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**ON A RELATION BETWEEN HIGHER ORDER
 ASYMPTOTIC RISK SUFFICIENCY AND HIGHER ORDER
 ASYMPTOTIC SUFFICIENCY IN A LOCAL SENSE**

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1. Introduction. In Takeuchi [4] higher order asymptotic risk sufficiency of maximum likelihood estimator has been discussed. In this paper we try to find some relations between asymptotic risk sufficiency with a special loss function and asymptotic sufficiency in a local sense.

Let $\mathcal{P}_n = \{P_{\theta,n}; \theta \in \Theta\}$ be a family of probability distributions on a measurable space $(\mathcal{X}, \mathcal{A}_n)$ with an index set Θ which is a subset of an Euclidean space with the usual norm $|\cdot|$. For a sub σ -field \mathcal{C} of \mathcal{A}_n , real number $c \geq 0$ and $\theta, \theta' \in \Theta$ let $r_n^{\mathcal{C}}(c: \theta, \theta') = \inf \{ (1+c)^{-1} \{ 1 - E_{P_{\theta,n}}(\phi) + c E_{P_{\theta',n}}(\phi); \phi \text{ are } \mathcal{C}\text{-measurable statistical test functions on } \mathcal{X} \} \}$. We note that $r_n^{\mathcal{C}}(c: \theta, \theta')$ means the Bayes risk of statistical problem of testing a hypothesis ' $P_{\theta',n}$ is true' against an alternative ' $P_{\theta,n}$ is true' with experiment $(\mathcal{X}, \mathcal{C}, \{P_{\theta',n}, P_{\theta,n}\})$ relative to a prior probability distribution $(c/(1+c), 1/(1+c))$ on $\{\theta', \theta\}$ provided that the loss function is simple.

Let $\{\mathcal{B}_n; n=1, 2, \dots\}$ be a sequence of sub σ -fields of $\{\mathcal{A}_n\}$ ($\mathcal{B}_n \subset \mathcal{A}_n$). In this paper we give a sufficient condition about the Bayes risk $r_n^{\mathcal{B}_n}$ for $\{\mathcal{B}_n\}$ to be higher order locally asymptotically sufficient sequence of σ -fields. More precisely our main result in this paper is the following: Under some conditions if for some positive number α $\sup_{c>0} \sup_{\theta^* \in K} \sup_{\theta: n^{1/2}|\theta - \theta^*| \leq b} \{r_n^{\mathcal{B}_n}(c: \theta, \theta^*) - r_n^{\mathcal{A}_n}(c: \theta, \theta^*)\} = o(n^{-\alpha})$ for every $b > 0$ and every compact subset K of Θ , then for every β satisfying $0 < \beta < 3^{-1}\alpha$ $\{\mathcal{B}_n\}$ is locally asymptotically sufficient for $\{\mathcal{P}_n\}$ with order $o(n^{-\beta})$ in the sense that for each $n=1, 2, \dots$ and each $\theta_0 \in \Theta$ there exists a family $\{Q_{\theta_0,n}^{\theta_0}; \theta \in \Theta\}$ of probability distributions on $(\mathcal{X}, \mathcal{A}_n)$ for which \mathcal{P}_n is sufficient σ -field and that for every $b > 0$

$$\sup_{\theta: n^{1/2}|\theta - \theta_0| \leq b} \|P_{\theta,n} - Q_{\theta_0,n}^{\theta_0}\|_{\mathcal{A}_n} = o(n^{-\beta})$$

uniformly in θ_0 over every compact subsets of Θ . Here $\|\cdot\|_{\mathcal{A}_n}$ means the total variation norm over \mathcal{A}_n .

We have discussed such a problem in the case $\alpha = \beta = 0$ in Suzuki [3] under

non-local situation. In LeCam [1], Chap. 5 he discusses some relations between insufficiency and deficiency in his terminology.

In Section 2 some auxiliary results about the order of asymptotic sufficiency are proved. The main theorem is stated and followed by some discussions about the asymptotic sufficiency in non-local sense in Section 3.

2. Auxiliary results. For each $n \in N = \{1, 2, \dots\}$ let $\mathcal{P}_n = \{P_{\theta, n}; \theta \in \Theta\}$ be a family of probability distributions on a measurable space $(\mathcal{X}, \mathcal{A}_n)$ with an index set Θ . For a subset $U(\neq \emptyset)$ of Θ we shall denote by \mathcal{P}_n^U the totality of $P_{\theta, n}$'s satisfying $\theta \in U$. We assume that for each $n \in N$ \mathcal{P}_n is dominated by a σ -finite measure μ_n on $(\mathcal{X}, \mathcal{A}_n)$. The probability density function of $P_{\theta, n}$ relative to μ_n will be denoted by $p_n(x, \theta)$. Without loss of generality we assume in the following that μ_n is a probability measure on $(\mathcal{X}, \mathcal{A}_n)$. For each $\theta, \theta' \in \Theta$ let $S_n(\theta) = \{x; p_n(x, \theta) > 0\}$ and let $h_n(x; \theta, \theta') = p_n(x, \theta)/p_n(x, \theta')$ if $x \in S_n(\theta')$, $= +\infty$ if $x \in S_n(\theta) \cap S_n(\theta')^c$, $= 1$ if $x \in S_n(\theta)^c \cap S_n(\theta')^c$. We put $\beta_n(\theta, \theta') = P_{\theta, n}\{S_n(\theta')^c\}$. For each $\theta, \theta' \in \Theta$ and real number $s \geq 1$ we define

$$J_n(s; \theta, \theta') = E_{P_{\theta', n}}[\{h_n(x; \theta, \theta')\}^s].$$

We note that $\beta_n(\theta, \theta') = 1 - J_n(1; \theta, \theta')$.

Let $\{U_n\}$ be a sequence of nonempty subsets of Θ . For $\{U_n\}$ we consider the following assumption.

ASSUMPTION 1. There exist a sequence $\{\theta_n^*\}_{n \in N} (\theta_n^* \in U_n)$ and a positive number γ such that

(a) For every $s \geq 1$

$$\limsup_{b \rightarrow \infty} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*) < \infty,$$

(b) $\sup_{\theta \in U_n} \beta_n(\theta, \theta_n^*) = o(n^{-\gamma})$.

For a sub σ -field \mathcal{C} of \mathcal{A}_n we denote by $\Phi(\mathcal{C})$ the family of \mathcal{C} -measurable statistical test functions on \mathcal{X} . For each $\theta, \theta' \in \Theta$ and each real number $c \geq 0$ we define

$$r_n^{\mathcal{C}}(c; \theta, \theta') = \inf (1+c)^{-1} \{1 - E_{P_{\theta, n}}(\phi) + c E_{P_{\theta', n}}(\phi); \phi \in \Phi(\mathcal{C})\}.$$

Let $\{\mathcal{B}_n\}$ be a sequence of sub σ -fields of $\{\mathcal{A}_n\}$ ($\mathcal{B}_n \subset \mathcal{A}_n$). For each $\theta \in \Theta$ define $\bar{p}_n(x, \theta) = E_{\mu_n}[p_n(x, \theta) | \mathcal{B}_n]$ the conditional expectation of $p_n(x, \theta)$ given \mathcal{B}_n with respect to μ_n and put $S'_n(\theta) = \{x; \bar{p}_n(x, \theta) > 0\}$. For $\theta, \theta' \in \Theta$ define $g_n(x; \theta, \theta') = \bar{p}_n(x, \theta)/\bar{p}_n(x, \theta')$ if $x \in S'_n(\theta')$, $= +\infty$ if $x \in S'_n(\theta')^c \cap S'_n(\theta)$, $= 1$ if $x \in S'_n(\theta')^c \cap S'_n(\theta)$. For $c > 0$ and $\delta > 0$ let $E_n(c, \theta, \delta) = \{x; g_n(x; \theta, \theta_n^*) < c < c + \delta \leq h_n(x; \theta, \theta_n^*)\}$ and $E'_n(c, \theta, \delta) = \{x; g_n(x; \theta, \theta_n^*) > c > c - \delta > h_n(x; \theta, \theta_n^*)\}$.

Proposition. Suppose that for some positive number α and a sequence $\{\theta_n^*\}_{n \in N}$ ($\theta_n^* \in U_n$)

$$(2.1) \quad \sup_{c>0} \sup_{\theta \in U_n} \{r_n^{\mathcal{B}_n}(c; \theta, \theta_n^*) - r_n^{\mathcal{A}_n}(c; \theta, \theta_n^*)\} = o(n^{-\alpha}).$$

Then we have

$$(2.2) \quad \begin{aligned} \sup_{c>0, \delta>0} \delta(1+c)^{-1} \lambda_n(c, \delta) &= o(n^{-\alpha}), \quad \text{and} \\ \sup_{c>0, \delta>0} \delta(1+c)^{-1} \lambda'_n(c, \delta) &= o(n^{-\alpha}) \end{aligned}$$

where $\lambda_n(c, \delta) = \sup_{\theta \in U_n} P_{\theta_n^*, n}(E_n(c, \theta, \delta))$ and $\lambda'_n(c, \delta) = \sup_{\theta \in U_n} P_{\theta_n^*, n}(E'_n(c, \theta, \delta))$.

This proposition can be proved in the same way as the proof of the first and second steps of Theorem 1 in Suzuki [3]. So we shall omit the proof of the proposition.

Theorem 1. Suppose that Assumption 1 is satisfied with a sequence $\{\theta_n^*\}_{n \in N}$ and $\gamma > 0$, and that $\{\mathcal{B}_n\}$ has the property (2.1) with $\beta > 0$. Then for every β satisfying $0 < \beta < 3^{-1}\alpha$ and $\beta \leq \gamma$, $\{\mathcal{B}_n\}$ is asymptotically sufficient for $\{\mathcal{P}_n^U\}$ with order $o(n^{-\beta})$ in the following sense: For each $n \in N$ there exists a family $\{q_n(x; \theta, \theta_n^*); \theta \in \Theta\}$ of probability density functions on $(\mathcal{X}, \mathcal{A}_n)$ relative to μ_n such that

(i) each q_n can be factorized as follows:

$$q_n(x; \theta, \theta_n^*) = r_n(x; \theta, \theta_n^*) p_n(x, \theta_n^*)$$

where r_n is a \mathcal{B}_n -measurable function, and

$$(ii) \quad \sup_{\theta \in U_n} \int_{\mathcal{X}} |p_n(x, \theta) - q_n(x; \theta, \theta_n^*)| d\mu_n = o(n^{-\beta}).$$

Proof. We shall divide the proof into several steps.

The first step. Suppose that Assumption 1 is satisfied with a sequence $\{\theta_n^*\}_{n \in N}$ and $\gamma > 0$, and that $\{\mathcal{B}_n\}_{n \in N}$ has the property (2.1) with $\alpha > 0$. Let β be any number satisfying $0 < \beta < 3^{-1}\alpha$ and $\beta \leq \gamma$. Take ε_1 be any number satisfying $0 < \varepsilon_1 < 3^{-1} \cdot (\alpha - 3\beta)$. Let $\alpha_n = n^{-\beta} (\log n)^{-1}$, $m_n = n^{\varepsilon_1}$ and $i_n = [m_n \alpha_n^{-1}] + 1$ where $[a]$ means the maximum integer not exceeding a . Put $(\gamma_n =) \gamma_n(x; \theta, \theta_n^*) = |h_n(x; \theta, \theta_n^*) - g_n(x; \theta, \theta_n^*)|$, $(\gamma'_n =) \gamma'_n(x; \theta, \theta_n^*) = |h_n(x; \theta, \theta_n^*) - I_{W_n}(x) g_n(x; \theta, \theta_n^*)|$ and

$$\rho_n(\theta, \theta_n^*) = \int_{\mathcal{X}} \gamma'_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n}$$

where $W_n = W_n(\theta, \theta_n^*) = \{x; g_n(x; \theta, \theta_n^*) \leq m_n\}$ and I_{W_n} means the indicator function of W_n .

We have

$$\begin{aligned}
\sup_{\theta \in U_n} \rho_n(\theta, \theta_n^*) &\leq \sup_{\theta \in U_n} \int_{W_n} \gamma_n dP_{\theta_n^*, n} + \sup_{\theta \in U_n} \int_{W_n^c} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} = J_n^* + J_n^{**}, \\
(2.3) \quad \text{and} \quad J_n^* &= \sup_{\theta \in U_n} \int_{W_n} \gamma_n dP_{\theta_n^*, n} \leq \alpha_n + \sup_{\theta \in U_n} \int_{D_n \cap W_n} \gamma_n dP_{\theta_n^*, n} = \alpha_n + I_n \\
(D_n &= \{x; \gamma_n \geq \alpha_n\}).
\end{aligned}$$

Furthermore we have

$$I_n = \sup_{\theta \in U_n} \int_{D_n \cap W_n} \gamma_n dP_{\theta_n^*, n} \leq \sup_{\theta \in U_n} \int_{D_n \cap \tilde{W}_n} \gamma_n dP_{\theta_n^*, n} + \sup_{\theta \in U_n} \int_{W_n^*} \gamma_n dP_{\theta_n^*, n}$$

where $W_n' = \{x; h_n(x; \theta, \theta_n^*) \leq m_n\}$, $\tilde{W}_n = W_n \cap W_n'$ and $W_n^* = W_n \cap (W_n')^c$.

The second step. It holds that

$$\begin{aligned}
I_n' &= \sup_{\theta \in U_n} \int_{D_n \cap \tilde{W}_n} \gamma_n dP_{\theta_n^*, n} \leq \sum_{i=1}^{2i_n-2} \sup_{\theta \in U_n} \int_{B_i} \gamma_n dP_{\theta_n^*, n} \\
(2.4) \quad &+ \sum_{i=0}^{2i_n-3} \sup_{\theta \in U_n} \int_{C_i} \gamma_n dP_{\theta_n^*, n} \\
&= I_{n,1}' + I_{n,2}'
\end{aligned}$$

where $B_i = \tilde{W}_n \cap \{x; h_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+1) \alpha_n, g_n(x; \theta, \theta_n^*) < 2^{-1}i \alpha_n\}$ and $C_i = \tilde{W}_n \cap \{x; h_n(x; \theta, \theta_n^*) < 2^{-1}(i+1) \alpha_n, g_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+2) \alpha_n\}$. Using the property (2.2) in Proposition we can evaluate $I_{n,i}' (i=1, 2)$ as follows. Taking account of $3\varepsilon_1 < \alpha - 3\beta$ we have

$$\begin{aligned}
I_{n,1}' &= \sum_{i=1}^{2i_n-2} \sup_{\theta \in U_n} \int_{B_i} \gamma_n dP_{\theta_n^*, n} \\
&\leq 2i_n m_n \left[\sup_{1 \leq i \leq 2i_n-2} \sup_{\theta \in U_n} P_{\theta_n^*, n} \{x; h_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+1) \alpha_n, \right. \\
&\quad \left. g_n(x; \theta, \theta_n^*) < 2^{-1}i \alpha_n\} \right] \\
(2.5) \quad &\leq 2i_n m_n \left[\sup_{1 \leq i \leq 2i_n-2} \lambda_n (2^{-1}i \alpha_n, 2^{-1} \alpha_n) \right] \\
&\leq 4i_n m_n \left[\sup_{1 \leq i \leq 2i_n-2} \alpha_n^{-1} (1 + 2^{-1}i \alpha_n) n^{-\alpha} \eta'_n \right] (\eta'_n = o(1)) \\
&\leq 4i_n^2 m_n n^{-\alpha} \eta'_n \\
&\leq A_1 \cdot n^{-(\alpha-2\beta-3\varepsilon_1)} (\log n)^2 \eta'_n \quad (A_1 \text{ is a constant}) \\
&= o(n^{-\beta}).
\end{aligned}$$

Similarly we have

$$(2.6) \quad I_{n,2}' = o(n^{-\beta}).$$

Thus from (2.4) and (2.5) we have

$$(2.7) \quad I_n' = o(n^{-\beta}).$$

The third step. Next we evaluate I_n'' as follows. For every $s > 1$ we have

$$\begin{aligned} I_n'' &= \sup_{\theta \in U_n} \int_{W_n''} \gamma_n dP_{\theta_n^*, n} \leq \sup_{\theta \in U_n} \int_{\{h_n > m_n\}} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ &\leq (m_n)^{1-s} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*). \end{aligned}$$

Hence we have

$$I_n'' \leq A_2(s) (m_n)^{1-s} = A_2(s) n^{(1-s)\epsilon_1}$$

where $A_2(s)$ is some constant depending only on s . We can choose $s > 1$ large enough so that

$$(2.8) \quad I_n'' = o(n^{-\beta}).$$

From (2.7) and (2.8) we have

$$I_n = o(n^{-\beta}).$$

Hence from (2.3) we have

$$J_n^* = o(n^{-\beta}).$$

Put $W_n'' = \{x; h_n(x; \theta, \theta_n^*) < 2^{-1} m_n\}$. Then we have

$$\begin{aligned} (2.9) \quad J_n^{**} &= \sup_{\theta \in U_n} \int_{W_n''} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ &\leq \sup_{\theta \in U_n} \int_{W_n^c \cap W_n''} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} + \sup_{\theta \in U_n} \int_{W_n^c \cap (W_n'')^c} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ &\leq 2^{-1} m_n \lambda_n'(m_n, m_n/2) + (m_n/2)^{1-s} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*). \end{aligned}$$

The first term on the right hand side is of order $o(n^{-\beta})$ by Proposition. The similar consideration as the evaluation of I_n'' implies that the second term of (2.9) is also of order $o(n^{-\beta})$ for sufficiently large number s . Thus we have

$$J_n^{**} = o(n^{-\beta}).$$

Hence it follows from (2.3) that

$$(2.10) \quad \sup_{\theta \in U_n} \rho_n(\theta, \theta_n^*) = o(n^{-\beta}).$$

The fourth step. Let $a_n(\theta, \theta_n^*) = [\int_{\mathcal{X}} I_{W_n}(x) g_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n}]^{-1} (\leq \infty)$ and let $r_n(x; \theta, \theta_n^*) = a_n(\theta, \theta_n^*) I_{W_n}(x) g_n(x; \theta, \theta_n^*)$ if $a_n(\theta, \theta_n^*) < \infty$, $= 1$ otherwise. Define $q_n(x; \theta, \theta_n^*) = r_n(x; \theta, \theta_n^*) p_n(x, \theta_n^*)$ and let $\mathcal{Q}_{\theta_n^*, n}$ be the probability distribution on $(\mathcal{X}, \mathcal{A}_n)$ with density $q_n(x; \theta, \theta_n^*)$ relative to μ_n . We note that \mathcal{D}_n is suffi-

cient σ -field for the family $\{Q_{\theta, n}^{\theta*}; \theta \in \Theta\}$ by the factorization theorem. It follows from (2.10) that there exists n_0 such that $a_n(\theta, \theta_n^*) < \infty$ for every $n \geq n_0$ and every $\theta \in U_n$. Therefore we can assume without loss of generality that $a_n(\theta, \theta_n^*) < \infty$ for every $\theta \in U_n$ and every $n \geq 1$.

Under this circumstances we have

$$\begin{aligned} \|P_{\theta, n} - Q_{\theta, n}^{\theta*}\|_{\mathcal{A}_n} &= \int_{\mathcal{X}} |p_n(x, \theta) - q_n(x; \theta, \theta_n^*)| d\mu_n \\ &\leq \int_{S_n(\theta_n^*)} |h_n(x; \theta, \theta_n^*) - a_n(\theta, \theta_n^*) I_{W_n}(x) g_n(x; \theta, \theta_n^*)| \\ &\quad \cdot p_n(x, \theta_n^*) d\mu_n + \beta_n(\theta, \theta_n^*) \\ &= \rho_n(\theta, \theta_n^*) + |1 - a_n(\theta, \theta_n^*)^{-1}| + \beta_n(\theta, \theta_n^*) \\ &\leq 2 \rho_n(\theta, \theta_n^*) + 2 \beta_n(\theta, \theta_n^*). \end{aligned}$$

Here $\|\nu\|_{\mathcal{A}_n}$ means the total variation norm of a signed measure ν on $(\mathcal{X}, \mathcal{A}_n)$. From Assumption 1, (b) and (2.10) we have

$$\sup_{\theta \in U_n} \|P_{\theta, n} - Q_{\theta, n}^{\theta*}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

This completes the proof of the theorem.

3. The order of local asymptotic sufficiency. In this section the index set Θ is assumed to be a subset of p -dimensional Euclidean space R^p . We denote by $|\cdot|$ the usual Euclidean norm in R^p . For $\theta \in \Theta$ and $b > 0$ let $U_n(\theta, b) = \{\theta' \in \Theta; n^{1/2}|\theta' - \theta| \leq b\}$.

Let $\{\mathcal{B}_n\}_{n \in \mathbb{N}}$ be the sequence of sub σ -fields $\mathcal{B}_n \subset \mathcal{A}_n$ as in the previous section. We consider the following assumption which will be used to prove our main theorem, Theorem 2.

ASSUMPTION 2. For every compact subset K of Θ and $b > 0$

- (a) $\limsup_{n \rightarrow \infty} \sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} J_n(s; \theta, \theta^*) < \infty \quad (\forall s > 1), \text{ and}$
- (b) $\sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} \beta_n(\theta, \theta^*) = o(n^{-\gamma}).$

Let α be a given positive number. We state a result about higher order locally asymptotic sufficiency of $\{\mathcal{B}_n\}$ for $\{\mathcal{P}_n\}$.

Theorem 2. Suppose that Assumption 2 is satisfied with $\gamma > 0$, and that for every compact subset K of Θ and every $b > 0$

$$(3.1) \quad \sup_{c > 0} \sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} \{r_n^{\mathcal{B}_n}(c; \theta, \theta^*) - r_n^{\mathcal{A}_n}(c; \theta, \theta^*)\} = o(n^{-\alpha}).$$

Then for every positive number β satisfying $\beta < 3^{-1}\alpha$ and $\beta \leq \gamma$ $\{\mathcal{B}_n\}_{n \in \mathbb{N}}$ is locally asymptotically sufficient for $\{\mathcal{P}_n\}$ with order $o(n^{-\beta})$ in the following sense: For each

$n \in N$ and each $\theta_0 \in \Theta$ there exists a family $\mathcal{Q}_{n^0} = \{Q_{\theta,n}^{\theta_0}; \theta \in \Theta\}$ of probability distributions on $(\mathcal{X}, \mathcal{A}_n)$ such that

- (i) \mathcal{B}_n is sufficient for \mathcal{Q}_{n^0} , and
- (ii) for every compact subset K of Θ and every $b > 0$

$$\sup_{\theta \in K} \sup_{\theta \in U_n(\theta, b)} \|P_{\theta,n} - Q_{\theta,n}^{\theta_0}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

Since the above result follows directly from Theorem 1 we shall omit the proof.

It is open problem whether non-local version of Theorem 2 still holds or not, i.e., whether any conditions such as in Theorem 2 imply the followings or not: There exists a sequence $\mathcal{Q}_n = \{Q_{\theta,n}; \theta \in \Theta\}$ of probability distributions on $(\mathcal{X}, \mathcal{A}_n)$ such that \mathcal{B}_n is sufficient for \mathcal{Q}_n , and that for every compact subset K of Θ

$$(3.2) \quad \sup_{\theta \in K} \|P_{\theta,n} - Q_{\theta,n}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

The case of $\alpha = \beta = 0$ has been discussed in Suzuki [3] in such a non-local situation.

It is well known that under some regularity conditions there exist a sequence $\{\hat{\theta}_n\}_{n \in N}$ of estimators of θ , a positive number γ and a number $v \geq 1$ having the following property: For every compact subset K of Θ there corresponds $a(K)$ such that

$$\sup_{\theta \in K} P_{\theta,n} \{n^{1/2} |\hat{\theta}_n(x) - \theta| \geq a(K) (\log n)^{v/2}\} = o(n^{-\gamma})$$

(c.f. Matsuda [2], Chap. 3).

Using such an estimator $\{\hat{\theta}_n\}$ we may be able to construct $\{Q_{\theta,n}; \theta \in \Theta\}$ satisfying the property (3.2), and for which \mathcal{B}_n is sufficient.

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