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## ASYMPTOTIC SUFFICIENCY I: REGULAR CASES

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**1. Introduction.** The concept of asymptotic sufficiency of maximum likelihood (m.l.) estimator is due to Wald [16] and this work was succeeded by LeCam [4] and Pfanzagl [10]. Higher order asymptotic sufficiency has been subsequently studied by Ghosh and Subramanyam [3], Michel [7] and Suzuki [14], [15].

Let  $\Theta$  be an open subset of the  $s$ -dimensional Euclidean space. Suppose that  $x_1, \dots, x_n$  are independent and identically distributed random variables with joint distribution  $P_{n,\theta}$ ,  $\theta \in \Theta$ , which has a constant support and satisfies certain regularity conditions. For  $\theta \in \Theta$  and  $z_n = (x_1, \dots, x_n)$  let  $G_n^{(m)}(z_n, \theta)$  denote the  $m$ -th derivative relative to  $\theta$  of the log-likelihood function. In Michel [7], it was shown that for  $k \geq 3$  a statistic  $T_{n,k} = (T_n, G_n^{(2)}(z_n, T_n), \dots, G_n^{(k)}(z_n, T_n))$ , where  $\{T_n\}$  is a sequence of asymptotic m.l. estimators of order  $o(n^{-(k-2)/2})$  (see Definition in Section 3), is asymptotically sufficient up to order  $o(n^{-(k-2)/2})$  in the following sense: For each  $n \in N$ ,  $T_{n,k}$  is sufficient for a family  $\{Q_{n,\theta}; \theta \in \Theta\}$  of probability distributions and for every compact subset  $K$  of  $\Theta$

$$\sup_{\theta \in K} \|P_{n,\theta} - Q_{n,\theta}\| = o(n^{-(k-2)/2}),$$

where  $\|\cdot\|$  means the total variation of a measure. Suzuki [14], [15] also showed that for  $k \in N$  a statistic  $(\hat{\theta}_n, G_n^{(1)}(z_n, \hat{\theta}_n), \dots, G_n^{(k)}(z_n, \hat{\theta}_n))$ , where  $\hat{\theta}_n$  is a reasonable estimator including m.l. estimator, is asymptotically sufficient up to order  $o(n^{-(k-1)/2})$  under a stronger moment condition than in Michel [7].

In this paper we give a refinement of their results on higher order asymptotic sufficiency. Our result includes that (1)  $T_{n,k} = (T_n, G_n^{(2)}(z_n, T_n), \dots, G_n^{(k)}(z_n, T_n))$  is asymptotically sufficient up to order  $O(n^{-k/2})$  for any sequence  $\{T_n\}$  of asymptotic m.l. estimators of order  $O(n^{-k/2})$  and (2) a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$  with some  $r \in (0, 1)$  is asymptotically sufficient up to order  $O(n^{-r/2})$  under mild moment conditions for the first and the second derivatives of the log-likelihood function.

In the case  $k=1$ , Pfanzagl ([10], Theorem 1) proved that a sequence of estimators with properties analogous to those of asymptotic m.l. estimators of order  $O(n^{-1/2})$  is asymptotically sufficient up to order  $O(n^{-1/2})$ , and showed in

[11] that this order of convergence cannot be improved in general. Thus our result is an extension of his and it seems to be impossible to improve the convergence order  $O(n^{-k/2})$ .

In Section 2 we present a result concerning probabilities of deviations for sums of independent and identically distributed random variables with a restricted moment. In Section 3 we investigate asymptotic sufficiency of  $T_{n,k}$  constructed by asymptotic m.l. estimators  $T_n$ . In the final Section 4 we give conditions under which a sequence of m.l. estimators becomes the one of asymptotic m.l. estimators of order  $O(n^{-r/2})$  with some  $r > 0$ .

**2. Probabilities of deviations.** Let  $Y_1, \dots, Y_n$  be a sequence of random variables (r.v.'s) and put  $S_m = \sum_{i=1}^m Y_i$ ,  $1 \leq m \leq n$ . Using the elementary inequality

$$E|S_n|^r \leq \sum_{i=1}^n E|Y_i|^r, \quad r \leq 1,$$

it follows from Markov's inequality that for  $x > 0$

$$(2.1) \quad P\{|S_n| \geq x\} \leq x^{-r} \sum_{i=1}^n E|Y_i|^r.$$

If the r.v.'s satisfy the relations

$$(2.2) \quad E(Y_{m+1}|S_m) = 0 \quad \text{a.s.} \quad 1 \leq m \leq n-1,$$

then von Bahr and Esseen [1] showed that

$$(2.3) \quad E|S_n|^r \leq 2 \sum_{i=1}^n E|Y_i|^r, \quad 1 \leq r \leq 2.$$

The condition (2.2) is fulfilled if the r.v.'s are independent and have zero means. In this case, (2.3) together with Markov's inequality implies the following inequality

$$(2.4) \quad P\{|S_n| \geq x\} \leq 2x^{-r} \sum_{i=1}^n E|Y_i|^r, \quad 1 \leq r \leq 2,$$

for  $x > 0$ .

The following theorem includes a uniform version of Corollary 2 in Nagaev [9].

**Theorem 1.** *Let  $Y_1, \dots, Y_n$  be a sequence of independent and identically distributed random variables with a common distribution  $P_\theta$ ,  $\theta \in K$ , where  $K$  is any set. Let  $h(y, \theta)$  be a measurable function of  $y$  for any fixed  $\theta \in K$  and put  $S_{n,\theta} = \sum_{i=1}^n h(Y_i, \theta)$ . If  $E_\theta(h(Y_1, \theta)) = 0$  for all  $\theta \in K$  and  $\xi_r = \sup_{\theta \in K} E_\theta|h(Y_1, \theta)|^r < \infty$  for some  $r > 0$ , then*

$$(2.5) \quad \sup_{\theta \in K} P_{\theta} \{ |S_{n,\theta}| \geq x \} = O(nx^{-r}), \quad 0 < r \leq 2,$$

for  $x > 0$ , and

$$(2.6) \quad \sup_{\theta \in K} P_{\theta} \{ |S_{n,\theta}| > \xi_r^{1/r} x \} = O(nx^{-r}), \quad r > 2,$$

for  $x \geq \sqrt{8(r-2)n \log n}$ .

Proof. (2.5) is an immediate consequence of (2.1) and (2.4).

For the proof of (2.6) we use the following inequality which is a slight modification of Theorem 1 in Nagaev [9]: For  $x > 0$  and  $y > 0$ ,

$$(2.7) \quad P_{\theta} \{ |S_{n,\theta}| > x \} < 2nP_{\theta} \{ |h(Y_1, \theta)| > y \} + 2 \left[ \frac{n\xi_{r,\theta} d_r}{y^r} \right]^{x/y} \\ \times \exp \left\{ 2n \left[ \frac{r \log y - \log (n\xi_{r,\theta} d_r)}{y} \right]^2 \xi_{r,\theta}^{2/r} + 1 \right\},$$

where  $d_r = 1 + (r+1)r^{r+2} \exp(-r)$  and  $\xi_{r,\theta} = E_{\theta} |h(Y_1, \theta)|^r$ . In order to show (2.7) it is enough to note that the relation (2.3) in [9] becomes

$$\left| \int_{-\infty}^{1/h} \exp \{ h[h(Y_1, \theta)] \} dP_{\theta} - 1 \right| < 2h^2 \xi_{r,\theta}^{2/r}.$$

Setting  $x = \xi_{r,\theta}^{1/r} n^{1/2} t$  and  $y = x/2$  for  $t \geq \sqrt{8(r-2) \log n}$  in (2.7), then we obtain

$$(2.8) \quad nP_{\theta} \{ |h(Y_1, \theta)| > y \} \leq n\xi_{r,\theta} y^{-r} \\ = 2^r n^{(2-r)/2} t^{-r}$$

and

$$(2.9) \quad \left[ \frac{n\xi_{r,\theta} d_r}{y^r} \right]^{x/y} = 2^{2r} d_r^2 n^{2-r} t^{-2r}.$$

Let us assume that  $n \geq \exp \left\{ \frac{1}{8(r-2)} \right\}$ . (For  $n < \exp \left\{ \frac{1}{8(r-2)} \right\}$ , (2.6) is trivially true.) Since  $0 \leq t^{-1} \log t < 1/2$  for  $t \geq 1$ , we have

$$2n \left[ \frac{r \log y - \log (n\xi_{r,\theta} d_r)}{y} \right]^2 \xi_{r,\theta}^{2/r} \\ = 8t^{-2} \left[ r \log t + \frac{r-2}{2} \log n - r \log 2 - \log d_r \right]^2 \\ \leq 8t^{-2} \left[ r^2 (\log t)^2 + \frac{(r-2)^2}{4} (\log n)^2 + r(r-2) \log n \log t + c_1 \right] \\ \leq 2r^2 + \frac{r-2}{4} \log n + r \log t + c_2,$$

where  $c_1$  and  $c_2$  denote positive constants depending only on  $r$ . From this fact and (2.9) it follows that the second term on the right side of (2.7) has an upper bound of the type  $c_3 n^{3(2-r)/4} t^{-r}$ . This, together with (2.8), implies (2.6).

REMARK. (1) In the case  $r \geq 3$ , Michel [7] showed a result analogous to Theorem 1 (cf. also Lemma 1 in Pfanzagl [12]).

(2) Let  $Y_1, \dots, Y_n$  be a sequence of independent r.v.'s with zero means. It follows from an inequality due to Marcinkiewicz and Zygmund [5] that

$$(2.10) \quad E \left| \sum_{i=1}^n Y_i \right|^r \leq c n^{(r-2)/2} \sum_{i=1}^n E |Y_i|^r, \quad r \geq 2,$$

where  $c$  is a positive constant depending only on  $r$  (see Chung [2], page 348). This leads to Lemma 2 in Pfanzagl [12] which requires a stronger moment condition than in Theorem 1 to evaluate probability of moderate deviations or large deviations.

**3. Asymptotic sufficiency.** Let  $\Theta$  be an open subset of the  $s$ -dimensional Euclidean space  $\mathbf{R}^s$  and for each  $\theta \in \Theta$ , let  $P_\theta$  be a probability measure on a measurable space  $(X, \mathcal{A})$ . It is assumed that  $P_\theta, \theta \in \Theta$ , is dominated by a  $\sigma$ -finite measure  $\mu$  on  $(X, \mathcal{A})$  and has a positive density  $p(x, \theta)$ . For each  $n \in N = \{1, 2, \dots\}$ , let  $(X^n, \mathcal{A}^n)$  be the Cartesian product of  $n$  copies of  $(X, \mathcal{A})$  and  $P_{n,\theta}$  be the product measure of  $n$  copies of  $P_\theta$ . Furthermore, let  $\mu_n$  denote the product measure of  $n$  copies of  $\mu$  and write  $p_n(z_n, \theta) = dP_{n,\theta} / d\mu_n$  for  $\theta \in \Theta$  and  $z_n = (x_1, \dots, x_n) \in X^n$ .

For a function  $h(z, \cdot): \mathbf{R}^s \rightarrow \mathbf{R}$  denote the  $m$ -th derivative relative to  $\theta$  of  $h(z, \theta)$  by

$$h^{(m)}(z, \theta) = \left( \frac{\partial^m}{\partial \theta_{i_1} \dots \partial \theta_{i_m}} h(z, \theta); i_1, \dots, i_m \in \{1, \dots, s\} \right).$$

In particular, we write

$$h^{(1)}(z, \theta) = \left( \frac{\partial}{\partial \theta_1} h(z, \theta), \dots, \frac{\partial}{\partial \theta_s} h(z, \theta) \right),$$

$$h^{(2)}(z, \theta) = \left( \frac{\partial^2}{\partial \theta_i \partial \theta_j} h(z, \theta) \right),$$

that is,  $h^{(1)}$  means a row vector and  $h^{(2)}$  a matrix. The Euclidean norm  $\|\cdot\|$  of  $h^{(m)}$  is defined by

$$\|h^{(m)}(z, \theta)\|^2 = \sum_{i_1, \dots, i_m=1}^s \left( \frac{\partial^m}{\partial \theta_{i_1} \dots \partial \theta_{i_m}} h(z, \theta) \right)^2.$$

For any  $\sigma = (\sigma_1, \dots, \sigma_s) \in \mathbf{R}^s$  define

$$h^{(m)}(z, \theta) \sigma^m = \sum_{i_1, \dots, i_m=1}^s \frac{\partial^m}{\partial \theta_{i_1} \dots \partial \theta_{i_m}} h(z, \theta) \prod_{p=1}^m \sigma_{i_p}.$$

Then, it is easy to see that

$$|h^{(m)}(z, \theta)\sigma^m| \leq \|h^{(m)}(z, \theta)\| \|\sigma\|^m.$$

Let  $k \in N$  and  $r > 0$  be fixed. We shall impose the following Conditions  $A$ ,  $B_r$  and  $C_{k,r}$  on  $p(x, \theta)$ .

Condition  $A$

(i) For each  $x \in X$ ,  $\theta \rightarrow p(x, \theta)$  admits continuous partial derivatives up to the order 2 on  $\Theta$ .

Let  $g(x, \theta) = \log p(x, \theta)$  and  $g^{(m)}$  be the  $m$ -th derivative of  $g$  defined above. Moreover, for  $\theta \in \Theta$  let  $J(\theta) = E_\theta(-g^{(2)}(\cdot, \theta))$ .

(ii) For every  $\theta \in \Theta$

(a)  $E_\theta(g^{(1)}(\cdot, \theta)) = 0$

(b)  $J(\theta)$  is positive definite.

Condition  $B_r$

For every compact  $K \subset \Theta$

$$\sup_{\theta \in K} E_\theta(\|g^{(1)}(\cdot, \theta)\|^{r+2}) < \infty.$$

Condition  $C_{k,r}$

(i) For each  $x \in X$ ,  $\theta \rightarrow p(x, \theta)$  admits continuous partial derivatives up to the order  $k+1$  on  $\Theta$ .

(ii) For every  $\theta \in \Theta$  there exist a neighborhood  $U_\theta$  of  $\theta$  and a measurable function  $\lambda(x, \theta)$  such that

(a) for all  $x \in X$ ,  $\tau, \sigma \in U_\theta$ ,  $\|g^{(k+1)}(x, \tau) - g^{(k+1)}(x, \sigma)\| \leq \|\tau - \sigma\| \lambda(x, \theta)$

(b) for every compact  $K \subset \Theta$ ,  $\sup_{\tau \in K} E_\tau(\lambda(\cdot, \theta)^{(r+2)/2}) < \infty$

(c)  $\sup_{\tau \in U_\theta} E_\tau(\|g^{(k+1)}(\cdot, \tau)\|^{v(r)}) < \infty$ ,

where

$$\begin{aligned} v(r) &= \frac{2+r}{2-r}, & \text{if } 0 < r < 1, \\ &= r+2, & \text{if } r \geq 1. \end{aligned}$$

(iii) For every compact  $K \subset \Theta$  there exist  $\delta_K > 0$  and  $\eta_K > 0$  such that  $\theta \in K$  and  $\tau \in \Theta$  with  $\|\theta - \tau\| < \delta_K$  imply

$$\|E_\theta(g^{(k+1)}(\cdot, \theta)) - E_\tau(g^{(k+1)}(\cdot, \tau))\| \leq \eta_K \|\theta - \tau\|.$$

REMARK. (1) Condition (iii) in  $C_{k,r}$  follows from conditions (3)(a) and (3)(b) in Suzuki [15] (see also (3.4) in [14]).

(2) It is easily seen that condition (ii) in  $C_{k,r}$  and the following condition (iii)' imply condition (iii) in  $C_{k,r}$ .

(iii)' For every  $\theta \in \Theta$  there exist a neighborhood  $U_\theta$  of  $\theta$  and a measurable function  $\lambda^*(x, \theta)$  such that for all  $x \in X$ ,  $\tau \in U_\theta$

$$|p(x, \tau)/p(x, \theta) - 1| \leq \|\tau - \theta\| \lambda^*(x, \theta)$$

and for every compact  $K \subset \Theta$

$$\sup_{\tau \in K} E_{\tau}(\lambda^*(\cdot, \theta))^{v(r)/(v(r)-1)} < \infty .$$

The following definition is due to Michel [7].

DEFINITION.  $T_n, n \in N$ , is a sequence of *asymptotic maximum likelihood (m.l.) estimators of order  $O(n^{-r/2})$ ,  $r > 0$* , if there exist positive constants  $\pi_1$  and  $\pi_2$  (depending on  $r$ ) such that for every compact  $K \subset \Theta$

$$(\alpha_r) \sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; n^{1/2} \|T_n(z_n) - \theta\| \geq (\log n)^{\pi_1}\} = O(n^{-r/2})$$

$$(\beta_r) \sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; n^{r/2} \|\sum_{i=1}^n g^{(1)}(x_i, T_n(z_n))\| \geq (\log n)^{\pi_2}\} = O(n^{-r/2}) .$$

Asymptotic m.l. estimators can be obtained from suitable initial estimators by applying a Newton-Raphson method (see Michel [6] and Pfanzagl [13]).

To simplify our notations we shall use  $n_K$  (depending on compact  $K$ ) as a generic constant instead of the phrase “for all sufficiently large  $n$ ”. In the same manner we shall use  $c_K$  as a generic constant to denote factors occurring in the bounds which depend on compact  $K$  but not on  $\theta \in K$  and  $n \in N$ .

**Lemma 1.** *Assume that Condition  $C_{k,r}$  is fulfilled for some  $k \in N$  and  $r > 0$ . Let  $T_n, n \in N$ , be a sequence of estimators with the property  $(\alpha_r)$ . Then for every compact  $K \subset \Theta$*

$$\sup_{\theta \in K} P_{n,\theta} \left\{ \sup_{n^{1/2} \|T_n - \tau\| \leq (\log n)^{\pi_1}} \left\| \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_{\tau}(g^{(k+1)}(\cdot, \tau))] \right\| \geq \psi(n, r) \right\} = O(n^{-r/2}),$$

where

$$\begin{aligned} \psi(n, r) &= n^{(2-r)/2}, & \text{if } 0 < r < 1, \\ &= n^{1/2}(\log n)^{\pi_1+1/2}, & \text{if } r \geq 1. \end{aligned}$$

Proof. Let  $0 < r < 1$  and  $K$  be a compact subset of  $\Theta$ . Condition (ii) implies that there exist  $d_K > 0$  and  $\lambda_K(x)$  such that  $\theta \in K$  and  $\tau \in \Theta$  with  $\|\theta - \tau\| < d_K$  imply  $\|g^{(k+1)}(x, \theta) - g^{(k+1)}(x, \tau)\| \leq \|\theta - \tau\| \lambda_K(x)$  for all  $x \in X$ , and such that  $\sup_{\theta \in K} E_{\theta}(\lambda_K(\cdot))^{(r+2)/2} < \infty$ . Let

$$D_{n,\theta,K} = \{z_n \in X^n; |\sum_{i=1}^n [\lambda_K(x_i) - E_{\theta}(\lambda_K(\cdot))]| < n\} .$$

According to Theorem 1

$$(3.1) \quad \sup_{\theta \in K} P_{n,\theta} \{(D_{n,\theta,K})^c\} = O(n^{-r/2}) .$$

Furthermore, Theorem 1 together with condition (ii) (c) implies that

$$(3.2) \quad \sup_{\theta \in K} P_{n,\theta} \{ (F_{n,\theta})^c \} = O(n^{-r/2}),$$

where

$$F_{n,\theta} = \{ z_n \in X^n; \| \sum_{i=1}^n [g^{(k+1)}(x_i, \theta) - E_\theta(g^{(k+1)}(\cdot, \theta))] \| < 1/2 n^{(2-r)/2} \}.$$

Let  $e_K > 0$  be such that  $\{ \tau \in \mathbf{R}^s; \inf_{\theta \in K} \|\theta - \tau\| \leq e_K \} \subset \Theta$ . Choose  $n_K$  to satisfy  $2n^{-1/2}(\log n)^{\pi_1} < \min \{ d_K, e_K, \delta_K \}$  for all  $n \geq n_K$ , where  $\delta_K$  appears in condition (iii). Then, by conditions (ii) and (iii), for  $n \geq n_K$ ,  $\theta \in K$ ,  $\tau \in \mathbf{R}^s$  with  $\|\theta - \tau\| \leq 2n^{-1/2}(\log n)^{\pi_1}$  and  $z_n \in D_{n,\theta,K} \cap F_{n,\theta}$

$$\begin{aligned} & \| \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_\tau(g^{(k+1)}(\cdot, \tau))] \| \\ & \leq \| \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - g^{(k+1)}(x_i, \theta)] \| + \| \sum_{i=1}^n [g^{(k+1)}(x_i, \theta) - E_\theta(g^{(k+1)}(\cdot, \theta))] \| \\ & \quad + n \| E_\theta(g^{(k+1)}(\cdot, \theta)) - E_\tau(g^{(k+1)}(\cdot, \tau)) \| \\ & \leq n [1 + \sup_{\theta \in K} E_\theta(\lambda_K(\cdot)) + \eta_K] \|\theta - \tau\| + 1/2 n^{(2-r)/2} \\ & < n^{(2-r)/2}. \end{aligned}$$

Taking account of (3.1) and (3.2), for every compact  $K \subset \Theta$  we obtain

$$\begin{aligned} & \sup_{\theta \in K} P_{n,\theta} \{ \sup_{n^{1/2}\|\theta - \tau\| \leq 2(\log n)^{\pi_1}} \| \sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_\tau(g^{(k+1)}(\cdot, \tau))] \| \geq n^{(2-r)/2} \} \\ & = O(n^{-r/2}). \end{aligned}$$

This together with the property  $(\alpha_r)$  leads to the desired assertion.

In the case  $r \geq 1$ , it is enough to show that there exists  $c_K > 0$  such that

$$(3.3) \quad \sup_{\theta \in K} P_{n,\theta} \{ (F_{n,\theta,K})^c \} = o(n^{-r/2}),$$

where

$$F_{n,\theta,K} = \{ z_n \in X^n; \| \sum_{i=1}^n [g^{(k+1)}(x_i, \theta) - E_\theta(g^{(k+1)}(\cdot, \theta))] \| \leq c_K (n \log n)^{1/2} \}.$$

This follows from Theorem 1 and condition (ii) (c).

**Lemma 2.** Assume that Conditions A, B, and  $C_{1,r}$  are fulfilled for some  $r > 0$ . Let  $T_n$ ,  $n \in N$ , be a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$ . Then for every compact  $K \subset \Theta$

$$\sup_{\theta \in K} P_{n,\theta} \{ z_n \in X^n; \| T_n(z_n) - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1} \| \geq \omega(n, r) \} = O(n^{-r/2}),$$

where



$$\begin{aligned} \omega(n, r) &= n^{-(r+1)/2}(\log n)^{2\pi_1}, & \text{if } 0 < r < 1, \\ &= n^{-1}(\log n)^{2\pi_1+1/2}, & \text{if } r \geq 1. \end{aligned}$$

Proof. Let  $0 < r < 1$  and  $K$  be a compact subset of  $\Theta$ . Condition (ii) in  $C_{1,r}$  implies that there exist  $d_K > 0$  and  $\lambda_K(x)$  such that  $x \in X$ ,  $\theta \in K$  and  $\tau \in \Theta$  with  $\|\theta - \tau\| < d_K$  imply  $\|g^{(2)}(x, \theta) - g^{(2)}(x, \tau)\| \leq \|\theta - \tau\| \lambda_K(x)$  and such that  $\sup_{\theta \in K} E_\theta(\lambda_K(\cdot)^{(r+2)/2}) < \infty$ . As in the proof of Lemma 1, we define

$$\begin{aligned} D_{n,\theta,K} &= \{z_n \in X^n; |\sum_{i=1}^n [\lambda_K(x_i) - E_\theta(\lambda_K(\cdot))]| < n\}, \\ F_{n,\theta} &= \{z_n \in X^n; \|\sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]\| < 1/2 n^{(2-r)/2}\}. \end{aligned}$$

It follows from Theorem 1, condition (ii) (a) in  $A$  and Condition  $B_r$ , that there exists  $c_K > 0$  such that

$$(3.4) \quad \sup_{\theta \in K} P_{n,\theta} \{(H_{n,\theta,K})^c\} = o(n^{-r/2}),$$

where

$$H_{n,\theta,K} = \{z_n \in X^n; \|\sum_{i=1}^n g^{(1)}(x_i, \theta)\| < c_K(n \log n)^{1/2}\}.$$

Let  $U_{n,\theta}$  and  $V_{n,r}$  be defined by

$$\begin{aligned} U_{n,\theta} &= \{z_n \in X^n; n^{1/2} \|T_n(z_n) - \theta\| < (\log n)^{\pi_1}\}, \\ V_{n,r} &= \{z_n \in X^n; n^{r/2} \|\sum_{i=1}^n g^{(1)}(x_i, T_n(z_n))\| < (\log n)^{\pi_2}\}. \end{aligned}$$

Choose  $e_K > 0$  such that  $K^* = \{\tau \in \mathbf{R}^s; \inf_{\theta \in K} \|\theta - \tau\| \leq e_K\} \subset \Theta$  and  $n_K$  such that  $n^{-1/2}(\log n)^{\pi_1} < \min\{d_K, e_K, \delta_K\}$  for all  $n \geq n_K$ , where  $\delta_K$  is determined by condition (iii) in  $C_{1,r}$ . It is obvious that  $n \geq n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta}$  imply  $T_n(z_n) \in K^*$ . Since  $K^*$  is a compact subset of  $\Theta$ , conditions (ii) (b) in  $A$  and (iii) in  $C_{1,r}$  imply that

$$\rho_{K^*} = \sup_{\tau \in K^*} \|J(\tau)^{-1}\| < \infty.$$

Using the equality

$$\sum_{i=1}^n g^{(1)}(x_i, \theta) = \sum_{i=1}^n g^{(1)}(x_i, T_n) + (\theta - T_n) \sum_{i=1}^n \bar{g}^{(2)}(x_i, T_n, \theta)$$

with  $\bar{g}^{(2)}(x, \theta, \sigma) = \int_0^1 g^{(2)}(x, (1-t)\theta + t\sigma) dt$ , we obtain

$$\begin{aligned} &T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1} + n^{-1} \sum_{i=1}^n g^{(1)}(x_i, T_n) J(T_n)^{-1} \\ &= T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) [J(\theta)^{-1} - J(T_n)^{-1}] + (T_n - \theta) n^{-1} \sum_{i=1}^n \bar{g}^{(2)}(x_i, T_n, \theta) J(T_n)^{-1} \end{aligned}$$

$$\begin{aligned}
&= (T_n - \theta)[J(T_n) - J(\theta)]J(T_n)^{-1} - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta)J(\theta)^{-1}[J(T_n) - J(\theta)]J(T_n)^{-1} \\
&\quad + (T_n - \theta)n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, T_n, \theta) - g^{(2)}(x_i, \theta)]J(T_n)^{-1} \\
&\quad + (T_n - \theta)n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]J(T_n)^{-1}.
\end{aligned}$$

Hence we have for  $n \geq n_K$ ,  $\theta \in K$  and  $z_n \in D_{n,\theta,K} \cap F_{n,\theta} \cap H_{n,\theta,K} \cap U_{n,\theta} \cap V_{n,r}$

$$\begin{aligned}
&\|T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta)J(\theta)^{-1}\| \\
&\leq \|n^{-1} \sum_{i=1}^n g^{(1)}(x_i, T_n)J(T_n)^{-1}\| + \|T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta)J(\theta)^{-1}\| \\
&\quad + \|n^{-1} \sum_{i=1}^n g^{(1)}(x_i, T_n)J(T_n)^{-1}\| \\
&\leq \rho_{K^*} n^{-(r+2)/2} (\log n)^{\alpha_2} + (c_K \eta_K \rho_{K^*}^2 n^{-1/2} (\log n)^{1/2} + 1/2 \rho_{K^*} n^{-r/2}) \|T_n - \theta\| \\
&\quad + \rho_{K^*} (1 + \eta_K + \sup_{\theta \in K} E_\theta(\lambda_K(\cdot))) \|T_n - \theta\|^2 \\
&\leq c_K n^{-(r+1)/2} (\log n)^{\alpha_1}.
\end{aligned}$$

This implies the desired result because of (3.1), (3.2), (3.4) and the properties  $(\alpha_r)$ ,  $(\beta_r)$ .

For the case  $r \geq 1$ , the proof is also similar except that  $F_{n,\theta}$  is replaced by  $F_{n,\theta,K}$  in (3.3) with  $k=1$ .

For simplicity, we write

$$G_n^{(m)}(z_n, \theta) = \sum_{i=1}^n g^{(m)}(x_i, \theta), \quad z_n = (x_1, \dots, x_n) \in X^n, \quad \theta \in \Theta.$$

Now we can present a result on asymptotic sufficiency of the statistic

$$\begin{aligned}
T_{n,k} &= T_n, & k=1, \\
&= (T_n, G_n^{(2)}(z_n, T_n), \dots, G_n^{(k)}(z_n, T_n)), & k \geq 2,
\end{aligned}$$

where  $T_n$ ,  $n \in N$ , is a sequence of asymptotic m.l. estimators.

**Theorem 2.** *Assume that Conditions  $A$ ,  $B_r$ ,  $C_{1,r}$  and  $C_{k,r}$  hold for some  $k \in N$  and  $r > 0$ . Let  $T_n$ ,  $n \in N$ , be a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$ . Then there exists a sequence of families of probability measures  $\{Q_{n,\theta}^k; \theta \in \Theta\}$ ,  $n \in N$ , such that*

- (a) *for each  $n \in N$ ,  $T_{n,k}$  is sufficient for  $\{Q_{n,\theta}^k; \theta \in \Theta\}$*
- (b) *for every compact  $K \subset \Theta$*

$$\begin{aligned} \sup_{\theta \in K} \|P_{n,\theta} - Q_{n,\theta}^k\| &= O(n^{-r/2}), & \text{if } r < k, \\ &= O(n^{-k/2}), & \text{if } r \geq k. \end{aligned}$$

Proof. Let

$$U_{n,\theta} = \{z_n \in X^n; n^{1/2} \|T_n - \theta\| < (\log n)^{\pi_1}\},$$

$$V_{n,r} = \{z_n \in X^n; n^{r/2} \|\sum_{i=1}^n g^{(1)}(x_i, T_n)\| < (\log n)^{\pi_2}\},$$

$$W_{n,r} = \{z_n \in X^n; \sup_{n^{1/2} \|T_n - \tau\| \leq (\log n)^{\pi_1}} \|\sum_{i=1}^n [g^{(k+1)}(x_i, \tau) - E_\tau(g^{(k+1)}(\cdot, \tau))]\| < \psi(n, r)\},$$

where  $\psi(n, r)$  is the same as in Lemma 1. We define

$$\begin{aligned} \bar{q}_{n,k}(z_n, \theta) &= I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}}(z_n) \exp \{G_n(z_n, T_n) + \sum_{m=2}^k \frac{1}{m!} G_n^{(m)}(z_n, T_n)(\theta - T_n)^m \\ &\quad + \frac{1}{(k+1)!} [E_\theta(G_n^{(k+1)}(\cdot, \theta))](\theta - T_n)^{k+1}\}, \end{aligned}$$

$$q_{n,k}(z_n, \theta) = v_n(\theta) \bar{q}_{n,k}(z_n, \theta),$$

where  $v_n(\theta) = [\int_{X^n} \bar{q}_{n,k}(z_n, \theta) d\mu_n]^{-1}$ . Here and hereafter  $I_U(\cdot)$  means the indicator function of a set  $U$ . For  $\theta \in \Theta$  and  $n \in N$ , we denote by  $\bar{Q}_{n,\theta}^k$  and  $Q_{n,\theta}^k$  the measures given by

$$\frac{d\bar{Q}_{n,\theta}^k}{d\mu_n} = \bar{q}_{n,k} \quad \text{and} \quad \frac{dQ_{n,\theta}^k}{d\mu_n} = q_{n,k}.$$

Then it follows from the factorization theorem that for each  $n \in N$ ,  $T_{n,k}$  is sufficient for  $\{Q_{n,\theta}^k; \theta \in \Theta\}$ .

In order to prove the second assertion (b) we fix a compact subset  $K$  of  $\Theta$ . Using the Taylor expansion

$$\begin{aligned} G_n(z_n, \theta) &= G_n(z_n, T_n) + \sum_{m=1}^k \frac{1}{m!} G_n^{(m)}(z_n, T_n)(\theta - T_n)^m \\ &\quad + \frac{1}{(k+1)!} G_n^{(k+1)}(z_n, T_n^*)(\theta - T_n)^{k+1} \end{aligned}$$

where  $\max \{\|T_n^* - \theta\|, \|T_n^* - T_n\|\} \leq \|T_n - \theta\|$ , we have for  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$

$$\begin{aligned} (3.5) \quad & \left| \log \frac{\bar{q}_{n,k}(z_n, \theta)}{\hat{p}_n(z_n, \theta)} \right| \\ & \leq \|G_n^{(1)}(z_n, T_n)\| \|T_n - \theta\| + \frac{1}{(k+1)!} \|G_n^{(k+1)}(z_n, T_n^*) \\ & \quad - E_\theta(G_n^{(k+1)}(\cdot, \theta))\| \|T_n - \theta\|^{k+1}. \end{aligned}$$

Since

$$\begin{aligned} & \|G_n^{(k+1)}(z_n, T_n^*) - E_\theta(G_n^{(k+1)}(\cdot, \theta))\| \\ & \leq \|G_n^{(k+1)}(z_n, T_n^*) - [E_\tau(G_n^{(k+1)}(\cdot, \tau))]_{\tau=T_n^*}\| \\ & \quad + \|[E_\tau(G_n^{(k+1)}(\cdot, \tau))]_{\tau=T_n^*} - E_\theta(G_n^{(k+1)}(\cdot, \theta))\|, \end{aligned}$$

it follows from (3.5) and condition (iii) in  $C_{k,r}$  that for  $n \geq n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$

$$(3.6) \quad \left| \log \frac{\bar{q}_{n,k}(z_n, \theta)}{p_n(z_n, \theta)} \right| \leq n^{-(r+1)/2} (\log n)^{\pi_1 + \pi_2} + n^{-(k+1)/2} (\log n)^{(k+1)\pi_1} \psi(n, r).$$

This implies that for  $n \geq n_K$ ,  $\theta \in K$  and  $z_n \in U_{n,\theta} \cap V_{n,r} \cap W_{n,r}$

$$(3.7) \quad \left| \log \frac{\bar{q}_{n,k}(z_n, \theta)}{p_n(z_n, \theta)} \right| \leq \log 2.$$

Using the inequality  $|1 - \exp(x)| \leq 2|x|$  for  $|x| \leq \log 2$ , then from (3.7) we have for  $n \geq n_K$  and  $\theta \in K$

$$\begin{aligned} (3.8) \quad & \|P_{n,\theta} - \bar{Q}_{n,\theta}\| \\ & \leq \int_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}} \left| 1 - \frac{\bar{q}_{n,k}(z_n, \theta)}{p_n(z_n, \theta)} \right| dP_{n,\theta} + P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^c\} \\ & \leq 2E_\theta \left[ \left| \log \frac{\bar{q}_{n,k}(\cdot, \theta)}{p_n(\cdot, \theta)} \right| I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}}(\cdot) \right] + P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^c\}. \end{aligned}$$

By the properties  $(\alpha_r)$ ,  $(\beta_r)$  and Lemma 1

$$(3.9) \quad \sup_{\theta \in K} P_{n,\theta}\{(U_{n,\theta} \cap V_{n,r} \cap W_{n,r})^c\} = O(n^{-r/2}).$$

Then it is obvious that the assertion (b) holds for the case  $1 \leq r < k$  and for the case  $k \geq 2$  and  $0 < r < 1$  because of (3.6), (3.8) and (3.9). It remains to prove the assertion (b) for the case  $r \geq k$  and for the case  $k=1$  and  $0 < r < 1$ .

In the case  $r \geq k$ , we shall show that the first term on the right side of (3.8) has upper bound of order  $O(n^{-k/2})$ . Because of condition (ii) in  $C_{k,r}$ , choose  $d_K > 0$ ,  $\lambda_K(x)$  and  $D_{n,\theta,K}$  as in the proof of Lemma 1. Let

$$M_{n,\theta} = \{z_n \in X^n; \|T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1}\| < n^{-1} (\log n)^{2\pi_1 + 1/2}\}.$$

According to Lemma 2

$$(3.10) \quad \sup_{\theta \in K} P_{n,\theta}\{M_{n,\theta}^c\} = O(n^{-k/2}).$$

We must again estimate the second term on the right side of (3.5). Since for  $\theta \in K$  and  $z_n \in M_{n,\theta}$

$$\|T_n - \theta\| < \rho_K n^{-1} \left\| \sum_{i=1}^n g^{(1)}(x_i, \theta) \right\| + n^{-1} (\log n)^{2\pi_1 + 1/2}$$

with  $\rho_K = \sup_{\theta \in K} \|J(\theta)^{-1}\|$ , it follows from Minkowski's inequality that

$$[E_\theta(\|T_n - \theta\|^{k+2} I_{M_{n,\theta}}(\cdot))]^{1/(k+2)} \leq \rho_K n^{-1} [E_\theta(\|\sum_{i=1}^n g^{(1)}(\cdot, \theta)\|^{k+2})]^{1/(k+2)} \\ + n^{-1} (\log n)^{2\pi_1 + 1/2}.$$

(2.10) with Condition  $B$ , implies that for  $\theta \in K$

$$E_\theta(\|\sum_{i=1}^n g^{(1)}(\cdot, \theta)\|^{k+2}) \leq c_K n^{(k+2)/2},$$

which leads to

$$(3.11) \quad [E_\theta(\|T_n - \theta\|^{k+2} I_{M_{n,\theta}}(\cdot))]^{1/(k+2)} \leq c_K n^{-1/2}.$$

Thus we have for  $n \geq n_K$  and  $\theta \in K$

$$(3.12) \quad E_\theta(\|G_n^{(k+1)}(\cdot, T_n^*) - G_n^{(k+1)}(\cdot, \theta)\| \|T_n - \theta\|^{k+1} I_{U_{n,\theta} \cap D_{n,\theta,K} \cap M_{n,\theta}}(\cdot)) \\ \leq (1 + \sup_{\theta \in K} E_\theta(\lambda_K(\cdot))) n E_\theta(\|T_n - \theta\|^{k+2} I_{M_{n,\theta}}(\cdot)) \leq c_K n^{-k/2}.$$

By Hölder's inequality

$$E_\theta(\|G_n^{(k+1)}(\cdot, \theta) - E_\theta(G_n^{(k+1)}(\cdot, \theta))\| \|T_n - \theta\|^{k+1} I_{M_{n,\theta}}(\cdot)) \\ \leq [E_\theta(\|G_n^{(k+1)}(\cdot, \theta) - E_\theta(G_n^{(k+1)}(\cdot, \theta))\|^{k+2})]^{1/(k+2)} [E_\theta(\|T_n - \theta\|^{k+2} I_{M_{n,\theta}})]^{(k+1)/(k+2)},$$

so that (2.10) with condition (ii) (c) in  $C_{k,r}$  and (3.11) imply that for  $\theta \in K$

$$(3.13) \quad E_\theta(\|G_n^{(k+1)}(\cdot, \theta) - E_\theta(G_n^{(k+1)}(\cdot, \theta))\| \|T_n - \theta\|^{k+1} I_{M_{n,\theta}}(\cdot)) \leq c_K n^{-k/2}.$$

Taking account of (3.7), we obtain for  $n \geq n_K$  and  $\theta \in K$

$$E_\theta\left(\left|\log \frac{\bar{q}_{n,k}(\cdot, \theta)}{p_n(\cdot, \theta)}\right| I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r}}(\cdot)\right) \\ \leq E_\theta\left(\left|\log \frac{\bar{q}_{n,k}(\cdot, \theta)}{p_n(\cdot, \theta)}\right| I_{U_{n,\theta} \cap V_{n,r} \cap W_{n,r} \cap D_{n,\theta,K} \cap M_{n,\theta}}(\cdot)\right) \\ + (\log 2) P_{n,\theta}\{(D_{n,\theta,K} \cap M_{n,\theta})^c\}.$$

Thus, the first term on the right side of (3.8) has upper bound of order  $O(n^{-k/2})$  because of (3.1), (3.5), (3.10), (3.12) and (3.13).

This, together with (3.8) and (3.9), implies that

$$\sup_{\theta \in K} \|P_{n,\theta} - \bar{Q}_{n,\theta}^k\| = O(n^{-k/2}).$$

Since

$$\sup_{\theta \in K} |1 - v_n(\theta)^{-1}| = \sup_{\theta \in K} |P_{n,\theta}\{X^n\} - \bar{Q}_{n,\theta}^k\{X^n\}| \\ = O(n^{-k/2}),$$

we have

$$\begin{aligned} \sup_{\theta \in \bar{K}} \|P_{n,\theta} - Q_{n,\theta}^k\| &\leq \sup_{\theta \in \bar{K}} \|P_{n,\theta} - \bar{Q}_{n,\theta}^k\| + \sup_{\theta \in \bar{K}} \|\bar{Q}_{n,\theta}^k - Q_{n,\theta}^k\| \\ &\leq \sup_{\theta \in \bar{K}} \|P_{n,\theta} - \bar{Q}_{n,\theta}^k\| + \sup_{\theta \in \bar{K}} |1 - v_n(\theta)^{-1}| \\ &= O(n^{-k/2}), \end{aligned}$$

which is the desired result.

In the case  $k=1$  and  $0 < r < 1$ ,  $M_{n,\theta}$  is replaced by the following set  $M_{n,\theta,r}$

$$M_{n,\theta,r} = \{z_n \in X^n; \|T_n - \theta - n^{-1} \sum_{i=1}^n g^{(1)}(x_i, \theta) J(\theta)^{-1}\| < n^{-(r+1)/2} (\log n)^{2\alpha_1}\}.$$

Then, a similar argument shows that

$$\sup_{\theta \in \bar{K}} \|P_{n,\theta} - Q_{n,\theta}^1\| = O(n^{-r/2}).$$

This completes the proof.

REMARK. (1) If  $r \geq k$ , it is possible to choose  $Q_{n,\theta}^k$  independent of  $r$  because  $V_{n,r}$  and  $W_{n,r}$  in the definition of  $\bar{q}_{n,k}$  can be replaced by  $V_{n,k}$  and  $W_{n,k}$ , respectively.

(2) In the case  $k=1$ , it follows from Theorem 2 that a sequence of asymptotic m.l. estimators of order  $O(n^{-r/2})$  is asymptotically sufficient up to order  $O(n^{-r/2})$  if  $0 < r < 1$  and  $O(n^{-1/2})$  if  $r=1$ . The latter result has been already shown by Pfanzagl [10] under similar circumstances to ours.

(3) Michel [7] showed that  $T_{n,k}$ ,  $k \geq 3$ , constructed by asymptotic m.l. estimators of order  $o(n^{-(k-2)/2})$  is asymptotically sufficient up to order  $o(n^{-(k-2)/2})$ . According to Theorem 2, the convergence order concerning asymptotic sufficiency of  $T_{n,k}$  can be improved up to  $O(n^{-k/2})$  if  $\{T_n\}$  is a sequence of asymptotic m.l. estimators with higher order than Michel's one.

(4) In [14], [15] Suzuki assumes the existence of moment generating function of  $g^{(k+1)}(x, \theta)$  to evaluate probability of large deviations. Of course this condition is stronger than ours.

**4. Properties of m.l. estimators.** We shall investigate conditions under which a sequence of m.l. estimators has the properties  $(\alpha_r)$  and  $(\beta_r)$  for some  $r > 0$ .

Let  $\bar{\Theta}$  denote the closure of  $\Theta$  in  $\bar{\mathbf{R}}^s = [-\infty, \infty]^s$ . Assume that  $g(\cdot, \theta): X \rightarrow \mathbf{R}$ ,  $\theta \in \Theta$ , admits a measurable extension  $g(\cdot, \theta): X \rightarrow \bar{\mathbf{R}}$ ,  $\theta \in \bar{\Theta}$ .

Condition  $A^*$

- (i)  $E_\theta(g(\cdot, \tau)) < E_\theta(g(\cdot, \theta))$  for all  $\theta \in \Theta$ ,  $\tau \in \bar{\Theta}$ ,  $\theta \neq \tau$ .
- (ii) For every  $x \in X$ ,  $\theta \rightarrow g(x, \theta)$  is continuous on  $\bar{\Theta}$ .

Condition  $B_r^*$

(i) For every  $\theta \in \Theta$  and every compact  $K \subset \Theta$

$$\sup_{\tau \in K} E_\tau(|g(\cdot, \theta)|^{(r+2)/2}) < \infty .$$

(ii) For every  $\theta \in \bar{\Theta}$  there exists a neighborhood  $U_\theta$  of  $\theta$  such that for every neighborhood  $U$  of  $\theta$ ,  $U \subset U_\theta$ , and every compact  $K \subset \Theta$

$$\sup_{\tau \in K} E_\tau(|\sup_{\sigma \in U} g(\cdot, \sigma)|^{(r+2)/2}) < \infty .$$

(iii) For each  $x \in X$ ,  $\theta \rightarrow g(x, \theta)$  admits continuous partial derivatives up to the order 2 on  $\Theta$ . For every  $\theta \in \Theta$  there exist a neighborhood  $U_\theta$  of  $\theta$  and a measurable function  $\lambda(x, \theta)$  such that

(a) for all  $x \in X$ ,  $\tau, \sigma \in U_\theta$ ,  $\|g^{(2)}(x, \tau) - g^{(2)}(x, \sigma)\| \leq \|\tau - \sigma\| \lambda(x, \theta)$

(b) for every compact  $K \subset \Theta$ ,  $\sup_{\tau \in K} E_\tau(\lambda(\cdot, \theta)^{(r+2)/2}) < \infty$

(c)  $\sup_{\tau \in U_\theta} E_\tau(\|g^{(2)}(\cdot, \tau)\|^{(r+2)/2}) < \infty .$

(iv)  $\theta \rightarrow J(\theta)$  is continuous on  $\Theta$ .

A *maximum likelihood estimator* for the sample size  $n$  is an estimator  $T_n$  for which  $T_n \in \bar{\Theta}$  and

$$\sum_{i=1}^n g(x_i, T_n) = \sup_{\theta \in \bar{\Theta}} \sum_{i=1}^n g(x_i, \theta) .$$

Condition (ii) in  $A^*$  insures that m.l. estimators for the sample size  $n$  exist. Let  $\hat{T}_n, n \in N$ , be a sequence of m.l. estimators.

The following lemma can be obtained in a way analogous to the one used in the proof of Lemma 4 in Michel and Pfanzagl [8] except that Theorem 1 is used instead of Chebyshev's inequality.

**Lemma 3.** *Let Condition  $A^*$  and conditions (i), (ii) in  $B_r^*$  be satisfied for some  $r > 0$ . Then for every  $\varepsilon > 0$  and every compact  $K \subset \Theta$*

$$\sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; \|\hat{T}_n(z_n) - \theta\| \geq \varepsilon\} = O(n^{-r/2}) .$$

The following proposition is an immediate consequence of Lemma 3.

**Proposition 1.** *Let Condition  $A^*$  and conditions (i), (ii) in  $B_r^*$  be satisfied for some  $r > 0$ . Moreover, assume that for each  $x \in X$ ,  $\theta \rightarrow g(x, \theta)$  is continuously differentiable on  $\Theta$ . Then for every compact  $K \subset \Theta$*

$$\sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; \|\sum_{i=1}^n g^{(1)}(x_i, \hat{T}_n(z_n))\| > 0\} = O(n^{-r/2}) .$$

**Lemma 4** (cf. Lemma 5 in Michel and Pfanzagl [8]). *Let Condition  $A^*$  and conditions (i)–(iii) in  $B_r^*$  be satisfied for some  $r > 0$ . Then for every  $\delta > 0$*

and every compact  $K \subset \Theta$  there exists  $d > 0$  such that

$$\sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; \sup_{\|\hat{T}_n - \tau\| \leq d} \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) + J(\theta)]\| \geq \delta\} = O(n^{-r/2}).$$

Proof. Let  $\delta > 0$  be given and  $K$  be a compact subset of  $\Theta$ . By condition (iii) in  $B_r^*$  we may choose  $d_K > 0$ ,  $\lambda_K(x)$  and  $D_{n,\theta,K}$  as in Lemma 1 with  $k=1$ . We write

$$F_{n,\theta,\delta} = \{z_n \in X^n; \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]\| < \delta/2\}.$$

From condition (iii) (c) in  $B_r^*$  it follows that

$$\sup_{\theta \in K} P_{n,\theta} \{(F_{n,\theta,\delta})^c\} = O(n^{-r/2}).$$

Taking  $2d = \min \{d_K, \delta/[2(1 + \sup_{\theta \in K} E_\theta(\lambda_K(\cdot)))]\}$ , we see that for  $z_n \in D_{n,\theta,K} \cap F_{n,\theta,\delta}$ ,  $\|\hat{T}_n - \theta\| < d$  and  $\|\hat{T}_n - \tau\| \leq d$

$$\begin{aligned} \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) + J(\theta)]\| &\leq \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) - g^{(2)}(x_i, \theta)]\| \\ &\quad + \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \theta) + J(\theta)]\| < \delta. \end{aligned}$$

This together with Lemma 3 implies the desired assertion.

Lemma 3 and Lemma 4 yield the following proposition.

**Proposition 2** (cf. Lemma 6 in [8] and Lemma 3 in Pfanzagl [12]). *Assume that Conditions A, A\*, B<sub>r</sub> and B<sub>r</sub>\* are fulfilled for some  $r > 0$ . Then for every compact  $K \subset \Theta$  there exists  $c_K > 0$  such that*

$$\sup_{\theta \in K} P_{n,\theta} \{z_n \in X^n; n^{1/2} \|\hat{T}_n(z_n) - \theta\| \geq c_K (\log n)^{1/2}\} = O(n^{-r/2}).$$

Proof. Let  $K$  be a fixed compact subset of  $\Theta$ . It follows from conditions (ii) (b) in  $A$  and (iv) in  $B_r^*$  that there exists  $\delta_K > 0$  such that  $\theta \in K$  and matrix  $J$  with  $\|J - J(\theta)\| < \delta_K$  imply that  $J$  is regular and  $\|J^{-1} - J(\theta)^{-1}\| < 1$ . Let

$$W_{n,\theta}^* = \{z_n \in X^n; \sup_{\|\hat{T}_n - \tau\| \leq d_K} \|n^{-1} \sum_{i=1}^n [g^{(2)}(x_i, \tau) + J(\theta)]\| < \delta_K\},$$

where  $d_K > 0$  is chosen to satisfy that

$$\sup_{\theta \in K} P_{n,\theta} \{(W_{n,\theta}^*)^c\} = O(n^{-r/2})$$

because of Lemma 4. Choose  $e_K > 0$  such that  $e_K \leq d_K$  and  $\{\tau \in \mathbf{R}^s; \inf_{\theta \in K} \|\theta - \tau\| \leq e_K\} \subset \Theta$ . Let



$$U_{n,\theta}^* = \{z_n \in X^n; \|\hat{T}_n - \theta\| < e_K\}.$$

In view of Lemma 3 we have

$$\sup_{\theta \in K} P_{n,\theta}\{(U_{n,\theta}^*)^c\} = O(n^{-r/2}).$$

Since for  $\theta \in K$  and  $z_n \in U_{n,\theta}^*$

$$\sum_{i=1}^n g^{(1)}(x_i, \theta) = (\theta - \hat{T}_n) \sum_{i=1}^n \bar{g}^{(2)}(x_i, \hat{T}_n, \theta),$$

it follows that for  $\theta \in K$  and  $z_n \in U_{n,\theta}^* \cap W_{n,\theta}^*$

$$\begin{aligned} \|n^{1/2}(\hat{T}_n - \theta)\| &\leq \|n^{-1/2} \sum_{i=1}^n g^{(1)}(x_i, \theta)\| \|(-n^{-1} \sum_{i=1}^n \bar{g}^{(2)}(x_i, \hat{T}_n, \theta))^{-1}\| \\ &\leq (1 + \sup_{\theta \in K} \|J(\theta)^{-1}\|) n^{-1/2} \|\sum_{i=1}^n g^{(1)}(x_i, \theta)\|. \end{aligned}$$

In order to complete the proof it is enough to note that there exists  $c_K > 0$  such that

$$\sup_{\theta \in K} P_{n,\theta}\{z_n \in X^n; \|\sum_{i=1}^n g^{(1)}(x_i, \theta)\| \geq c_K(n \log n)^{1/2}\} = o(n^{-r/2}).$$

This follows from Theorem 1, condition (ii) (a) in *A* and Condition *B*.

REMARK. (1) Proposition 2 remains to hold for a sequence of minimum contrast estimators with obvious modification.

(2) If every  $(r+2)/2$  in Condition  $B_r^*$  is replaced by a number greater than it, then Proposition 2 holds with  $o(n^{-r/2})$  instead of  $O(n^{-r/2})$ .

(3) Proposition 2 improves Lemma 3 of Pfanzagl [12] in the following sense:

(a) This result still holds for  $0 < r < 1$ .

(b) In the case  $r \geq 1$ , the moment conditions used in Proposition 2 are weaker than in [12] because of the use of Theorem 1 instead of Lemma 2 of [12] (see Remark (2) of Theorem 1).

From Theorem 2, Proposition 1 and Proposition 2, the following theorem is immediate.

**Theorem 3.** *Assume that Conditions A, A\*, B<sub>r</sub>, (i), (ii) in B<sub>r</sub><sup>\*</sup>, C<sub>1,r</sub> and C<sub>k,r</sub> are fulfilled for some k ∈ N and r > 0. Then,  $\hat{T}_{n,k} = (\hat{T}_n, G_n^{(2)}(z_n, \hat{T}_n), \dots, G_n^{(k)}(z_n, \hat{T}_n))$  is asymptotically sufficient up to order  $O(n^{-r/2})$  if  $r < k$  and  $O(n^{-k/2})$  if  $r \geq k$ . Here  $\hat{T}_{n,1}$  means  $\hat{T}_n$ .*

It is remarked that we need the  $(2+r)$ -th absolute moment of  $g^{(1)}$  and the  $(2+r)/(2-r)$ -th absolute moment of  $g^{(2)}$  in order to show that a sequence

of m.l. estimators is asymptotically sufficient up to order  $O(n^{-r/2})$  with  $0 < r \leq 1$ .

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