$

This shows M_t^N converges to M_t uniformly in probability $\tilde{P}$. On the other hand, $\{M_t^N\}_{t\in[0,T]}$ are square integrable martingale, and so they have continuous paths. Hense M_t has continuous paths.

Next, if we choose the sequence of stopping times $\{T_n\}_{n=1,2\cdots}$ such that

$$T_n = \begin{cases} \min\{t;\, M_t \geqq n\} \\ T \qquad \text{if} \quad \{t;\, M_t \geqq n\} \neq \phi, \end{cases}$$

then T_n converges to T and $\{M_{t\wedge T_n}\}_{t\in[0,T]}$ $(n=1, 2\cdots)$ are square integrable martingales, because of the cotinuity of paths of M_t.