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# ASYMPTOTIC SUFFICIENCY I: REGULAR CASES

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1. Introduction. The concept of asymptotic sufficiency of maximum likelihood (m.l.) estimator is due to Wald [16] and this work was succeeded by LeCam [4] and Pfanzagl [10]. Higher order asymptotic sufficiency has been subsequently studied by Ghosh and Subramanyam [3], Michel [7] and Suzuki [14], [15].

Let  $\Theta$  be an open subset of the *s*-dimensional Euclidean space. Suppose that  $x_1, \dots, x_n$  are independent and identically distributed random variables with joint distribution  $P_{n,\theta}$ ,  $\theta \in \Theta$ , which has a constant support and satisfies certain regularity conditions. For  $\theta \in \Theta$  and  $z_n = (x_1, \dots, x_n)$  let  $G_n^{(m)}(z_n, \theta)$ denote the *m*-th derivative relative to  $\theta$  of the log-likelihood function. In Michel [7], it was shown that for  $k \ge 3$  a statistic  $T_{n,k} = (T_n, G_n^{(2)}(z_n, T_n), \dots,$  $G_n^{(k)}(z_n, T_n))$ , where  $\{T_n\}$  is a sequence of asymptotic m.l. estimators of order  $o(n^{-(k-2)/2})$  (see Definition in Section 3), is asymptotically sufficient up to order  $o(n^{-(k-2)/2})$  in the following sense: For each  $n \in N$ ,  $T_{n,k}$  is sufficient for a family  $\{Q_{n,\theta}; \theta \in \Theta\}$  of probability distributions and for every compact subset K of  $\Theta$ 

$$\sup_{\theta \in K} ||P_{n,\theta} - Q_{n,\theta}|| = o(n^{-(k-2)/2}),$$

where  $||\cdot||$  means the total variation of a measure. Suzuki [14], [15] also showed that for  $k \in N$  a statistic  $(\hat{\theta}_n, G_n^{(1)}(z_n, \hat{\theta}_n), \dots, G_n^{(k)}(z_n, \hat{\theta}_n))$ , where  $\hat{\theta}_n$  is a reasonable estimator including m.l. estimator, is asymptotically sufficient up to order  $o(n^{-(k-1)/2})$  under a stronger moment condition than in Michel [7].

In this paper we give a refinement of their results on higher order asymptotic sufficiency. Our result includes that (1)  $T_{n,k} = (T_n, G_n^{(2)}(z_n, T_n), \cdots, G_n^{(k)}(z_n, T_n))$  is asymptotically sufficient up to order  $O(n^{-k/2})$  for any sequence  $\{T_n\}$  of asymptotic m.l. estimators of order  $O(n^{-k/2})$  and (2) a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$  with some  $r \in (0, 1)$  is asymptotically sufficient up to order  $O(n^{-r/2})$  and the second derivatives of the log-likelihood function.

In the case k=1, Pfanzagl ([10], Theorem 1) proved that a sequence of estimators with properties analogous to those of asymptotic m.l. estimators of order  $O(n^{-1/2})$  is asymptotically sufficient up to order  $O(n^{-1/2})$ , and showed in

[11] that this order of convergence cannot be improved in general. Thus our result is an extension of his and it seems to be impossible to improve the convergence order  $O(n^{-k/2})$ .

In Section 2 we present a result concerning probabilities of deviations for sums of independent and identically distributed random variables with a restricted moment. In Section 3 we investigate asymptotic sufficiency of  $T_{n,k}$  constructed by asymptotic m.l. estimators  $T_n$ . In the final Section 4 we give conditions under which a sequence of m.l. estimators becomes the one of asymptotic m.l. estimators of order  $O(n^{-r/2})$  with some r > 0.

2. Probabilities of deviations. Let  $Y_1, \dots, Y_n$  be a sequence of random variables (r.v.'s) and put  $S_m = \sum_{i=1}^m Y_i$ ,  $1 \le m \le n$ . Using the elementary inequality

$$E|S_n|^r \leq \sum_{i=1}^n E|Y_i|^r, \quad r \leq 1,$$

it follows from Markov's inequality that for x > 0

(2.1) 
$$P\{|S_n| \ge x\} \le x^{-r} \sum_{i=1}^n E|Y_i|^r.$$

If the r.v.'s satisfy the relations

(2.2) 
$$E(Y_{m+1}|S_m) = 0 \text{ a.s. } 1 \leq m \leq n-1,$$

then von Bahr and Esseen [1] showed that

(2.3) 
$$E |S_n|^r \leq 2 \sum_{i=1}^n E |Y_i|^r, \quad 1 \leq r \leq 2.$$

The condition (2.2) is fulfilled if the r.v.'s are independent and have zero means. In this case, (2.3) together with Markov's inequality implies the following inequality

(2.4) 
$$P\{|S_n| \ge x\} \le 2x^{-r} \sum_{i=1}^n E|Y_i|^r, \quad 1 \le r \le 2,$$

for x > 0.

The following theorem includes a uniform version of Corollary 2 in Nagaev [9].

**Theorem 1.** Let  $Y_1, \dots, Y_n$  be a sequence of independent and identically distributed random variables with a common distribution  $P_{\theta}$ ,  $\theta \in K$ , where K is any set. Let  $h(y, \theta)$  be a measurable function of y for any fixed  $\theta \in K$  and put  $S_{n,\theta} = \sum_{i=1}^{n} h(Y_i, \theta)$ . If  $E_{\theta}(h(Y_1, \theta)) = 0$  for all  $\theta \in K$  and  $\xi_r = \sup_{\theta \in K} E_{\theta} |h(Y_1, \theta)|^r < \infty$  for some r > 0, then

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(2.5) 
$$\sup_{\theta \in K_{r}} P_{\theta}\{|S_{n,\theta}| \geq x\} = O(nx^{-r}), \quad 0 < r \leq 2,$$

for x > 0, and

(2.6) 
$$\sup_{\theta \in \mathcal{K}} P_{\theta}\{|S_{n,\theta}| > \xi_r^{1/r}x\} = O(nx^{-r}), \quad r > 2,$$

for  $x \ge \sqrt{8(r-2)n \log n}$ .

Proof. (2.5) is an immediate consequence of (2.1) and (2.4).

For the proof of (2.6) we use the following inequality which is a slight modification of Theorem 1 in Nagaev [9]: For x>0 and y>0,

(2.7) 
$$P_{\theta}\{|S_{n,\theta}| > x\} < 2nP_{\theta}\{|h(Y_{1}, \theta)| > y\} + 2\left[\frac{n\xi_{r,\theta}d_{r}}{y^{r}}\right]^{x/y} \\ \times \exp\left\{2n\left[\frac{r\log y - \log(n\xi_{r,\theta}d_{r})}{y}\right]^{2}\xi_{r,\theta}^{2/r} + 1\right\},$$

where  $d_r = 1 + (r+1)^{r+2} \exp(-r)$  and  $\xi_{r,\theta} = E_{\theta} |h(Y_1, \theta)|^r$ . In order to show (2.7) it is enough to note that the relation (2.3) in [9] becomes

$$\left|\int_{-\infty}^{1/h} \exp\left\{h[h(Y_1,\theta)]\right\} dP_{\theta} - 1\right| < 2h^2 \xi_{r,\theta}^{2/r}.$$

Setting  $x = \xi_{r,\theta} l^{r} n^{1/2} t$  and y = x/2 for  $t \ge \sqrt{8(r-2)\log n}$  in (2.7), then we obtain

(2.8) 
$$nP_{\theta}\{|h(Y_1, \theta)| > y\} \leq n\xi_{r,\theta}y^{-r}$$
$$= 2^r n^{(2-r)/2} t^{-r}$$

and

(2.9) 
$$\left[\frac{n\xi_{r,\theta}d_r}{y^r}\right]^{x/y} = 2^{2r}d_r^2n^{2-r}t^{-2r}.$$

Let us assume that  $n \ge \exp\left\{\frac{1}{8(r-2)}\right\}$ . (For  $n < \exp\left\{\frac{1}{8(r-2)}\right\}$ , (2.6) is trivially true.) Since  $0 \le t^{-1} \log t < 1/2$  for  $t \ge 1$ , we have

$$2n \left[ \frac{r \log y - \log (n\xi_{r,\theta}d_r)}{y} \right]^2 \xi_{r,\theta}^{2/r}$$
  
=  $8t^{-2} [r \log t + \frac{r-2}{2} \log n - r \log 2 - \log d_r]^2$   
 $\leq 8t^{-2} [r^2 (\log t)^2 + \frac{(r-2)^2}{4} (\log n)^2 + r(r-2) \log n \log t + c_1]$   
 $\leq 2r^2 + \frac{r-2}{4} \log n + r \log t + c_2,$ 

where  $c_1$  and  $c_2$  denote positive constants depending only on r. From this fact and (2.9) it follows that the second term on the right side of (2.7) has an upper bound of the type  $c_3 n^{3(2-r)/4} t^{-r}$ . This, together with (2.8), implies (2.6).

REMARK. (1) In the case  $r \ge 3$ , Michel [7] showed a result analogous to Theorem 1 (cf. also Lemma 1 in Pfanzagl [12]).

(2) Let  $Y_1, \dots, Y_n$  be a sequence of independent r.v.'s with zero means. It follows from an inequality due to Marcinkiewicz and Zygmund [5] that

(2.10) 
$$E |\sum_{i=1}^{n} Y_{i}|^{r} \leq c n^{(r-2)/2} \sum_{i=1}^{n} E |Y_{i}|^{r}, \quad r \geq 2,$$

where c is a positive constant depending only on r (see Chung [2], page 348). This leads to Lemma 2 in Pfanzagl [12] which requires a stronger moment condition than in Theorem 1 to evaluate probability of moderate deviations or large deviations.

3. Asymptotic sufficiency. Let  $\Theta$  be an open subset of the *s*-dimensional Euclidean space  $\mathbb{R}^s$  and for each  $\theta \in \Theta$ , let  $P_{\theta}$  be a probability measure on a measurable space  $(X, \mathcal{A})$ . It is assumed that  $P_{\theta}, \theta \in \Theta$ , is dominated by a  $\sigma$ -finite measure  $\mu$  on  $(X, \mathcal{A})$  and has a positive density  $p(x, \theta)$ . For each  $n \in N = \{1, 2, \dots\}$ , let  $(X^n, \mathcal{A}^n)$  be the Cartesian product of *n* copies of  $(X, \mathcal{A})$  and  $P_{n,\theta}$  be the product measure of *n* copies of  $P_{\theta}$ . Furthermore, let  $\mu_n$  denote the product measure of *n* copies of  $\mu$  and write  $p_n(z_n, \theta) = dP_{n,\theta}/d\mu_n$  for  $\theta \in \Theta$  and  $z_n = (x_1, \dots, x_n) \in X^n$ .

For a function  $h(z, \cdot)$ :  $\mathbf{R}^s \to \mathbf{R}$  denote the *m*-th derivative relative to  $\theta$  of  $h(z, \theta)$  by

$$h^{(m)}(z, \theta) = \left(\frac{\partial^m}{\partial \theta_{i_1} \cdots \partial \theta_{i_m}} h(z, \theta); i_1, \cdots, i_m \in \{1, \cdots, s\}\right).$$

In particular, we write

$$egin{aligned} h^{(1)}(z,\, heta) &= \left(rac{\partial}{\partial heta_1}h(z,\, heta),\,\cdots,\,rac{\partial}{\partial heta_s}h(z,\, heta)
ight),\ h^{(2)}(z,\, heta) &= \left(rac{\partial^2}{\partial heta_i\partial heta_j}h(z,\, heta)
ight), \end{aligned}$$

that is,  $h^{(1)}$  means a row vector and  $h^{(2)}$  a matrix. The Euclidean norm  $||\cdot||$  of  $h^{(m)}$  is defined by

$$||h^{(m)}(z, \theta)||^2 = \sum_{i_1, \cdots, i_m=1}^s \left(\frac{\partial^m}{\partial \theta_{i_1} \cdots \partial \theta_{i_m}} h(z, \theta)\right)^2.$$

For any  $\sigma = (\sigma_1, \dots, \sigma_s) \in \mathbf{R}^s$  define

$$h^{(m)}(z, \theta)\sigma^{m} = \sum_{i_{1}, \cdots, i_{m}=1}^{s} \frac{\partial^{m}}{\partial \theta_{i_{1}} \cdots \partial \theta_{i_{m}}} h(z, \theta) \prod_{p=1}^{m} \sigma_{i_{p}}.$$

Then, it is easy to see that

$$|h^{(m)}(z, \theta)\sigma^{m}| \leq ||h^{(m)}(z, \theta)|| ||\sigma||^{m}.$$

Let  $k \in N$  and r > 0 be fixed. We shall impose the following Conditions  $A, B_r$  and  $C_{k,r}$  on  $p(x, \theta)$ .

Condition A

(i) For each  $x \in X$ ,  $\theta \to p(x, \theta)$  admits continuous partial derivatives up to the order 2 on  $\Theta$ .

Let  $g(x, \theta) = \log p(x, \theta)$  and  $g^{(m)}$  be the *m*-th derivative of g defined above. Moreover, for  $\theta \in \Theta$  let  $J(\theta) = E_{\theta}(-g^{(2)}(\cdot, \theta))$ .

- (ii) For every  $\theta \in \Theta$
- (a)  $E_{\theta}(g^{(1)}(\cdot, \theta))=0$
- (b)  $J(\theta)$  is positive definite.

Condition  $B_r$ 

For every compact  $K \subset \Theta$ 

$$\sup_{\theta\in K} E_{\theta}(||g^{(1)}(\cdot, \theta)||^{r+2}) < \infty$$

Condition  $C_{k,r}$ 

(i) For each  $x \in X$ ,  $\theta \rightarrow p(x, \theta)$  admits continuous partial derivatives up to the order k+1 on  $\Theta$ .

(ii) For every  $\theta \in \Theta$  there exist a neighborhood  $U_{\theta}$  of  $\theta$  and a measurable function  $\lambda(x, \theta)$  such that

- (a) for all  $x \in X$ ,  $\tau$ ,  $\sigma \in U_{\theta}$ ,  $||g^{(k+1)}(x, \tau) g^{(k+1)}(x, \sigma)|| \leq ||\tau \sigma||\lambda(x, \theta)|$
- (b) for every compact  $K \subset \Theta$ ,  $\sup_{\tau \in K} E_{\tau}(\lambda(\cdot, \theta)^{(r+2)/2}) < \infty$

(c) 
$$\sup_{\tau\in\overline{U}_{\mathfrak{g}}} E_{\tau}(||g^{(k+1)}(\cdot, \tau)||^{\nu(\tau)}) < \infty,$$

where

$$\nu(r) = \frac{2+r}{2-r}, \quad \text{if } 0 < r < 1, \\
= r+2, \quad \text{if } r \ge 1.$$

(iii) For every compact  $K \subset \Theta$  there exist  $\delta_K > 0$  and  $\eta_K > 0$  such that  $\theta \in K$  and  $\tau \in \Theta$  with  $||\theta - \tau|| < \delta_K$  imply

$$||E_{\theta}(g^{(k+1)}(\cdot, \theta)) - E_{\tau}(g^{(k+1)}(\cdot, \tau))|| \leq \eta_{K} ||\theta - \tau||.$$

REMARK. (1) Condition (iii) in  $C_{k,r}$  follows from conditions (3)(a) and (3) (b) in Suzuki [15] (see also (3.4) in [14]).

(2) It is easily seen that condition (ii) in  $C_{k,r}$  and the following condition (iii)' imply condition (iii) in  $C_{k,r}$ .

(iii)' For every  $\theta \in \Theta$  there exist a neighborhood  $U_{\theta}$  of  $\theta$  and a measurable function  $\lambda^*(x, \theta)$  such that for all  $x \in X$ ,  $\tau \in U_{\theta}$ 

 $|p(x, \tau)/p(x, \theta)-1| \leq ||\tau-\theta||\lambda^*(x, \theta)$ 

and for every compact  $K \subset \Theta$ 

$$\sup_{\tau\in\kappa}E_{\tau}(\lambda^*(\cdot,\theta)^{\nu(r)/(\nu(r)-1)})<\infty.$$

The following definition is due to Michel [7].

DEFINITION.  $T_n$ ,  $n \in N$ , is a sequence of asymptotic maximum likelihood (m.l.) estimators of order  $O(n^{-r/2})$ , r>0, if there exist positive constants  $\pi_1$  and  $\pi_2$  (depending on r) such that for every compact  $K \subset \Theta$ 

$$\begin{aligned} &(\alpha_r) \quad \sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; \, n^{1/2} || T_n(z_n) - \theta || \ge (\log n)^{\pi_1} \} = O(n^{-r/2}) \\ &(\beta_r) \quad \sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; \, n^{r/2} || \sum_{i=1}^n g^{(1)}(x_i, \, T_n(z_n)) || \ge (\log n)^{\pi_2} \} = O(n^{-r/2}) \,. \end{aligned}$$

Asymptotic m.l. estimators can be obtained from suitable initial estimators by applying a Newton-Raphson method (see Michel [6] and Pfanzagl [13]).

To simplify our notations we shall use  $n_K$  (depending on compact K) as a generic constant instead of the phrase "for all sufficiently large n". In the same manner we shall use  $c_K$  as a generic constant to denote factors occurring in the bounds which depend on compact K but not on  $\theta \in K$  and  $n \in N$ .

**Lemma 1.** Assume that Condition  $C_{k,r}$  is fulfilled for some  $k \in N$  and r>0. Let  $T_n$ ,  $n \in N$ , be a sequence of estimators with the property  $(\alpha_r)$ . Then for every compact  $K \subset \Theta$ 

$$\sup_{\theta \in K} P_{n,\theta} \{ \sup_{n^{1/2} \| T_n - \tau \| \leq (\log n)^{\pi_1}} || \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_{\tau}(g^{(k+1)}(\cdot, \tau))] || \geq \psi(n, r) \} = O(n^{-r/2}),$$

where

$$\psi(n, r) = n^{(2-r)/2}, \quad \text{if } 0 < r < 1,$$
  
=  $n^{1/2} (\log n)^{\pi_1 + 1/2}, \quad \text{if } r \ge 1.$ 

Proof. Let 0 < r < 1 and K be a compact subset of  $\Theta$ . Condition (ii) implies that there exist  $d_K > 0$  and  $\lambda_K(x)$  such that  $\theta \in K$  and  $\tau \in \Theta$  with  $||\theta - \tau|| < d_K$  imply  $||g^{(k+1)}(x, \theta) - g^{(k+1)}(x, \tau)|| \le ||\theta - \tau||\lambda_K(x)$  for all  $x \in X$ , and such that  $\sup_{\theta \in K} E_{\theta}(\lambda_K(\cdot)^{(r+2)/2}) < \infty$ . Let

$$D_{n,\theta,K} = \{z_n \in X^n; |\sum_{i=1}^n [\lambda_K(x_i) - E_{\theta}(\lambda_K(\cdot))]| < n\}.$$

According to Theorem 1

(3.1) 
$$\sup_{\theta \in K} P_{n,\theta}\{(D_{n,\theta,K})^c\} = O(n^{-r/2}).$$

Furthermore, Theorem 1 together with condition (ii) (c) implies that

(3.2) 
$$\sup_{\theta \in K} P_{n,\theta}\{(F_{n,\theta})^c\} = O(n^{-r/2}),$$

where

$$F_{n,\theta} = \{z_n \in X^n; \|\sum_{i=1}^n [g^{(k+1)}(x_i, \theta) - E_{\theta}(g^{(k+1)}(\cdot, \theta))]\| < 1/2 n^{(2-r)/2} \}.$$

Let  $e_K > 0$  be such that  $\{\tau \in \mathbf{R}^s; \inf_{\theta \in K} ||\theta - \tau|| \leq e_K\} \subset \Theta$ . Choose  $n_K$  to satisfy  $2n^{-1/2}(\log n)^{\pi_1} < \min \{d_K, e_K, \delta_K\}$  for all  $n \geq n_K$ , where  $\delta_K$  appears in condition (iii). Then, by conditions (ii) and (iii), for  $n \geq n_K$ ,  $\theta \in K$ ,  $\tau \in \mathbf{R}^s$  with  $||\theta - \tau|| \leq 2n^{-1/2}(\log n)^{\pi_1}$  and  $z_n \in D_{n,\theta,K} \cap F_{n,\theta}$ 

$$\begin{split} &\|\sum_{i=1}^{n} \left[ g^{(k+1)}(x_{i}, \tau) - E_{\tau}(g^{(k+1)}(\cdot, \tau)) \right] \| \\ &\leq \|\sum_{i=1}^{n} \left[ g^{(k+1)}(x_{i}, \tau) - g^{(k+1)}(x_{i}, \theta) \right] \| + \|\sum_{i=1}^{n} \left[ g^{(k+1)}(x_{i}, \theta) - E_{\theta}(g^{(k+1)}(\cdot, \theta)) \right] \| \\ &+ n \| E_{\theta}(g^{(k+1)}(\cdot, \theta)) - E_{\tau}(g^{(k+1)}(\cdot, \tau)) \| \\ &\leq n [1 + \sup_{\theta \in \mathcal{K}} E_{\theta}(\lambda_{\mathcal{K}}(\cdot)) + \eta_{\mathcal{K}}] \| \theta - \tau \| + 1/2 n^{(2-r)/2} \\ &< n^{(2-r)/2} . \end{split}$$

Taking account of (3.1) and (3.2), for every compact  $K \subset \Theta$  we obtain

$$\sup_{\theta \in \mathcal{K}} P_{n,\theta} \{ \sup_{n^{1/2} \| \theta - \tau \| \le 2(\log n)^{\kappa_1}} \| \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_{\tau}(g^{(k+1)}(\cdot, \tau))] \| \ge n^{(2-r)/2} \}$$
  
=  $O(n^{-r/2}).$ 

This together with the property  $(\alpha_r)$  leads to the desired assertion.

In the case  $r \ge 1$ , it is enough to show that there exists  $c_K > 0$  such that

(3.3) 
$$\sup_{\boldsymbol{\theta}\in\boldsymbol{F}} P_{\boldsymbol{n},\boldsymbol{\theta}}\{(F_{\boldsymbol{n},\boldsymbol{\theta},\boldsymbol{K}})^c\} = o(\boldsymbol{n}^{-r/2}),$$

where

$$F_{n,\theta,K} = \{z_n \in X^n; || \sum_{i=1}^n [g^{(k+1)}(x_i, \theta) - E_{\theta}(g^{(k+1)}(\cdot, \theta))] || \leq c_K (n \log n)^{1/2} \}.$$

This follows from Theorem 1 and condition (ii) (c).

**Lemma 2.** Assume that Conditions A,  $B_r$  and  $C_{1,r}$  are fulfilled for some r>0. Let  $T_n$ ,  $n \in N$ , be a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$ . Then for every compact  $K \subset \Theta$ 

$$\sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; ||T_n(z_n) - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1} || \ge \omega(n, r) \} = O(n^{-r/2}),$$

where

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$$\begin{split} \omega(n, r) &= n^{-(r+1)/2} (\log n)^{2\pi_1}, \quad \text{if } 0 < r < 1, \\ &= n^{-1} (\log n)^{2\pi_1 + 1/2}, \quad \text{if } r \ge 1. \end{split}$$

Proof. Let 0 < r < 1 and K be a compact subset of  $\Theta$ . Condition (ii) in  $C_{1,r}$  implies that there exist  $d_K > 0$  and  $\lambda_K(x)$  such that  $x \in X$ ,  $\theta \in K$  and  $\tau \in \Theta$ with  $||\theta - \tau|| < d_K$  imply  $||g^{(2)}(x,\theta) - g^{(2)}(x,\tau)|| \le ||\theta - \tau||\lambda_K(x)$  and such that  $\sup_{\theta \in K} E_{\theta}(\lambda_K(\cdot)^{(r+2)/2}) < \infty$ . As in the proof of Lemma 1, we define

$$D_{n,\theta,K} = \{z_n \in X^n; |\sum_{i=1}^n [\lambda_K(x_i) - E_{\theta}(\lambda_K(\cdot))]| < n\},$$
  
$$F_{n,\theta} = \{z_n \in X^n; ||\sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)|| < 1/2 n^{(2-r)/2}\}.$$

It follows from Theorem 1, condition (ii) (a) in A and Condition  $B_r$ , that there exists  $c_R > 0$  such that

(3.4) 
$$\sup_{\theta \in K} P_{n,\theta} \{ (H_{n,\theta,K})^c \} = o(n^{-r/2}),$$

where

$$H_{n,\theta,K} = \{z_n \in X^n; ||\sum_{i=1}^n g^{(1)}(x_i, \theta)|| < c_K(n \log n)^{1/2}\}.$$

Let  $U_{n,\theta}$  and  $V_{n,r}$  be defined by

$$U_{n,\theta} = \{z_n \in X^n; n^{1/2} || T_n(z_n) - \theta || < (\log n)^{\pi_1} \},\$$
  
$$V_{n,r} = \{z_n \in X^n; n^{r/2} || \sum_{i=1}^n g^{(1)}(x_i, T_n(z_n)) || < (\log n)^{\pi_2} \}.$$

Choose  $e_K > 0$  such that  $K^* = \{\tau \in \mathbb{R}^s; \inf_{\theta \in K} ||\theta - \tau|| \leq e_K\} \subset \Theta$  and  $n_K$  such that  $n^{-1/2}(\log n)^{\pi_1} < \min\{d_K, e_K, \delta_K\}$  for all  $n \geq n_K$ , where  $\delta_K$  is determined by condition (iii) in  $C_{1,r}$ . It is obvious that  $n \geq n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta}$  imply  $T_n(z_n) \in K^*$ . Since  $K^*$  is a compact subset of  $\Theta$ , conditions (ii) (b) in A and (iii) in  $C_{1,r}$  imply that

$$ho_{K*} = \sup_{ au \in K*} ||J( au)^{-1}|| < \infty$$

Using the equality

$$\sum_{i=1}^{n} g^{(1)}(x_{i}, \theta) = \sum_{i=1}^{n} g^{(1)}(x_{i}, T_{n}) + (\theta - T_{n}) \sum_{i=1}^{n} \overline{g}^{(2)}(x_{i}, T_{n}, \theta)$$

with  $g^{(2)}(x, \theta, \sigma) = \int_0^1 g^{(2)}(x, (1-t)\theta + t\sigma)dt$ , we obtain

$$T_{n} - \theta - n^{-1} \sum_{i=1}^{n} g^{(1)}(x_{i}, \theta) J(\theta)^{-1} + n^{-1} \sum_{i=1}^{n} g^{(1)}(x_{i}, T_{n}) J(T_{n})^{-1}$$
  
=  $T_{n} - \theta - n^{-1} \sum_{i=1}^{n} g^{(1)}(x_{i}, \theta) [J(\theta)^{-1} - J(T_{n})^{-1}] + (T_{n} - \theta) n^{-1} \sum_{i=1}^{n} \bar{g}^{(2)}(x_{i}, T_{n}, \theta) J(T_{n})^{-1}$ 

$$= (T_n - \theta)[J(T_n) - J(\theta)]J(T_n)^{-1} - n^{-1}\sum_{i=1}^n g^{(1)}(x_i, \theta)J(\theta)^{-1}[J(T_n) - J(\theta)]J(T_n)^{-1} + (T_n - \theta)n^{-1}\sum_{i=1}^n [g^{(2)}(x_i, T_n, \theta) - g^{(2)}(x_i, \theta)]J(T_n)^{-1} + (T_n - \theta)n^{-1}\sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]J(T_n)^{-1}.$$

Hence we have for  $n \ge n_K$ ,  $\theta \in K$  and  $z_n \in D_{n,\theta,K} \cap F_{n,\theta} \cap H_{n,\theta,K} \cap U_{n,\theta} \cap V_{n,r}$ 

$$\begin{split} ||T_{n}-\theta-n^{-1}\sum_{i=1}^{n}g^{(1)}(x_{i},\theta)J(\theta)^{-1}|| \\ &\leq ||n^{-1}\sum_{i=1}^{n}g^{(1)}(x_{i},T_{n})J(T_{n})^{-1}||+||T_{n}-\theta-n^{-1}\sum_{i=1}^{n}g^{(1)}(x_{i},\theta)J(\theta)^{-1} \\ &+n^{-1}\sum_{i=1}^{n}g^{(1)}(x_{i},T_{n})J(T_{n})^{-1}|| \\ &\leq \rho_{K*}n^{-(r+2)/2}(\log n)^{\pi_{2}}+(c_{K}\eta_{K}\rho_{K*}^{2}n^{-1/2}(\log n)^{/12}+1/2|\rho_{K*}n^{-r/2})||T_{n}-\theta|| \\ &+\rho_{K*}(1+\eta_{K}+\sup_{\theta\in K}E_{\theta}(\lambda_{K}(\cdot)))||T_{n}-\theta||^{2} \\ &\leq c_{K}n^{-(r+1)/2}(\log n)^{\pi_{1}}. \end{split}$$

This implies the desired result because of (3.1), (3.2), (3.4) and the properties  $(\alpha_r), (\beta_r)$ .

For the case  $r \ge 1$ , the proof is also similar except that  $F_{n,\theta}$  is replaced by  $F_{n,\theta,K}$  in (3.3) with k=1.

For simplicity, we write

$$G_n^{(m)}(z_n, \theta) = \sum_{i=1}^n g^{(m)}(x_i, \theta), \quad z_n = (x_1, \cdots, x_n) \in X^n, \quad \theta \in \Theta.$$

Now we can present a result on asymptotic sufficiency of the statistic

$$\begin{split} T_{n,k} &= T_n, & k = 1, \\ &= (T_n, \, G_n^{(2)}(z_n, \, T_n), \, \cdots, \, G_n^{(k)}(z_n, \, T_n)), & k \ge 2, \end{split}$$

where  $T_n$ ,  $n \in N$ , is a sequence of asymptotic m.l. estimators.

**Theorem 2.** Assume that Conditions A,  $B_r$ ,  $C_{1,r}$  and  $C_{k,r}$  hold for some  $k \in N$  and r > 0. Let  $T_n$ ,  $n \in N$ , be a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$ . Then there exists a sequence of families of probability measures  $\{Q_{n,\theta}^k; \theta \in \Theta\}, n \in N$ , such that

- (a) for each  $n \in N$ ,  $T_{n,k}$  is sufficient for  $\{Q_{n,\theta}^k; \theta \in \Theta\}$
- (b) for every compact  $K \subset \Theta$

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$$\sup_{\theta \in \mathcal{K}} ||P_{n,\theta} - Q_{n,\theta}^k|| = O(n^{-r/2}), \quad \text{if } r < k,$$
$$= O(n^{-k/2}), \quad \text{if } r \ge k.$$

Proof. Let

$$\begin{split} U_{n,\theta} &= \{z_n \in X^n; n^{1/2} || T_n - \theta || < (\log n)^{\pi_1} \}, \\ V_{n,r} &= \{z_n \in X^n; n^{r/2} || \sum_{i=1}^n g^{(1)}(x_i, T_n) || < (\log n)^{\pi_2} \}, \\ W_{n,r} &= \{z_n \in X^n; \sup_{\substack{n^{1/2} || T_n - \tau || \le (\log n)^{\pi_1} i = 1}} || \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_\tau(g^{(k+1)}(\cdot, \tau))] || < \psi(n, r) \}, \end{split}$$

where  $\psi(n, r)$  is the same as in Lemma 1. We define

$$\overline{q}_{n,k}(z_n, \theta) = I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}}(z_n) \exp \{G_n(z_n, T_n) + \sum_{m=2}^k \frac{1}{m!} G_n^{(m)}(z_n, T_n)(\theta - T_n)^m + \frac{1}{(k+1)!} [E_{\theta}(G_n^{(k+1)}(\cdot, \theta))](\theta - T_n)^{k+1}\},$$

$$q_{n,k}(z_n,\theta) = v_n(\theta)\overline{q}_{n,k}(z_n,\theta),$$

where  $v_n(\theta) = [\int_{X^n} \bar{q}_{n,k}(z_n, \theta) d\mu_n]^{-1}$ . Here and hereafter  $I_U(\cdot)$  means the indicator function of a set U. For  $\theta \in \Theta$  and  $n \in N$ , we denote by  $\bar{Q}_{n,\theta}^k$  and  $Q_{n,\theta}^k$  the measures given by

$$\frac{d\overline{Q}_{n,\theta}^{\ k}}{d\mu_n} = \overline{q}_{n,k} \quad \text{and} \quad \frac{dQ_{n,\theta}^{\ k}}{d\mu_n} = q_{n,k} \,.$$

Then it follows from the factorization theorem that for each  $n \in N$ ,  $T_{n,k}$  is sufficient for  $\{Q_{n,\theta}^k; \theta \in \Theta\}$ .

In order to prove the second assertion (b) we fix a compact subset K of  $\Theta$ . Using the Taylor expansion

$$G_n(z_n, \theta) = G_n(z_n, T_n) + \sum_{m=1}^k \frac{1}{m!} G_n^{(m)}(z_n, T_n) (\theta - T_n)^m + \frac{1}{(k+1)!} G_n^{(k+1)}(z_n, T_n^*) (\theta - T_n)^{k+1}$$

where max  $\{||T_n^* - \theta||, ||T_n^* - T_n||\} \leq ||T_n - \theta||$ , we have for  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$ 

(3.5) 
$$\begin{vmatrix} \log \frac{\overline{q}_{n,k}(z_n, \theta)}{p_n(z_n, \theta)} \end{vmatrix} \\ \leq ||G_n^{(1)}(z_n, T_n)|| \, ||T_n - \theta|| + \frac{1}{(k+1)!} ||G_n^{(k+1)}(z_n, T_n^*) \\ - E_{\theta}(G_n^{(k+1)}(\cdot, \theta))|| \, ||T_n - \theta||^{k+1}. \end{cases}$$

Since

$$\begin{aligned} & \|G_{n}^{(k+1)}(z_{n}, T_{n}^{*}) - E_{\theta}(G_{n}^{(k+1)}(\cdot, \theta))\| \\ & \leq \|G_{n}^{(k+1)}(z_{n}, T_{n}^{*}) - [E_{\tau}(G_{n}^{(k+1)}(\cdot, \tau))]_{\tau=T_{n}^{*}}\| \\ & + \|[E_{\tau}(G_{n}^{(k+1)}(\cdot, \tau))]_{\tau=T_{n}^{*}} - E_{\theta}(G_{n}^{(k+1)}(\cdot, \theta))\| \end{aligned}$$

it follows from (3.5) and condition (iii) in  $C_{k,r}$  that for  $n \ge n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$ 

(3.6) 
$$\left|\log \frac{\overline{q}_{n,k}(z_n, \theta)}{p_n(z_n, \theta)}\right| \leq n^{-(r+1)/2} (\log n)^{\pi_1 + \pi_2} + n^{-(k+1)/2} (\log n)^{(k+1)\pi_1} \psi(n, r).$$

This implies that for  $n \ge n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$ 

(3.7) 
$$\left|\log \frac{\overline{q}_{n,k}(z_n,\theta)}{p_n(z_n,\theta)}\right| \leq \log 2$$

Using the inequality  $|1-\exp(x)| \leq 2|x|$  for  $|x| \leq \log 2$ , then from (3.7) we have for  $n \geq n_K$  and  $\theta \in K$ 

$$(3.8) \qquad ||P_{n,\theta} - \overline{Q}_{n,\theta}^{k}|| \\ \leq \int_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}} \left|1 - \frac{\overline{q}_{n,k}(z_{n},\theta)}{p_{n}(z_{n},\theta)}\right| dP_{n,\theta} + P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^{c}\} \\ \leq 2E_{\theta}\left[\left|\log \frac{\overline{q}_{n,k}(\cdot,\theta)}{p_{n}(\cdot,\theta)}\right| I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}}(\cdot)\right] + P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^{c}\}.$$

By the properties  $(\alpha_r)$ ,  $(\beta_r)$  and Lemma 1

(3.9) 
$$\sup_{\theta \in K} P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^c\} = O(n^{-r/2}).$$

Then it is obvious that the assertion (b) holds for the case  $1 \le r < k$  and for the case  $k \ge 2$  and 0 < r < 1 because of (3.6), (3.8) and (3.9). It remains to prove the assertion (b) for the case  $r \ge k$  and for the case k=1 and 0 < r < 1.

In the case  $r \ge k$ , we shall show that the first term on the right side of (3.8) has upper bound of order  $O(n^{-k/2})$ . Because of condition (ii) in  $C_{k,r}$ , choose  $d_K > 0$ ,  $\lambda_K(x)$  and  $D_{n,\theta,K}$  as in the proof of Lemma 1. Let

$$M_{n,\theta} = \{z_n \in X^n; ||T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1}|| < n^{-1} (\log n)^{2\pi_1 + 1/2} \}.$$

According to Lemma 2

(3.10) 
$$\sup_{\theta \in K} P_{n,\theta}\{(M_{n,\theta})^c\} = O(n^{-k/2}).$$

We must again estimate the second term on the right side of (3.5). Since for  $\theta \in K$  and  $z_n \in M_{n,\theta}$ 

$$||T_n - \theta|| < \rho_K n^{-1} || \sum_{i=1}^n g^{(1)}(x_i, \theta) || + n^{-1} (\log n)^{2\pi_1 + 1/2}$$

with  $\rho_{\kappa} = \sup_{\theta \in K} ||J(\theta)^{-1}||$ , it follows from Minkowski's inequality that

$$\begin{split} [E_{\theta}(||T_{n}-\theta||^{k+2}I_{M_{n},\theta}(\cdot))]^{1/(k+2)} &\leq \rho_{K}n^{-1}[E_{\theta}(||\sum_{i=1}^{n}g^{(1)}(\cdot,\theta)||^{k+2})]^{1/(k+2)} \\ &+ n^{-1}(\log n)^{2\pi_{1}+1/2} \,. \end{split}$$

(2.10) with Condition  $B_r$  implies that for  $\theta \in K$ 

$$E_{\theta}(\|\sum_{i=1}^{n} g^{(1)}(\cdot, \theta)\|^{k+2}) \leq c_{\kappa} n^{(k+2)/2},$$

which leads to

(3.11) 
$$[E_{\theta}(||T_n - \theta||^{k+2}I_{M_{n,\theta}}(\cdot))]^{1/(k+2)} \leq c_K n^{-1/2}$$

Thus we have for  $n \ge n_K$  and  $\theta \in K$ 

(3.12) 
$$E_{\theta}(||G_{n}^{(k+1)}(\cdot, T_{n}^{*}) - G_{n}^{(k+1)}(\cdot, \theta)|| ||T_{n} - \theta||^{k+1}I_{U_{n,\theta} \cap D_{n,\theta,\kappa} \cap M_{n,\theta}}(\cdot)) \leq (1 + \sup_{\theta \in \kappa} E_{\theta}(\lambda_{\kappa}(\cdot)))nE_{\theta}(||T_{n} - \theta||^{k+2}I_{M_{n,\theta}}(\cdot)) \leq c_{\kappa}n^{-k/2}.$$

By Hölder's inequality

$$\begin{split} & E_{\theta}(||G_{n}^{(k+1)}(\cdot,\theta) - E_{\theta}(G_{n}^{(k+1)}(\cdot,\theta))|| ||T_{n} - \theta||^{k+1}I_{M_{n},\theta}(\cdot)) \\ & \leq [E_{\theta}(||G_{n}^{(k+1)}(\cdot,\theta) - E_{\theta}(G_{n}^{(k+1)}(\cdot,\theta))||^{k+2})]^{1/(k+2)}[E_{\theta}(||T_{n} - \theta||^{k+2}I_{M_{n},\theta})]^{(k+1)/(k+2)}, \\ & \text{so that (2.10) with condition (ii) (c) in } C_{k,r} \text{ and (3.11) imply that for } \theta \in K \\ & (3.13) \quad E_{\theta}(||G_{n}^{(k+1)}(\cdot,\theta) - E_{\theta}(G_{n}^{(k+1)}(\cdot,\theta))|| ||T_{n} - \theta||^{k+1}I_{M_{n},\theta}(\cdot)) \leq c_{K}n^{-k/2}. \end{split}$$

Taking account of (3.7), we obtain for  $n \ge n_K$  and  $\theta \in K$ 

$$E_{\theta}\left(\left|\log\frac{\overline{q}_{n,k}(\cdot,\theta)}{p_{n}(\cdot,\theta)}\right|I_{U_{n,\theta}\cap V_{n,r}\cap W_{n,r}}(\cdot)\right)$$

$$\leq E_{\theta}\left(\left|\log\frac{\overline{q}_{n,k}(\cdot,\theta)}{p_{n}(\cdot,\theta)}\right|I_{U_{n,\theta}\cap V_{n,r}\cap W_{n,r}\cap D_{n,\theta,K}\cap M_{n,\theta}}(\cdot)\right)$$

$$+\left(\log 2\right)P_{n,\theta}\left\{\left(D_{n,\theta,K}\cap M_{n,\theta}\right)^{c}\right\}.$$

Thus, the first term on the right side of (3.8) has upper bound of order  $O(n^{-k/2})$  because of (3.1), (3.5), (3.10), (3.12) and (3.13).

This, together with (3.8) and (3.9), implies that

$$\sup_{\theta\in K} ||P_{n,\theta} - \overline{Q}_{n,\theta}^k|| = O(n^{-k/2}).$$

Since

$$\sup_{\theta \in \mathcal{K}} |1 - v_n(\theta)^{-1}| = \sup_{\theta \in \mathcal{K}} |P_{n,\theta}\{X^n\} - \overline{Q}_{n,\theta}^k\{X^n\}|$$
$$= O(n^{-k/2}),$$

we have

$$\begin{split} \sup_{\theta \in \mathcal{K}} ||P_{n,\theta} - Q_{n,\theta}^k|| &\leq \sup_{\theta \in \mathcal{K}} ||P_{n,\theta} - \overline{Q}_{n,\theta}^k|| + \sup_{\theta \in \mathcal{K}} ||\overline{Q}_{n,\theta}^k - Q_{n,\theta}^k|| \\ &\leq \sup_{\theta \in \mathcal{K}} ||P_{n,\theta} - \overline{Q}_{n,\theta}^k|| + \sup_{\theta \in \mathcal{K}} |1 - v_n(\theta)^{-1}| \\ &= O(n^{-k/2}) \,, \end{split}$$

which is the desired result.

In the case k=1 and 0 < r < 1,  $M_{n,\theta}$  is replaced by the following set  $M_{n,\theta,r}$ 

$$M_{n,\theta,r} = \{z_n \in X^n; ||T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1}|| < n^{-(r+1)/2} (\log n)^{2\pi}\}.$$

Then, a similar argument shows that

$$\sup_{\theta\in\mathcal{K}}||P_{n,\theta}-Q_{n,\theta}^1||=O(n^{-r/2}).$$

This completes the proof.

REMARK. (1) If  $r \ge k$ , it is possible to choose  $Q_{n,\theta}^k$  independent of r because  $V_{n,r}$  and  $W_{n,r}$  in the definition of  $\overline{q}_{n,k}$  can be replaced by  $V_{n,k}$  and  $W_{n,k}$ , respectively.

(2) In the case k=1, it follows from Theorem 2 that a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$  is asymptotically sufficient up to order  $O(n^{-r/2})$  if 0 < r < 1 and  $O(n^{-1/2})$  if r=1. The latter result has been already shown by Pfanzagl [10] under similar circumstances to ours.

(3) Michel [7] showed that  $T_{n,k}$ ,  $k \ge 3$ , constructed by asymptotic m.l. estimators of order  $o(n^{-(k-2)/2})$  is asymptotically sufficient up to order  $o(n^{-(k-2)/2})$ . According to Theorem 2, the convergence order concerning asymptotic sufficiency of  $T_{n,k}$  can be improved up to  $O(n^{-k/2})$  if  $\{T_n\}$  is a sequence of asymptotic m.l. estimators with higher order than Michel's one.

(4) In [14], [15] Suzuki assumes the existence of moment generating function of  $g^{(k+1)}(x, \theta)$  to evaluate probability of large deviations. Of course this condition is stronger than ours.

4. Properties of m.l. estimators. We shall investigate conditions under which a sequence of m.l. estimators has the properties  $(\alpha_r)$  and  $(\beta_r)$  for some r>0.

Let  $\overline{\Theta}$  denote the closure of  $\Theta$  in  $\overline{\mathbf{R}}^s = [-\infty, \infty]^s$ . Assume that  $g(\cdot, \theta)$ :  $X \to \mathbf{R}, \theta \in \Theta$ , admits a measurable extension  $g(\cdot, \theta): X \to \overline{\mathbf{R}}, \theta \in \overline{\Theta}$ .

Condition  $A^*$ 

(i)  $E_{\theta}(g(\cdot, \tau)) < E_{\theta}(g(\cdot, \theta))$  for all  $\theta \in \Theta, \tau \in \overline{\Theta}, \theta \neq \tau$ .

(ii) For every  $x \in X$ ,  $\theta \rightarrow g(x, \theta)$  is continuous on  $\overline{\Theta}$ .

Condition  $B_r^*$ 

(i) For every  $\theta \in \Theta$  and every compact  $K \subset \Theta$ 

$$\sup_{\tau\in\kappa}E_{\tau}(|g(\cdot,\theta)|^{(r+2)/2})<\infty.$$

(ii) For every  $\theta \in \overline{\Theta}$  there exists a neighborhood  $U_{\theta}$  of  $\theta$  such that for every neighborhood U of  $\theta$ ,  $U \subset U_{\theta}$ , and every compact  $K \subset \Theta$ 

$$\sup_{\tau\in K} E_{\tau}(|\sup_{\sigma\in \overline{\sigma}}g(\cdot,\sigma)|^{(r+2)/2}) < \infty.$$

(iii) For each  $x \in X$ ,  $\theta \to g(x, \theta)$  admits continuous partial derivatives up to the order 2 on  $\Theta$ . For every  $\theta \in \Theta$  there exist a neighborhood  $U_{\theta}$  of  $\theta$  and a measurable function  $\lambda(x, \theta)$  such that

- (a) for all  $x \in X$ ,  $\tau, \sigma \in U_{\theta}$ ,  $||g^{(2)}(x, \tau) g^{(2)}(x, \sigma)|| \leq ||\tau \sigma||\lambda(x, \theta)|$
- (b) for every compact  $K \subset \Theta$ ,  $\sup_{\tau \in K} E_{\tau}(\lambda(\cdot, \theta)^{(r+2)/2}) < \infty$
- (c)  $\sup_{\tau \in \overline{U}_{\theta}} E_{\tau}(||g^{(2)}(\cdot, \tau)||^{(r+2)/2}) < \infty$ .
- (iv)  $\theta \rightarrow J(\theta)$  is continuous on  $\Theta$ .

A maximum likelihood estimator for the sample size n is an estimator  $T_n$  for which  $T_n \in \overline{\Theta}$  and

$$\sum_{i=1}^{n} g(x_i, T_n) = \sup_{\theta \in \bar{\Theta}} \sum_{i=1}^{n} g(x_i, \theta) .$$

Condition (ii) in  $A^*$  insures that m.l. estimators for the sample size n exist. Let  $\hat{T}_n, n \in N$ , be a sequence of m.l. estimators.

The following lemma can be obtained in a way analogous to the one used in the proof of Lemma 4 in Michel and Pfanzagl [8] except that Theorem 1 is used instead of Chebyshev's inequality.

**Lemma 3.** Let Condition  $A^*$  and conditions (i), (ii) in  $B_r^*$  be satisfied for some r>0. Then for every  $\varepsilon > 0$  and every compact  $K \subset \Theta$ 

$$\sup_{\theta\in\mathcal{K}}P_{n,\theta}\{z_n\in X^n; ||\hat{T}_n(z_n)-\theta||\geq\varepsilon\}=O(n^{-r/2}).$$

The following proposition is an immediate consequence of Lemma 3.

**Proposition 1.** Let Condition  $A^*$  and conditions (i), (ii) in  $B^*_r$  be satisfied for some r>0. Moreover, assume that for each  $x \in X$ ,  $\theta \rightarrow g(x, \theta)$  is continuously differentiable on  $\Theta$ . Then for every compact  $K \subset \Theta$ 

$$\sup_{\theta \in \mathcal{K}} P_{n,\theta} \{ z_n \in X^n; || \sum_{i=1}^n g^{(1)}(x_i, \hat{T}_n(z_n)) || > 0 \} = O(n^{-r/2})$$

**Lemma 4** (cf. Lemma 5 in Michel and Pfanzagl [8]). Let Condition  $A^*$ and conditions (i)–(iii) in  $B_r^*$  be satisfied for some r>0. Then for every  $\delta>0$ 

and every compact  $K \subset \Theta$  there exists d > 0 such that

$$\sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; \sup_{||\hat{T}_n - \tau|| \leq d} ||n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) + J(\theta)]|| \geq \delta \} = O(n^{-r/2}).$$

Proof. Let  $\delta > 0$  be given and K be a compact subset of  $\Theta$ . By condition (iii) in  $B_r^*$  we may choose  $d_K > 0$ ,  $\lambda_K(x)$  and  $D_{n,\theta,K}$  as in Lemma 1 with k=1. We write

$$F_{n,\theta,\delta} = \{z_n \in X^n; ||n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]|| < \delta/2\}.$$

From condition (iii) (c) in  $B_r^*$  it follows that

$$\sup_{\theta\in K} P_{n,\theta}\{(F_{n,\theta,\delta})^c\} = O(n^{-r/2}).$$

Taking  $2d = \min \{ d_K, \delta/[2(1 + \sup_{\theta \in K} E_{\theta}(\lambda_K(\cdot)))] \}$ , we see that for  $z_n \in D_{n,\theta,K} \cap F_{n,\theta,\delta}, ||\hat{T}_n - \theta|| < d$  and  $||\hat{T}_n - \tau|| \leq d$ 

$$||n^{-1}\sum_{i=1}^{n} [g^{(2)}(x_{i}, \tau) + J(\theta)]|| \leq ||n^{-1}\sum_{i=1}^{n} [g^{(2)}(x_{i}, \tau) - g^{(2)}(x_{i}, \theta)]|| + ||n^{-1}\sum_{i=1}^{n} [g^{(2)}(x_{i}, \theta) + J(\theta)]|| < \delta.$$

This together with Lemma 3 implies the desired assertion.

Lemma 3 and Lemma 4 yield the following proposition.

**Proposition 2** (cf. Lemma 6 in [8] and Lemma 3 in Pfanzagl [12]). Assume that Conditions A,  $A^*$ , B, and  $B_r^*$  are fulfilled for some r>0. Then for every compact  $K \subset \Theta$  there exists  $c_K > 0$  such that

$$\sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; n^{1/2} || \hat{T}_n(z_n) - \theta || \ge c_K (\log n)^{1/2} \} = O(n^{-r/2}) .$$

Proof. Let K be a fixed compact subset of  $\Theta$ . It follows from conditions (ii) (b) in A and (iv) in  $B_r^*$  that there exists  $\delta_K > 0$  such that  $\theta \in K$  and matrix J with  $||J-J(\theta)|| < \delta_K$  imply that J is regular and  $||J^{-1}-J(\theta)^{-1}|| < 1$ . Let

$$W_{n,\theta}^* = \{z_n \in X^n; \sup_{\|\hat{T}_n - \tau\| \leq d_K} \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) + J(\theta)]\| < \delta_K\},\$$

where  $d_{\kappa} > 0$  is chosen to satisfy that

$$\sup_{\theta\in K} P_{n,\theta}\{(W_{n,\theta}^*)^c\} = O(n^{-r/2})$$

because of Lemma 4. Choose  $e_{\kappa} > 0$  such that  $e_{\kappa} \leq d_{\kappa}$  and  $\{\tau \in \mathbb{R}^{s}; \inf_{\theta \in \kappa} ||\theta - \tau|| \leq e_{\kappa}\} \subset \Theta$ . Let

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$$U_{n,\theta}^* = \{z_n \in X^n; ||\hat{T}_n - \theta|| < e_K\}.$$

In view of Lemma 3 we have

$$\sup_{\theta \in \mathcal{K}} P_{n,\theta}\{(U_{n,\theta}^*)^c\} = O(n^{-r/2}).$$

Since for  $\theta \in K$  and  $z_n \in U_{n,\theta}^*$ 

$$\sum_{i=1}^{n} g^{(1)}(x_{i}, \theta) = (\theta - \hat{T}_{n}) \sum_{i=1}^{n} \bar{g}^{(2)}(x_{i}, \hat{T}_{n}, \theta) ,$$

it follows that for  $\theta \in K$  and  $z_n \in U_{n,\theta}^* \cap W_{n,\theta}^*$ 

$$\begin{aligned} ||n^{1/2}(\hat{T}_n - \theta)|| &\leq ||n^{-1/2} \sum_{i=1}^n g^{(1)}(x_i, \theta)|| ||(-n^{-1} \sum_{i=1}^n \bar{g}^{(2)}(x_i, \hat{T}_n, \theta))^{-1}|| \\ &\leq (1 + \sup_{\theta \in K} ||J(\theta)^{-1}||)n^{-1/2}||\sum_{i=1}^n g^{(1)}(x_i, \theta)|| . \end{aligned}$$

In order to complete the proof it is enough to note that there exists  $c_{\kappa} > 0$  such that

$$\sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; \| \sum_{i=1}^n g^{(1)}(x_i, \theta) \| \ge c_K (n \log n)^{1/2} \} = o(n^{-r/2}).$$

This follows from Theorem 1, condition (ii) (a) in A and Condition  $B_r$ .

REMARK. (1) Proposition 2 remains to hold for a sequence of minimum contrast estimators with obvious modification.

(2) If every (r+2)/2 in Condition  $B_r^*$  is replaced by a number greater than it, then Proposition 2 holds with  $o(n^{-r/2})$  instead of  $O(n^{-r/2})$ .

(3) Proposition 2 improves Lemma 3 of Pfanzagl [12] in the following sense:

(a) This result still holds for 0 < r < 1.

(b) In the case  $r \ge 1$ , the moment conditions used in Proposition 2 are weaker than in [12] because of the use of Theorem 1 instead of Lemma 2 of [12] (see Remark (2) of Theorem 1).

From Theorem 2, Proposition 1 and Proposition 2, the following theorem is immediate.

**Theorem 3.** Assume that Conditions A,  $A^*$ ,  $B_r$ , (i), (ii) in  $B_r^*$ ,  $C_{1,r}$  and  $C_{k,r}$  are fulfilled for some  $k \in N$  and r > 0. Then,  $\hat{T}_{n,k} = (\hat{T}_n, G_n^{(2)}(z_n, \hat{T}_n), \cdots, G_n^{(k)}(z_n, \hat{T}_n))$  is asymptotically sufficient up to order  $O(n^{-r/2})$  if r < k and  $O(n^{-k/2})$  if  $r \geq k$ . Here  $\hat{T}_{n,1}$  means  $\hat{T}_n$ .

It is remarked that we need the (2+r)-th absolute moment of  $g^{(1)}$  and the (2+r)/(2-r)-th absolute moment of  $g^{(2)}$  in order to show that a sequence

of m.l. estimators is asymptotically sufficient up to order  $O(n^{-r/2})$  with  $0 < r \le 1$ . Acknowledgment. The author wishes to express his hearty thanks to

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