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# ON A RELATION BETWEEN HIGHER ORDER ASYMPTOTIC RISK SUFFICIENCY AND HIGHER ORDER ASYMPTOTIC SUFFICIENCY IN A LOCAL SENSE

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**1. Introduction.** In Takeuchi [4] higher order asymptotic risk sufficiency of maximum likelihood estimator has been discussed. In this paper we try to find some relations between asymptotic risk sufficiency with a special loss function and asymptotic sufficiency in a local sense.

Let  $\mathcal{P}_n = \{P_{\theta,n}; \theta \in \Theta\}$  be a family of probability distributions on a measurable space  $(\mathcal{X}, \mathcal{A}_n)$  with an index set  $\Theta$  which is a subset of an Euclidean space with the usual norm  $|\cdot|$ . For a sub  $\sigma$ -field  $\mathcal{C}$  of  $\mathcal{A}_n$ , real number  $c \geq 0$  and  $\theta, \theta' \in \Theta$  let  $r_n^{\mathcal{C}}(c; \theta, \theta') = \inf (1+c)^{-1} \{1 - E_{P_{\theta,n}}(\phi) + cE_{P_{\theta',n}}(\phi)\}$ ;  $\phi$  are  $\mathcal{C}$ -measurable statistical test functions on  $\mathcal{X}$ . We note that  $r_n^{\mathcal{C}}(c; \theta, \theta')$  means the Bayes risk of statistical problem of testing a hypothesis ' $P_{\theta',n}$  is true' against an alternative ' $P_{\theta,n}$  is true' with experiment  $(\mathcal{X}, \mathcal{C}, \{P_{\theta',n}, P_{\theta,n}\})$  relative to a prior probability distribution  $(c/(1+c), 1/(1+c))$  on  $\{\theta', \theta\}$  provided that the loss function is simple.

Let  $\{\mathcal{B}_n; n=1, 2, \dots\}$  be a sequence of sub  $\sigma$ -fields of  $\{\mathcal{A}_n\}$  ( $\mathcal{B}_n \subset \mathcal{A}_n$ ). In this paper we give a sufficient condition about the Bayes risk  $r_n^{\mathcal{B}_n}$  for  $\{\mathcal{B}_n\}$  to be higher order locally asymptotically sufficient sequence of  $\sigma$ -fields. More precisely our main result in this paper is the following: Under some conditions if for some positive number  $\alpha$   $\sup_{c>0} \sup_{\theta^* \in K} \sup_{\theta: n^{1/2}|\theta-\theta^*| \leq b} \{r_n^{\mathcal{B}_n}(c; \theta, \theta^*) - r_n^{\mathcal{A}_n}(c; \theta, \theta^*)\} = o(n^{-\alpha})$  for every  $b>0$  and every compact subset  $K$  of  $\Theta$ , then for every  $\beta$  satisfying  $0 < \beta < 3^{-1}\alpha$   $\{\mathcal{B}_n\}$  is locally asymptotically sufficient for  $\{\mathcal{P}_n\}$  with order  $o(n^{-\beta})$  in the sense that for each  $n=1, 2, \dots$  and each  $\theta_0 \in \Theta$  there exists a family  $\{Q_{\theta,n}^0; \theta \in \Theta\}$  of probability distributions on  $(\mathcal{X}, \mathcal{A}_n)$  for which  $\mathcal{P}_n$  is sufficient  $\sigma$ -field and that for every  $b>0$

$$\sup_{\theta: n^{1/2}|\theta-\theta_0| \leq b} \|P_{\theta,n} - Q_{\theta,n}^0\|_{\mathcal{A}_n} = o(n^{-\beta})$$

uniformly in  $\theta_0$  over every compact subsets of  $\Theta$ . Here  $\|\cdot\|_{\mathcal{A}_n}$  means the total variation norm over  $\mathcal{A}_n$ .

We have discussed such a problem in the case  $\alpha=\beta=0$  in Suzuki [3] under

non-local situation. In LeCam [1], Chap. 5 he discusses some relations between insufficiency and deficiency in his terminology.

In Section 2 some auxiliary results about the order of asymptotic sufficiency are proved. The main theorem is stated and followed by some discussions about the asymptotic sufficiency in non-local sense in Section 3.

**2. Auxiliary results.** For each  $n \in N = \{1, 2, \dots\}$  let  $\mathcal{P}_n = \{P_{\theta,n}; \theta \in \Theta\}$  be a family of probability distributions on a measurable space  $(\mathcal{X}, \mathcal{A}_n)$  with an index set  $\Theta$ . For a subset  $U (\neq \emptyset)$  of  $\Theta$  we shall denote by  $\mathcal{P}_n^U$  the totality of  $P_{\theta,n}$ 's satisfying  $\theta \in U$ . We assume that for each  $n \in N$   $\mathcal{P}_n$  is dominated by a  $\sigma$ -finite measure  $\mu_n$  on  $(\mathcal{X}, \mathcal{A}_n)$ . The probability density function of  $P_{\theta,n}$  relative to  $\mu_n$  will be denoted by  $p_n(x, \theta)$ . Without loss of generality we assume in the following that  $\mu_n$  is a probability measure on  $(\mathcal{X}, \mathcal{A}_n)$ . For each  $\theta, \theta' \in \Theta$  let  $S_n(\theta) = \{x; p_n(x, \theta) > 0\}$  and let  $h_n(x; \theta, \theta') = p_n(x, \theta)/p_n(x, \theta')$  if  $x \in S_n(\theta')$ ,  $= +\infty$  if  $x \in S_n(\theta) \cap S_n(\theta')^c$ ,  $= 1$  if  $x \in S_n(\theta)^c \cap S_n(\theta')^c$ . We put  $\beta_n(\theta, \theta') = P_{\theta,n}\{S_n(\theta')^c\}$ . For each  $\theta, \theta' \in \Theta$  and real number  $s \geq 1$  we define

$$J_n(s; \theta, \theta') = E_{P_{\theta',n}}[\{h_n(x; \theta, \theta')\}^s].$$

We note that  $\beta_n(\theta, \theta') = 1 - J_n(1; \theta, \theta')$ .

Let  $\{U_n\}$  be a sequence of nonempty subsets of  $\Theta$ . For  $\{U_n\}$  we consider the following assumption.

**ASSUMPTION 1.** There exist a sequence  $\{\theta_n^*\}_{n \in N} (\theta_n^* \in U_n)$  and a positive number  $\gamma$  such that

- (a) For every  $s \geq 1$ 

$$\lim_{b \rightarrow \infty} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*) < \infty,$$
- (b)  $\sup_{\theta \in U_n} \beta_n(\theta, \theta_n^*) = o(n^{-\gamma}).$

For a sub  $\sigma$ -field  $\mathcal{C}$  of  $\mathcal{A}_n$  we denote by  $\Phi(\mathcal{C})$  the family of  $\mathcal{C}$ -measurable statistical test functions on  $\mathcal{X}$ . For each  $\theta, \theta' \in \Theta$  and each real number  $c \geq 0$  we define

$$r_n^{\mathcal{C}}(c; \theta, \theta') = \inf (1+c)^{-1} \{1 - E_{P_{\theta,n}}(\phi) + c E_{P_{\theta',n}}(\phi); \phi \in \Phi(\mathcal{C})\}.$$

Let  $\{\mathcal{B}_n\}$  be a sequence of sub  $\sigma$ -fields of  $\{\mathcal{A}_n\}$  ( $\mathcal{B}_n \subset \mathcal{A}_n$ ). For each  $\theta \in \Theta$  define  $\bar{p}_n(x, \theta) = E_{\mu_n}[p_n(x, \theta) | \mathcal{B}_n]$  the conditional expectation of  $p_n(x, \theta)$  given  $\mathcal{B}_n$  with respect to  $\mu_n$  and put  $S'_n(\theta) = \{x; \bar{p}_n(x, \theta) > 0\}$ . For  $\theta, \theta' \in \Theta$  define  $g_n(x; \theta, \theta') = \bar{p}_n(x, \theta)/\bar{p}_n(x, \theta')$  if  $x \in S'_n(\theta')$ ,  $= +\infty$  if  $x \in S'_n(\theta')^c \cap S'_n(\theta)$ ,  $= 1$  if  $x \in S'_n(\theta')^c \cap S'_n(\theta)^c$ . For  $c > 0$  and  $\delta > 0$  let  $E_n(c, \theta, \delta) = \{x; g_n(x; \theta, \theta_n^*) < c < c + \delta \leq h_n(x; \theta, \theta_n^*)\}$  and  $E_n(c, \theta, \delta) = \{x; g_n(x; \theta, \theta_n^*) > c > c - \delta > h_n(x; \theta, \theta_n^*)\}$ .

**Proposition.** Suppose that for some positive number  $\alpha$  and a sequence  $\{\theta_n^*\}_{n \in N}$  ( $\theta_n^* \in U_n$ )

$$(2.1) \quad \sup_{c>0} \sup_{\theta \in U_n} \{r_n^{\mathcal{B}_n}(c; \theta, \theta_n^*) - r_n^{\mathcal{A}_n}(c; \theta, \theta_n^*)\} = o(n^{-\alpha}).$$

Then we have

$$(2.2) \quad \begin{aligned} \sup_{c>0, \delta>0} \delta(1+c)^{-1} \lambda_n(c, \delta) &= o(n^{-\alpha}), \quad \text{and} \\ \sup_{c>0, \delta>0} \delta(1+c)^{-1} \lambda'_n(c, \delta) &= o(n^{-\alpha}) \end{aligned}$$

where  $\lambda_n(c, \delta) = \sup_{\theta \in U_n} P_{\theta_n^*, n}(E_n(c, \theta, \delta))$  and  $\lambda'_n(c, \delta) = \sup_{\theta \in U_n} P_{\theta_n^*, n}(E'_n(c, \theta, \delta))$ .

This proposition can be proved in the same way as the proof of the first and second steps of Theorem 1 in Suzuki [3]. So we shall omit the proof of the proposition.

**Theorem 1.** Suppose that Assumption 1 is satisfied with a sequence  $\{\theta_n^*\}_{n \in N}$  and  $\gamma > 0$ , and that  $\{\mathcal{B}_n\}$  has the property (2.1) with  $\beta > 0$ . Then for every  $\beta$  satisfying  $0 < \beta < 3^{-1}\alpha$  and  $\beta \leq \gamma$ ,  $\{\mathcal{B}_n\}$  is asymptotically sufficient for  $\{\mathcal{P}_n^{U_n}\}$  with order  $o(n^{-\beta})$  in the following sense: For each  $n \in N$  there exists a family  $\{q_n(x; \theta, \theta_n^*)\}$  of probability density functions on  $(\mathcal{X}, \mathcal{A}_n)$  relative to  $\mu_n$  such that

(i) each  $q_n$  can be factorized as follows:

$$q_n(x; \theta, \theta_n^*) = r_n(x; \theta, \theta_n^*) p_n(x, \theta_n^*)$$

where  $r_n$  is a  $\mathcal{B}_n$ -measurable function, and

$$(ii) \quad \sup_{\theta \in U_n} \int_{\mathcal{X}} |p_n(x, \theta) - q_n(x; \theta, \theta_n^*)| d\mu_n = o(n^{-\beta}).$$

**Proof.** We shall divide the proof into several steps.

The first step. Suppose that Assumption 1 is satisfied with a sequence  $\{\theta_n^*\}_{n \in N}$  and  $\gamma > 0$ , and that  $\{\mathcal{B}_n\}_{n \in N}$  has the property (2.1) with  $\alpha > 0$ . Let  $\beta$  be any number satisfying  $0 < \beta < 3^{-1}\alpha$  and  $\beta \leq \gamma$ . Take  $\varepsilon_1$  be any number satisfying  $0 < \varepsilon_1 < 3^{-1} \cdot (\alpha - 3\beta)$ . Let  $\alpha_n = n^{-\beta}(\log n)^{-1}$ ,  $m_n = n^{\varepsilon_1}$  and  $i_n = [m_n \alpha_n^{-1}] + 1$  where  $[a]$  means the maximum integer not exceeding  $a$ . Put  $(\gamma_n =) \gamma_n(x; \theta, \theta_n^*) = |h_n(x; \theta, \theta_n^*) - g_n(x; \theta, \theta_n^*)|$ ,  $(\gamma'_n =) \gamma'_n(x; \theta, \theta_n^*) = |h_n(x; \theta, \theta_n^*) - I_{W_n}(x; \theta, \theta_n^*)|$  and

$$\rho_n(\theta, \theta_n^*) = \int_{\mathcal{X}} \gamma'_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n}$$

where  $W_n = W_n(\theta, \theta_n^*) = \{x; g_n(x; \theta, \theta_n^*) \leq m_n\}$  and  $I_{W_n}$  means the indicator function of  $W_n$ .

We have

$$\begin{aligned}
\sup_{\theta \in U_n} \rho_n(\theta, \theta_n^*) &\leq \sup_{\theta \in U_n} \int_{W_n} \gamma_n dP_{\theta_n^*, n}^* + \sup_{\theta \in U_n} \int_{W_n^c} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n}^* = J_n^* + J_n^{**}, \\
(2.3) \quad \text{and} \quad J_n^* &= \sup_{\theta \in U_n} \int_{W_n} \gamma_n dP_{\theta_n^*, n}^* \leq \alpha_n + \sup_{\theta \in U_n} \int_{D_n \cap W_n} \gamma_n dP_{\theta_n^*, n}^* = \alpha_n + I_n \\
(D_n &= \{x; \gamma_n \geq \alpha_n\}).
\end{aligned}$$

Furthermore we have

$$I_n = \sup_{\theta \in U_n} \int_{D_n \cap W_n} \gamma_n dP_{\theta_n^*, n}^* \leq \sup_{\theta \in U_n} \int_{D_n \cap \tilde{W}_n} \gamma_n dP_{\theta_n^*, n}^* + \sup_{\theta \in U_n} \int_{W_n^*} \gamma_n dP_{\theta_n^*, n}^*$$

where  $W_n' = \{x; h_n(x; \theta, \theta_n^*) \leq m_n\}$ ,  $\tilde{W}_n = W_n \cap W_n'$  and  $W_n^* = W_n \cap (W_n')^c$ .

The second step. It holds that

$$\begin{aligned}
I_n' &= \sup_{\theta \in U_n} \int_{D_n \cap \tilde{W}_n} \gamma_n dP_{\theta_n^*, n}^* \leq \sum_{i=1}^{2i_n-2} \sup_{\theta \in U_n} \int_{B_i} \gamma_n dP_{\theta_n^*, n}^* \\
(2.4) \quad &+ \sum_{i=0}^{2i_n-3} \sup_{\theta \in U_n} \int_{C_i} \gamma_n dP_{\theta_n^*, n}^* \\
&= I_{n,1}' + I_{n,2}'
\end{aligned}$$

where  $B_i = \tilde{W}_n \cap \{x; h_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+1)\alpha_n, g_n(x; \theta, \theta_n^*) < 2^{-1}i\alpha_n\}$  and  $C_i = \tilde{W}_n \cap \{x; h_n(x; \theta, \theta_n^*) < 2^{-1}(i+1)\alpha_n, g_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+2)\alpha_n\}$ . Using the property (2.2) in Proposition we can evaluate  $I_{n,i}' (i=1, 2)$  as follows. Taking account of  $3\epsilon_1 < \alpha - 3\beta$  we have

$$\begin{aligned}
I_{n,1}' &= \sum_{i=1}^{2i_n-2} \sup_{\theta \in U_n} \int_{B_i} \gamma_n dP_{\theta_n^*, n}^* \\
&\leq 2i_n m_n \left[ \sup_{1 \leq i \leq 2i_n-2} \sup_{\theta \in U_n} P_{\theta_n^*, n}^* \{x; h_n(x; \theta, \theta_n^*) \geq 2^{-1}(i+1)\alpha_n, \right. \\
&\quad \left. g_n(x; \theta, \theta_n^*) < 2^{-1}i\alpha_n\} \right] \\
(2.5) \quad &\leq 2i_n m_n \left[ \sup_{1 \leq i \leq 2i_n-2} \lambda_n(2^{-1}i\alpha_n, 2^{-1}\alpha_n) \right] \\
&\leq 4i_n m_n \left[ \sup_{1 \leq i \leq 2i_n-2} \alpha_n^{-1} (1 + 2^{-1}i\alpha_n) n^{-\alpha} \eta_n' \right] (\eta_n' = o(1)) \\
&\leq 4i_n^2 m_n n^{-\alpha} \eta_n' \\
&\leq A_1 \cdot n^{-(\alpha-2\beta-3\epsilon_1)} (\log n)^2 \eta_n' \quad (A_1 \text{ is a constant}) \\
&= o(n^{-\beta}).
\end{aligned}$$

Similarly we have

$$(2.6) \quad I_{n,2}' = o(n^{-\beta}).$$

Thus from (2.4) and (2.5) we have

$$(2.7) \quad I_n' = o(n^{-\beta}).$$

The third step. Next we evaluate  $I'_n$  as follows. For every  $s > 1$  we have

$$\begin{aligned} I''_n &= \sup_{\theta \in U_n} \int_{W_n^*} \gamma_n dP_{\theta_n^*, n} \leq \sup_{\theta \in U_n} \int_{\{h_n > m_n\}} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ &\leq (m_n)^{1-s} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*). \end{aligned}$$

Hence we have

$$I''_n \leq A_2(s) (m_n)^{1-s} = A_2(s) n^{(1-s)\epsilon_1}$$

where  $A_2(s)$  is some constant depending only on  $s$ . We can choose  $s > 1$  large enough so that

$$(2.8) \quad I''_n = o(n^{-\beta}).$$

From (2.7) and (2.8) we have

$$I_n = o(n^{-\beta}).$$

Hence from (2.3) we have

$$J_n^* = o(n^{-\beta}).$$

Put  $W''_n = \{x; h_n(x; \theta, \theta_n^*) < 2^{-1} m_n\}$ . Then we have

$$\begin{aligned} J_n^{**} &= \sup_{\theta \in U_n} \int_{W_n^c} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ (2.9) \quad &\leq \sup_{\theta \in U_n} \int_{W_n^c \cap W''_n} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} + \sup_{\theta \in U_n} \int_{W_n^c \cap (W''_n)^c} h_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n} \\ &\leq 2^{-1} m_n \lambda'_n(m_n, m_n/2) + (m_n/2)^{1-s} \sup_{\theta \in U_n} J_n(s; \theta, \theta_n^*). \end{aligned}$$

The first term on the right hand side is of order  $o(n^{-\beta})$  by Proposition. The similar consideration as the evaluation of  $I''_n$  implies that the second term of (2.9) is also of order  $o(n^{-\beta})$  for sufficiently large number  $s$ . Thus we have

$$J_n^{**} = o(n^{-\beta}).$$

Hence it follows from (2.3) that

$$(2.10) \quad \sup_{\theta \in U_n} \rho_n(\theta, \theta_n^*) = o(n^{-\beta}).$$

The fourth step. Let  $a_n(\theta, \theta_n^*) = [\int_{\mathcal{X}} I_{W_n}(x) g_n(x; \theta, \theta_n^*) dP_{\theta_n^*, n}]^{-1} (\leq \infty)$  and let  $r_n(x; \theta, \theta_n^*) = a_n(\theta, \theta_n^*) I_{W_n}(x) g_n(x; \theta, \theta_n^*)$  if  $a_n(\theta, \theta_n^*) < \infty$ ,  $= 1$  otherwise. Define  $q_n(x; \theta, \theta_n^*) = r_n(x; \theta, \theta_n^*) p_n(x, \theta_n^*)$  and let  $Q_{\theta_n^*, n}^*$  be the probability distribution on  $(\mathcal{X}, \mathcal{A}_n)$  with density  $q_n(x; \theta, \theta_n^*)$  relative to  $\mu_n$ . We note that  $\mathcal{B}_n$  is suffi-

cient  $\sigma$ -field for the family  $\{Q_{\theta,n}^{\theta^*}; \theta \in \Theta\}$  by the factorization theorem. It follows from (2.10) that there exists  $n_0$  such that  $a_n(\theta, \theta_n^*) < \infty$  for every  $n \geq n_0$  and every  $\theta \in U_n$ . Therefore we can assume without loss of generality that  $a_n(\theta, \theta_n^*) < \infty$  for every  $\theta \in U_n$  and every  $n \geq 1$ .

Under this circumstances we have

$$\begin{aligned} \|P_{\theta,n} - Q_{\theta,n}^{\theta^*}\|_{\mathcal{A}_n} &= \int_{\mathcal{X}} |p_n(x, \theta) - q_n(x; \theta, \theta_n^*)| d\mu_n \\ &\leq \int_{S_n(\theta_n^*)} |h_n(x; \theta, \theta_n^*) - a_n(\theta, \theta_n^*) I_{W_n}(x) g_n(x; \theta, \theta_n^*)| \\ &\quad \cdot p_n(x, \theta_n^*) d\mu_n + \beta_n(\theta, \theta_n^*) \\ &= \rho_n(\theta, \theta_n^*) + |1 - a_n(\theta, \theta_n^*)^{-1}| + \beta_n(\theta, \theta_n^*) \\ &\leq 2 \rho_n(\theta, \theta_n^*) + 2 \beta_n(\theta, \theta_n^*). \end{aligned}$$

Here  $\|\nu\|_{\mathcal{A}_n}$  means the total variation norm of a signed measure  $\nu$  on  $(\mathcal{X}, \mathcal{A}_n)$ . From Assumption 1, (b) and (2.10) we have

$$\sup_{\theta \in U_n} \|P_{\theta,n} - Q_{\theta,n}^{\theta^*}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

This completes the proof of the theorem.

**3. The order of local asymptotic sufficiency.** In this section the index set  $\Theta$  is assumed to be a subset of  $p$ -dimensional Euclidean space  $R^p$ . We denote by  $|\cdot|$  the usual Euclidean norm in  $R^p$ . For  $\theta \in \Theta$  and  $b > 0$  let  $U_n(\theta, b) = \{\theta' \in \Theta; n^{1/2}|\theta' - \theta| \leq b\}$ .

Let  $\{\mathcal{B}_n\}_{n \in N}$  be the sequence of sub  $\sigma$ -fields  $\mathcal{B}_n \subset \mathcal{A}_n$  as in the previous section. We consider the following assumption which will be used to prove our main theorem, Theorem 2.

**ASSUMPTION 2.** For every compact subset  $K$  of  $\Theta$  and  $b > 0$

- (a)  $\limsup_{n \rightarrow \infty} \sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} J_n(s; \theta, \theta^*) < \infty \quad (\forall s > 1), \text{ and}$
- (b)  $\sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} \beta_n(\theta, \theta^*) = o(n^{-\gamma}).$

Let  $\alpha$  be a given positive number. We state a result about higher order locally asymptotic sufficiency of  $\{\mathcal{B}_n\}$  for  $\{\mathcal{P}_n\}$ .

**Theorem 2.** Suppose that Assumption 2 is satisfied with  $\gamma > 0$ , and that for every compact subset  $K$  of  $\Theta$  and every  $b > 0$

$$(3.1) \quad \sup_{c > 0} \sup_{\theta^* \in K} \sup_{\theta \in U_n(\theta^*, b)} \{r_n^{\mathcal{B}_n}(c; \theta, \theta^*) - r_n^{\mathcal{A}_n}(c; \theta, \theta^*)\} = o(n^{-\alpha}).$$

Then for every positive number  $\beta$  satisfying  $\beta < 3^{-1}\alpha$  and  $\beta \leq \gamma$   $\{\mathcal{B}_n\}_{n \in N}$  is locally asymptotically sufficient for  $\{\mathcal{P}_n\}$  with order  $o(n^{-\beta})$  in the following sense: For each

$n \in N$  and each  $\theta_0 \in \Theta$  there exists a family  $\mathcal{Q}_n^{\theta_0} = \{Q_{\theta,n}^{\theta_0}; \theta \in \Theta\}$  of probability distributions on  $(\mathcal{X}, \mathcal{A}_n)$  such that

- (i)  $\mathcal{B}_n$  is sufficient for  $\mathcal{Q}_n^{\theta_0}$ , and
- (ii) for every compact subset  $K$  of  $\Theta$  and every  $b > 0$

$$\sup_{\theta \in K} \sup_{\theta \in U_n(\theta, b)} \|P_{\theta,n} - Q_{\theta,n}^{\theta_0}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

Since the above result follows directly from Theorem 1 we shall omit the proof.

It is open problem whether non-local version of Theorem 2 still holds or not, i.e., whether any conditions such as in Theorem 2 imply the followings or not: There exists a sequence  $\mathcal{Q}_n = \{Q_{\theta,n}; \theta \in \Theta\}$  of probability distributions on  $(\mathcal{X}, \mathcal{A}_n)$  such that  $\mathcal{B}_n$  is sufficient for  $\mathcal{Q}_n$ , and that for every compact subset  $K$  of  $\Theta$

$$(3.2) \quad \sup_{\theta \in K} \|P_{\theta,n} - Q_{\theta,n}\|_{\mathcal{A}_n} = o(n^{-\beta}).$$

The case of  $\alpha = \beta = 0$  has been discussed in Suzuki [3] in such a non-local situation.

It is well known that under some regularity conditions there exist a sequence  $\{\hat{\theta}_n\}_{n \in N}$  of estimators of  $\theta$ , a positive number  $\gamma$  and a number  $v \geq 1$  having the following property: For every compact subset  $K$  of  $\Theta$  there corresponds  $a(K)$  such that

$$\sup_{\theta \in K} P_{\theta,n} \{n^{1/2} |\hat{\theta}_n(x) - \theta| \geq a(K) (\log n)^{v/2}\} = o(n^{-\gamma})$$

(c.f. Matsuda [2], Chap. 3).

Using such an estimator  $\{\hat{\theta}_n\}$  we may be able to construct  $\{Q_{\theta,n}; \theta \in \Theta\}$  satisfying the property (3.2), and for which  $\mathcal{B}_n$  is sufficient.

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